United States
Department of
Agriculture
Food and
Nutrition
Service
Office of
Analysis,
Nutrition, and
Evaluation

# Household Food Security in the United States, 1995-1997 

## Technical Issues and Statistical Report

# Measuring Food Security in the United States 

# Household Food Security in the United States, 1995-1997: Technical Issues and Statistical Report 

Final Report of the Project<br>to Analyze 1996 and 1997<br>Food Security Data

December 2001

James Ohls
Larry Radbill
Allen Schirm

Submitted to:
U.S. Department of Agriculture

Food and Nutrition Service
3101 Park Center Drive
Alexandria, VA 22302

Submitted by:
Mathematica Policy Research, Inc.
P.O. Box 2393

Princeton, NJ 08543-2393
Telephone: (609) 799-3535
Facsimile: (609) 799-0005
Project Director: James Ohls

## ACKNOWLEDGMENTS

The authors would like to thank Gary Bickel, the project officer for the study, for guidance and suggestions throughout the work. Valuable advice and comments were also received from Mark Nord and Steven Carlson. The advice of our Advisory Committee members also proved extremely useful at certain key points in the research. They included Christine Olson, Ben Wright, and Cheryl Wehler. Robert Mislevy also provided critical suggestions on Rasch modeling and the interpretation of Rasch model results, and Don Rubin made suggestions about dealing with missing data. Jeff Passel assisted in calculating correct weights from the CPS files. In addition to people already named, the following provided important advice at a research conference where parts of this material were presented: Susan Mayer, Christopher Jencks, Margaret Andrews, David Smallwood, Jean Pierre Habicht, and Helen Jensen.

Extraordinary programming support was provided by Abhijay Prakash, Jean Knab, and Jeanne Bellotti.

Roy Grisham edited the manuscript and substantially improved the clarity of the exposition. Key production support was provided by Jane Nelson, Jill Miller, and Marjorie Mitchell.

## CONTENTS

## Chapter

Page
EXECUTIVE SUMMARY ..... XV
I INTRODUCTION ..... 1
A. PURPOSE OF THE REPORT ..... 1
B. OVERVIEW OF REPORT ..... 2
II BACKGROUND ..... 5
A. A BRIEF SUMMARY OF THE LITERATURE INFORMING THE FOOD SECURITY CONCEPT ..... 6
B. THE CPS DATA ON FOOD SECURITY ..... 10
C. THE ONE-PARAMETER LOGISTIC ITEM RESPONSE THEORY (RASCH) MODEL ..... 13
D. RELATIONSHIP BETWEEN IRT MODELING AND THE FOOD INSECURITY PREVALENCE ESTIMATES ..... 15
E. CONVENTIONS USED IN THIS REPORT ..... 17
III STABILITY OF ESTIMATED MODEL PARAMETERS OVER TIME ..... 19
IV STANDARD BENCHMARK METHODOLOGY ..... 23
A. THE BASIC GOVERNMENT METHOD (NO RECALIBRATION OF THE MODEL) ..... 24

1. Check and Edit the Data ..... 24
2. Construct Binary Variables from the Raw Data for Each of the 18 Food Security Variables ..... 24
3. Screening ..... 24
4. Impute Missing Data ..... 25

## CONTENTS (continued)

Chapter Page
IV
(Continued)
5. For Each Household, Count the Number of Affirmative Answers to the 18 Food Security Variables ..... 25
6. Determine the Continuous Scale Score and Food Security Classification ..... 25
B. CHECKING THE CALIBRATION OF THE MODEL ON THE DATA SETS ..... 30

1. Edit the Data ..... 30
2. Construct Binary Variables from the Raw Data for Each of the 18 Food Security Variables ..... 30
3. Screening ..... 30
4. Estimate the Rasch Model Parameters ..... 30
V DEFINING "CUTPOINTS" OVER TIME ..... 33
5. Earlier Work ..... 33
6. Generalization of Cutpoints to Other Years ..... 35
7. Alternative Approaches to Setting Cutpoints ..... 37
VI CONCLUSION ..... 41
A. STRENGTHS AND LIMITATIONS OF THE METHODOLOGY ..... 41
B. DIRECTIONS FOR FURTHER RESEARCH ..... 44
REFERENCES ..... 47
APPENDIX A: SUMMARY OF SALIENT PARTS OF ITEM RESPONSE THEORY AND RELATED MEASUREMENT ISSUES ..... 51
APPENDIX B: METHOD FOR SCREENING HOUSEHOLDS INTO THE FOOD SECURITY ANALYSIS ..... 79
APPENDIX C: IMPUTING MISSING FOOD SECURITY DATA ..... 91
APPENDIX D: HOW "MODAL" ARE HOUSEHOLD RESPONSE PATTERNS? ..... 97

## CONTENTS (continued)

Chapter Page
APPENDIX E: ITEM SEVERITY LEVELS BY SUBGROUPS . . . . . . . . . . 107
$\begin{aligned} \text { APPENDIX F: } & \text { PREVALENCE ESTIMATES AND THEIR } \\ & \text { STANDARD ERRORS . . . . . . . . . . . . . . . . . . . . . . . . . . } 129\end{aligned}$

## TABLES

Table Page
II. 1 PREVALENCE OF FOOD SECURITY FOR HOUSEHOLDS, BY YEAR ..... 9
II. 2 CPS SAMPLE SIZES ..... 12
III. 1 COMPARISON OF ITEM CALIBRATIONS ESTIMATED FROM APRIL 1995, SEPTEMBER 1996, AND APRIL 1997 CPS FOOD SECURITY DATA ..... 20
IV. 1 THRESHOLDS USED TO CLASSIFY HOUSEHOLDS AS "POOR," AND DHHS POVERTY GUIDELINES AT TIME OF SURVEY ..... 26
IV. 2 CONTINUOUS SCALE SCORES BASED ON THE NUMBER OF AFFIRMATIVE ANSWERS ..... 29
B. 1 THRESHOLDS USED TO CLASSIFY HOUSEHOLDS AS "POOR" AND DHHS POVERTY GUIDELINES AT TIME OF SURVEY ..... 82
B. 2 CRITERIA USED IN 1995-1997 SURVEYS TO SCREEN HOUSEHOLDS INTO THE FOOD SECURITY MODULE ..... 84
B. 3 FOOD SECURITY PREVALENCE ESTIMATES WITH ALTERNATIVE DEFINITIONS OF THE ANALYSIS SAMPLE ..... 87
D. 1 PERCENT OF HOUSEHOLDS BY THE MINIMUM NUMBER OF NONMODAL RESPONSES TO THE FOOD SECURITY ITEMS ..... 99
D. 2 ANALYSIS OF MODALITY BY NUMBER OF YES RESPONSES TO FOOD SECURITY ITEMS ..... 100
D. 3 MINIMUM AND MAXIMUM FOOD INSECURITY PREVALENCE ESTIMATES ..... 102
E. 1 THRESHOLD PARAMETER ESTIMATES FOR 18-ITEM 12-MONTH FOOD SECURITY SCALE BY RACE/ETHNIC GROUP ..... 108
E. 2 ITEM ORDERING FOR RACE/ETHNIC GROUPS AND FULL POPULATION ..... 109
E. 3 CROSS-GROUP DIFFERENCES IN THRESHOLD PARAMETER ESTIMATES BY RACE/ETHNIC GROUP ..... 110
Table Page
E. 4 THRESHOLD PARAMETER ESTIMATES FOR 12-MONTH FOOD SECURITY SCALE BY HOUSEHOLD TYPES ..... 114
E. 5 ITEM ORDERING FOR THREE HOUSEHOLD TYPES AND FULL POPULATION ..... 115
E. 6 CROSS-GROUP DIFFERENCES IN THRESHOLD PARAMETER ESTIMATES BY HOUSEHOLD TYPE ..... 116
E. 7 THRESHOLD PARAMETER ESTIMATES FOR 18-ITEM 12 MONTH FOOD SECURITY SCALE BY METROPOLITAN STATUS ..... 119
E. 8 ITEM ORDERING BY METROPOLITAN STATUS AND FULL POPULATION ..... 120
E. 9 CROSS-GROUP DIFFERENCES IN THRESHOLD PARAMETER ESTIMATES BY METROPOLITAN STATUS ..... 121
E. 10 THRESHOLD PARAMETER ESTIMATES FOR 18-ITEM 12-MONTH FOOD SECURITY SCALE BY REGION ..... 123
E. 11 ITEM ORDERING FOR 4 REGIONS AND FULL POPULATION ..... 124
E. 12 CROSS-GROUP DIFFERENCES IN THRESHOLD PARAMETER ESTIMATES BY REGION ..... 125

## FIGURES

Figure Page
A. 1 ITEM CHARACTERISTIC CURVES ..... 57
A. 2 ITEM CHARACTERISTIC CURVES ..... 58
A. 3 TEST CHARACTERISTIC CURVES ..... 60
A. 4 TEST CHARACTERISTIC CURVES ..... 61
A. 5 THREE EQUIVALENT IRT SCALES ..... 64
A. 6 ITEM INFORMATION CURVES ..... 66
A. 7 ITEM INFORMATION CURVES ..... 67
A. 8 TEST INFORMATION CURVE ..... 69
A. 9 TEST INFORMATION CURVE ..... 70
A. 10 18- AND 10-ITEM TEST CHARACTERISTIC CURVES ..... 71
E. 1 ITEM CALIBRATIONS FOR 3 RACE/ETHNIC GROUPS ..... 112
E. 2 ITEM CALIBRATIONS FOR HOUSEHOLD TYPE GROUPS ..... 117
E. 3 ITEM CALIBRATIONS BY METROPOLITAN STATUS ..... 122
E. 4 ITEM CALIBRATIONS BY REGION ..... 126

## EXHIBITS

Exhibit Page
IV. 1 SCREENS USED IN THE CPS SURVEYS . . . . . . . . . . . . . . . . . . . . . . . . . . . . . 27
IV. 2 ALGORITHM FOR IMPUTING MISSING FOOD SECURITY DATA . . . . . . 28

## EXECUTIVE SUMMARY

This is the Final Report for the project, "Analysis of the Current Population Survey Data for Food Security and Hunger Measurement" conducted by Mathematica Policy Research, Inc. (MPR) for the USDA Food and Nutrition Service (FNS), beginning in 1997. The project provided USDA with technical support and statistical estimation work for analyzing the 1996 and 1997 data on food security collected in the U.S. Census Bureau, Current Population Survey (CPS) Food Security Supplement. More broadly, the work examined a number of analytic and empirical issues relevant to analyzing the first three years of CPS food security data available-those for 1995, 1996, and 1997.

It was originally intended that the Final Report would provide the main vehicle for dissemination of the substantive findings on the prevalence of food insecurity based on the 1996 and 1997 data. However, because of the importance of making these results available as early as possible, USDA elected to issue an "Advance Report," thus making the results of the 1996 and 1997 analyses conducted by MPR available before completion of the overall project. In addition, since 1999, a number of publications have become available that present estimates of food insecurity prevalence, as well as discussions of the methods used in computing food security estimates in general. Most important, Andrews et al. (2000) provides a comparative analysis of the annual data for the five-year period 1995 through 1999, while Bickel et al. (2000) provides a how-to guide for measuring food security that incorporates relevant work done prior to that time, including earlier work from the current project. Selected issues in food security are also considered in Ohls et al. 1999.

In light of these developments, USDA suggested that MPR recast this Final Report to focus on several selected topics related to the 1995-1997 data, rather than provide comprehensive treatment of the overall research, much of which has since been incorporated in later publications. The Final Report has been organized around these suggestions.

Among the issues addressed in the report are:

C The stability of the food security measurement scale over time
C Temporal adjustments to the categories or designated ranges of severity on the underlying continuous scale used to classify households by food security status

C Screening issues related to ensuring a strictly comparable analysis sample over the 1995-1997 CPS food security samples

C Alternative imputation strategies for dealing with missing data
C The degree to which household responses to the food security questions are "modal," in the sense that households consistently respond affirmatively to questions involving
less severe food insecurity whenever they respond affirmatively to questions involving relatively more severe food insecurity

C The degree to which the estimated parameters of the model used to measure the severity of food insecurity vary across different groups of households, defined by ethnicity and other characteristics

The first section below provides background information about the analysis. Subsequent sections summarize findings on each of the above issues.

## BACKGROUND

The analysis in this report is based on a statistical procedure which assigns households to food security status, based on their answers to a series of 18 survey questions. The food security categories used are:

C Food secure
C Food insecure-no hunger
C Food insecure with hunger

The data used for national-level analysis of food security are from annual supplements to the CPS, which is fielded monthly to more than 40,000 U.S. households.

Households are classified by food security status in the analysis, based on a procedure, Rasch modeling, which has a long history in the statistical literature. The first work in applying the Rasch model to food security data was undertaken under an earlier contract let by USDA (Hamilton et al. 1997). The Rasch model, as used in that work, posits that there is a single, one-dimensional attribute among households that indicates food insecurity. The model then uses a set of assumptions and statistical methods to assign "severity levels" to each of a series of 18 survey questions relating to food insecurity and hunger. A continuous food security measure is then assigned to each household in the data set, based on households' replies to the 18 questions. Supplemental procedures developed by Hamilton et al. are then used to translate the continuous scale score into a limited number of discrete food security statuses.

The objectives of the current project were to extend the analysis to 1996 and 1997 data and to address a number of related issues associated with measuring food security over time. Our findings in selected areas are summarized below.

## STABILITY OF THE PARAMETERS OF THE MODEL OVER TIME

An important issue in examining the validity of the Rasch modeling approach is whether the model parameter estimates are stable over time. The underlying theory on which the Rasch model is based posits that, if the wording of an item does not change, its estimated level of severity should not change. For example, even if food insecurity became more prevalent over time, a household at a given level of insecurity this year is expected to answer each item the same way a household at that level of insecurity did a year earlier. Due to sampling variability and other factors, such as minor wording changes, we do not expect estimated model parameters to remain exactly the same over time; but a finding of major changes over time would call into question the validity of the model. Particularly problematic would be a finding of important changes in the ordering of the items by severity.

To examine issues of model stability, we estimated the model independently on three CPS data sets (1995-1997), using consistent conventions as to statistical scaling. Some variation across years was found, as expected. In general, however, the estimated parameters of the model were quite stable. Also, the estimated order of severity of the different questions remained largely constant, with the only changes in severity order occurring among questions that were very close to each other on the severity scale in the original estimation work. The conclusion of this component of the research is that the food security model is sufficiently temporally stable to make it a reasonable tool to use in time series analysis.

## ADJUSTING "CUT POINTS" USED TO CLASSIFY HOUSEHOLDS INTO A LIMITED NUMBER OF FOOD SECURITY STATUS CATEGORIES

The Rasch model places each household on a continuous numeric food security scale. For purposes of policy analysis, it is also useful to establish numerical "cut points" that assign households to a small number of designated categories which summarize their food security status. To create this categorical measure, Hamilton et al. (1997a and 1997b) specified four categories: food secure, food insecure without hunger, food insecure with moderate hunger, and food insecure with severe hunger. More recently, the latter two categories have usually been collapsed to form a single category, while additional scale development work has identified a new nested category, food insecure with children's hunger (Nord and Bickel 1999 and 2001).

A key issue that arises in this work is whether it is appropriate to keep the same continuous scale cut points over time, or whether, alternatively, some temporal adjustments may be needed. The analysis of the body of the report concludes that, at least in some situations, it is not optimal to attempt to classify households based on the same cutpoints over time.

While the Rasch model places households on a continuous food security scale, due to certain statistical properties of the model substantial numbers of households tend to be clustered at certain points in the scale. If cut points are held constant, there is a risk that, because of chance statistical variation, the score assigned to one of these clusters of households might accidentally cross one of the cutpoints in a given year, causing considerable instability in estimates of food security prevalence.

Chapter V of the Final Report identifies several technical approaches for avoiding this difficulty. The discussion is based on the principle that a household with a given pattern of survey answers should always be classified into the same food security grouping, independent of when the data are collected.

## SCREENING HOUSEHOLDS INTO THE SAMPLE IN THE 1995-1997 SURVEYS

The food security supplements in the 1995-1997 CPS had two general sections. The first section gathered information about food expenditures, participation in several programs aimed at providing food to needy families (for example, food stamps and school meal programs), and the sufficiency of food eaten during the preceding 12 months. The second section gathered more-detailed information about food insecurity and coping behaviors during the previous 12 months and prior 30 days. Not all households were asked this second set of questions, which includes the questions used to construct the food security scale. In order to minimize respondent burden, households who, on the basis of earlier questions, appeared to have a high likelihood of being food secure were excluded from the more detailed questions and were assumed to be food secure in the analysis. This prescreening applied to higher income households in all three years, 1995-1997, and in one year, 1996, it was applied to lower-income households as well. Beginning in 1998 and continuing consistently since then, the CPS Food Security Supplement has included a new, less restrictive, pre-screen applied to higher-income households.

To ensure comparability in the analysis samples for the three years, the current research developed a common screen, such that any households giving survey answers that passed this common screen would have been tracked into the food security module in any of the three years. Households that did not pass the common screen were, for purposes of the analysis, treated as if they had not been tracked into the food security module of the survey-essentially, they were assumed to be food secure. Technical details concerning how this common screen was constructed are provided in Appendix B of the Final Report.

While use of the common screen has the desired effect of ensuring consistency in the 1995-1997 analysis samples, it also has the effect of treating as food secure a number of households who, during the survey, gave indications of experiencing food insecurity. Across the three years, use of the common screen was found to result in estimates of the prevalence of food insecurity which are between 1.0 and 1.5 percentage points lower than those that are obtained when the maximum available samples are used in the estimation. ${ }^{1}$

[^0]
## IMPUTING MISSING DATA

Most households gave complete answers to the food security questions they were asked in the CPS; however, a limited number did not. Appendix C of the report examines a number of alternative approaches for including households with partially missing data in the analysis. One approach is reliance on the Rasch model itself, which has the capacity to assign food security scale scores to observations with incomplete data. However, as is noted in Appendix C, in some instances, the determinations made within the model for cases with substantial amounts of missing data may lack face validity. Also, as a practical matter, many researchers may not have easy access to the software needed to implement the model.

An alternative algorithm for dealing with missing data has therefore been developed. Depending on the exact configuration of food security module answers given by the respondent, this alternative algorithm essentially involves imputing missing data items based on either (a) the highest severity item, in terms of level of food security severity, that the respondent answered positively; or (b) the lowest severity item answered negatively.

## "MODALITY" OF HOUSEHOLD FOOD SECURITY RESPONSE PATTERNS

The Rasch model implies that many households will exhibit item response patterns that are reasonably "modal" in the sense that if a household answers "yes" to any of the items, it will tend to answer "yes" to the less severe items, then answer "no" to the more severe items. A household that exhibits this pattern exactly-a string of all "yes" answers followed by a string of all "no" answers-is said to be a "modal" household. There is nothing in Rasch theory that predicts that all households will be modal; indeed, the model cannot be estimated if all households are exactly modal. Still, it is of interest in understanding the data to examine the degree of modality present. A large number of strongly nonmodal response patterns could call into question the validity of the model.

Analysis of the 1997 data indicates that most household response patterns tend to be either exactly or approximately modal. Of those households in the 1997 data who gave an affirmative answer to at least one question, approximately 39 percent households provided answer patterns that were exactly modal, while another 36 percent gave sets of answers which had only a single nonmodal response.

## CONSISTENCY OF ESTIMATED FOOD SECURITY MODEL PARAMETERS ACROSS POPULATION SUBGROUPS

Essentially, the analysis conducted with the aggregated CPS data sets assumes that different subgroups of the population are similar with regard to how they experience food insecurity. To test this assumption, the Rasch model was estimated separately for subgroups of the population, defined according to (a) race/ethnicity; (b) household composition; (c) metropolitan status; and (d) region of country.

The results of this analysis indicate considerable robustness of the analysis to this kind of disaggregation. In general, estimated severity levels for the individual questions were found to have consistent patterns across different subgroups, and the magnitudes of the parameters do not change substantially.

There is no clear statistical test of how much difference in the estimated subpopulation models would affect confidence in the overall modeling approach. However, the judgment of statistical experts who have used the Rasch model extensively in other contexts is that the findings of the subgroup analyses can reasonably be judged to be highly consistent with one another.

## CONCLUSION: REFLECTIONS ON THE STRENGTHS AND LIMITATIONS OF THE FOOD SECURITY METHODOLOGY

We conclude by discussing the strengths and limitations of the use of the Rasch model as a basis for food security measurement. Possible directions for future research are also noted.

The food security scale reflects more than 10 years of methodological development by both government and private groups. The use of the Rasch model methodology has made it possible to guide the development of the food security estimates with a thoroughly studied model that has wellunderstood statistical properties. In terms of goodness-of-fit criteria, the mathematical form of the measurement model shows strong correspondence on "fit" to the empirical data. The approach has undergone extensive review by experts in both the public and the private sector. In general, these experts consider the model an appropriate application of the IRT methodology, and they have viewed the analysis results as reasonable.

Another important strength, as established by the current project, is that the estimated item parameters of the IRT model are robust across time and population subgroups. The values obtained from the 1996 and 1997 data are essentially the same as the original 1995 values. ${ }^{2}$ In addition to stability over time, there is stability across subgroups, defined by such characteristics as race/ethnicity, household composition, and region of the country.

Tempering these strengths are a number of limitations which should also be recognized. Most of these, if not all, are a matter of careful interpretation of what the food security measure does and does not do. For example, the CPS indicator questions for food insecurity and hunger and the scale developed from them are designed to provide a household-level measure of the severity of conditions as experienced within U.S. households. This is in line with the conceptual understanding of food insecurity as a condition of deprivation or stress experienced by households in meeting members' basic food needs. However, the experience of hunger as such, which appears only at a more severe stage of food insecurity, is strictly individual. The household classification, "food insecure with hunger" refers to that more severe range where evidence of reduced food intake and hunger has appeared for one or more household members. But this is a collective measure which may apply to all household members, to adult members only, or to as few as one (adult) member.

[^1]Second, the basic measure is designed to capture respondents' experiences over the course of a year, while household circumstances can change markedly during such a period. Accordingly, the 12-month measure-designed to provide reliable benchmark and trend figures-may not represent the current situation of given households. Similarly, the "food insecure with hunger"designation can, in principle, result from just one serious episode during the year, although for most such households evidence of a repeated pattern of reduced (adult) food intakes during the year must be established.

In addition, a number of issues of interpretation flow from the need to have a simple categorical measure as a means of classifying households for purposes of manageable data reporting and monitoring, in addition to the underlying continuous scaled measure. The categorical measure was created to make the scale more accessible to non-technical users and more convenient to users whose needs could be better served by a simple categorical variable than by the detailed continuous measure. The categorical measure as such is straightforward: it represents designated ranges of severity along the continuous scale (i.e., qualitatively differing severity levels of "food insecure"), plus the category of households that either show no evidence of food problems within the CPS data set, and hence can be deemed to be "food secure," or that show only one or two indications of food stress, which is deemed insufficient as evidence to establish confidently their status as "food insecure."

The interpretive problems with the categorical measure stem from at least three sources. First, the designation of appropriate severity ranges, and their exact delineation in operational form based on the available set of indicators, is inherently judgmental and thus leaves room for disagreement.

Second, the Rasch model employs a probabilistic logic in generating the continuous scale measure of severity of household food insecurity; similarly, the corresponding severity-range summary categories share this probabilistic nature. However, the naming conventions adopted for the severity-range categories are determinate in form, which can be misleading.

Thirdly, a misplaced specificity and determinateness can easily be attributed to the individual indicator items as well, causing a misunderstanding of their actual role in the measurement process.

To illustrate this last point, straightforward names adopted for the severity-range categories raise issues of face validity when they seemingly contradict the clear language of particular indicator items embedded within the measurement scale. For instance, it is technically possible for a household to be classified "food insecure with hunger," even though the respondent has answered "no" to the particular question, "In the last 12 months were you ever hungry but didn't eat because you couldn't afford enough food?" In this case, the respondent either must have replied "yes" to a series of increasingly severe indicators of food insufficiency, including at least three items indicating reduced food intake for themselves and/or other adult members of the household, one of which establishes a repeated pattern of such reduced intakes over the year, or they must have replied "yes" to most of the foregoing, plus one or more of the items that are more severe than the explicit hunger question. The categorical measure (and its naming convention) reflects the judgment that, on the balance of this evidence, one or more adult members of the household has, with high probability, experienced resource-constrained hunger sometime during the year. Conversely, the opposite case also can occur: the household can be classified "food insecure without hunger," based on its overall pattern of
response and the resulting scale score, even though the respondent has answered "yes" to the explicit hunger question as such.

In creating the scale, a number of steps were taken to minimize the effects of these factors. For instance the numerical cutpoints defining the categories were set to be conservative, in the sense that there must be three answers to questions thought to indicate food insecurity before a household is classified as food insecure, and similarly for the hunger classification. Also, analysis presented in the text of the report indicates that substantial numbers of respondents follow close-to-expected, response patterns, which do not lead to any apparent anomalies in classification. Nevertheless, room for disagreement remains as to what types of answers to the questions should be construed as reflecting the language used in designating the three scale categories.

A possible solution to some of these issues would be to state the category names in more-probabilistic terms, such as "probably food insecure" or "a high likelihood of hunger." This would be in keeping with the probabilistic nature of the underlying model, and it might help ease the concerns of those who are bothered by the anomalies posed by apparently inconsistent patterns of question responses. However, such category name changes might also interfere with the clarity of meaning of the categories themselves, thus reducing their effectiveness.

Overall, it is important to recognize that these limitations have not prevented the food security scale from becoming an important, widely used research and policy tool. Questions to support the scale have been included in an increasing number of national surveys and scale results are frequently cited in the policy process. This evidence suggests that many policy analysts have found the scale to be a valuable tool for measuring an important aspect of material deprivation among America's poor.

## I. INTRODUCTION

For decades, ensuring that everyone in the United States has effective access to adequate, nutritious food has been a major policy goal of the federal government. More recently, increased attention has been paid to developing ways of measuring how well this goal is being achieved. A growing body of literature has emerged from work done in the late 1980s, aimed at measuring households' "food security." As part of this literature, the U.S. Department of Agriculture (USDA) has begun publishing estimated rates of food security and food insecurity in the United States. The research reported below was undertaken to continue this work.

## A. PURPOSE OF THE REPORT

This is the Final Report for the project, "Analysis of the Current Population Survey Data for Food Security and Hunger Measurement" conducted by Mathematica Policy Research, Inc. (MPR) for the USDA Food and Nutrition Service (FNS) beginning in 1997. The project provided USDA with technical support and statistical estimation work for analyzing the 1996 and 1997 data on food security collected in the annual Food Security Supplement to the U.S. Census Bureau, Current Population Survey (CPS). More broadly, the work examined a number of analytic and empirical issues relevant to analyzing the first three years of CPS food security data available-those for 1995, 1996, and 1997.

It was originally intended that the Final Report would provide the vehicle for the main dissemination of the substantive analytical findings on the prevalence of food insecurity based on the 1996 and 1997 data. However, because of the importance of making these results available as early as possible, USDA eventually elected to issue an "Advance Report," thus making the results of the 1996 and 1997 analyses conducted by MPR available before completion of the overall project.

The Advance Report also included comparable results based on 1998 data. In addition, since 1999, a number of other publications have become available which present estimates of food insecurity prevalence, as well as discussions of the methods used in computing food security estimates in general. Most important, Andrews et al. (2000) provide a comparative analysis of the annual data for the five-year period 1995 through 1999, while Bickel et al. (2000) provide a "how to" guide for measuring food security that incorporates relevant work done prior to that time, including earlier work from the current project. Other important contributions include Ohls et al. (1999) and Prell and Andrews (2001).

In light of these developments, USDA suggested that MPR recast this final report to focus on several selected topics related to the 1995-1997 data, rather than providing a comprehensive treatment of the overall research, much of which has either been reported previously or incorporated in later publications. This report has been organized around these suggestions.

## B. OVERVIEW OF REPORT

Chapter II provides a brief summary of the past work in the food security area and of the methods currently being used to measure food security; it thus provides a context for the subsequent material. Chapter III presents results that describe the stability of the estimated food security scale across different years. Chapter IV provides a summary of "benchmark" methods that can be used to estimate the food security scale on different data sets and with different assumptions as to which parameters are to be held constant and which are not. Chapter V discusses issues related to adjusting the food security model over time, to accommodate both the availability of new data and possible developments in the underlying methodology. Chapter VI provides some broad conclusions about the food security scale.

Appendices provide new tabulations of the data and explore additional methodological issues, such as:

- Summary of salient parts of item response theory and related measurement issues
- Screening households into the food security analysis
- Imputing missing food security data
- How "modal" are household response patterns?
- Item-severity levels, by subgroup


## II. BACKGROUND

The publication of the U.S. Department of Agriculture (USDA) 1997 report on food security levels in the United States (Hamilton et al. 1997a and 1997b; and Price et al. 1997) has spurred widespread interest in measuring food security for various groups in the U.S. population. Using data from the April 1995 Current Population Survey (CPS), that set of reports presented a comprehensive method for measuring food security levels. Other major surveys that have measured food security or plan to do so are the Panel Study of Income Dynamics, the National Health and Nutrition Examination Survey, the Continuing Survey of Food Intakes by Individuals, the Survey of Program Dynamics, and the Early Childhood Longitudinal Study.

The 1997 USDA report was based on a single CPS data set, that for April 1995. Important next steps in food security research have been to extend that analysis to later years, and to develop a method for measuring changes in food security over time. Major research questions include:

- Are estimated model parameters stable over time?
- What method is most appropriate for assessing changes in the prevalence of food insecurity in the U.S. population?
- How sensitive are prevalence estimates to alternative ways of implementing the procedures used in the 1997 report?

This chapter gives background information that will be useful in addressing these issues. Section A provides background for our work, describing the roots of the food security concept in prior research. The CPS data used for the current study are described in Section B, while Section C describes the Rasch model used in both the original analysis of the 1995 data and the current work measuring the severity of food insecurity as experienced within households.

## A. A BRIEF SUMMARY OF THE LITERATURE INFORMING THE FOOD SECURITY CONCEPT

Although hunger has long been a concern of American social and nutrition policy, attempts to measure it systematically have posed major challenges to advocates and policy analysts alike. Early attempts to equate hunger directly to malnutrition were not successful, because they encountered conceptual difficulties in defining malnutrition and operational difficulties in developing reliable and inexpensive ways of measuring people's nutrient intake. Furthermore, as additional discussion took place, it was recognized that feeling physical "hunger," a sensation experienced by most people fairly frequently, is not equivalent to the experience of hunger resulting from lack of resources to obtain food, a situation more closely related to economic deprivation. Obviously, refinement of the concept was needed.

From the late 1970s through the early 1980s, there was growing interest in broadening the concept of hunger to the more general construct of resource-constrained food insecurity. This broader concept came to be defined in terms of the phenomena and experiences associated with food stress for the households, as well as household members actually experiencing hunger. Lacking access to food because of resource constraints also came to be included in most analysts' definition of "hunger" as a policy issue.

The broadening of the relevant concepts took place partly within the U.S. government, with the inclusion of a basic question related to food insecurity in the two most recent administrations of the Nationwide Food Consumption Survey and a small set of such questions in the Third National Health and Nutrition Examination Survey and the Survey of Income and Program Participation. Two private research efforts gave substantial impetus to the evolving focus on food insecurity. First, the Community Childhood Hunger Identification Project (CCHIP)—organized by the Food Research and Action Center and funded by local and national business and philanthropic
organizations-demonstrated that reasonable and consistent answers could be obtained using a set of survey questions designed to measure food insecurity (Wehler et al. 1991). Second, work at Cornell University provided additional theoretical support and advanced the development of measurement scales based on answers to survey questions about food security (Radimer et al. 1992).

Beginning in 1992, a federal interagency working group on food security measurement-under the leadership of the USDA Food and Nutrition Service (FNS) and the DHHS Centers for Disease Control and Prevention, National Center for Health Statistics (NCHS)—began a systematic effort to develop a battery of questions about food insecurity based on the prior research, which could be administered regularly in government-conducted surveys. Drawing on previous research findings about food insecurity, together with additional research commissioned from outside researchers, USDA staff assembled the full range of food security survey questions that had been used, and identified sets of items that held promise as reliable indicators for use with U.S. state- or nationallevel populations. The Federal Interagency Group was assisted in this work by an expert panel that included many leading food security researchers.

FNS passed an important milestone when it won approval from the U.S. Office of Management and Budget (OMB) for a supplement to the April 1995 CPS containing a set of questions designed to measure food security. The supplement gathered information about households' shopping patterns and various aspects of food insufficiency and insecurity during the 30 days and 12 months prior to the interview.

In 1995, Abt Associates, assisted by staff from Tufts University and Cornell University and CAW and Associates, Inc., was engaged by the USDA to analyze the 1995 CPS data. Faced with a questionnaire containing more than 50 items, the Abt team worked with the USDA to further refine the underlying concepts of food insufficiency and food insecurity. Along with this conceptual work,
the team had to identify which of the CPS questionnaire items could reliably measure food insecurity. In the early stages of their work, they relied heavily on factor analysis to identify a group of items that, taken together, appeared to measure food insecurity. Next, the Abt team applied a scaling procedure (described later) to assign a food security measure to each household. Based on these measures, households were classified into four categories-food secure; food insecure without hunger; food insecure with moderate hunger; and food insecure with severe hunger-and the Abt team was able to estimate the prevalence of these four designated levels of food security/insecurity.

Table II. 1 summarizes the published food insecurity prevalence estimates for the five years for which data have been analyzed. As shown in the table, depending on the year of estimation, approximately 89 to 91 percent of U.S. households are estimated to be food secure. The rate of food insecurity with hunger is estimated to be in the range of 3 to 4 percent. The remaining households, approximately 5 to 7 percent, are in the less severe range where there are signs of food insecurity but the households are not classified as hungry.

This report extends the research of the CCHIP, Cornell, USDA, and Abt researchers. The work done by the Abt team was focused on developing and implementing a measure of the severity of food insecurity using data from the April 1995 CPS. The present report analyzes data from two additional rounds of surveys: the September 1996 and April 1997 CPS. ${ }^{1}$ With these additional data, the focus shifts to issues that arise in the development of a stable and consistent ongoing social indicator. The issues examined here are of the kind that can be critical when prevalence is measured on a routine basis and changes in prevalence are closely monitored by policymakers. The availability of food-

[^2]TABLE II. 1
PREVALENCE OF FOOD SECURITY FOR HOUSEHOLDS, BY YEAR (Percentage)

|  |  | Food Insecure |  |  |
| :--- | :---: | ---: | :---: | :---: |
| Year | Food Secure | All | Without Hunger | With Hunger |
| 1995 | 89.7 | 10.3 | 6.4 | 3.9 |
| 1996 | 89.6 | 10.4 | 6.3 | 4.1 |
| 1997 | 91.3 | 8.7 | 5.6 | 3.1 |
| 1998 | 89.8 | 10.2 | 6.6 | 3.6 |
| 1999 | 91.3 | 8.7 | 5.9 | 2.8 |

Source: Andrews et al. 2000.
security data from three years in sequence has made it possible to address issues that arise when tracking changes over time.

## B. THE CPS DATA ON FOOD SECURITY

Data for the current study come from the USDA-sponsored Food Security Supplement to the CPS. The CPS is a monthly survey of about 50,000 households conducted by the U.S. Census Bureau for the Bureau of Labor Statistics. The sample is designed to represent the civilian, noninstitutional population of the United States. Each monthly sample is divided into eight representative subsamples, or rotation groups. A given rotation group is interviewed for a total of eight months: it is in the sample for 4 consecutive months, leaves the sample during the following 8 months, then returns for another 4 consecutive months. In each monthly sample, one of the eight rotation groups is in the first month, another rotation group is in the second month, and so on. ${ }^{2}$ Under this system, 75 percent of the sample is common from month to month, and 50 percent is common from year to year for the same month.

The primary purpose of the CPS is to provide information about the labor force characteristics of the U.S. population. In each month, however, a supplement is added to the core questionnaire. In March of each year, for instance, the Census Bureau sponsors the Annual Demographic Supplement. This survey is the data source for the official income and poverty statistics published by the Census Bureau each year. In April 1995, September 1996, April 1997, August 1998, April 1999, September 2000, and April 2001, a special supplement was added to the CPS core
${ }^{2}$ More formally, the CPS sample is actually one of geographic addresses rather than households. If sample members move to a new address, they are not interviewed at that new address and they thus leave the sample. However, the address those sample members moved from remains in the sample, and the new residents are interviewed. These are known as "replacement households."
questionnaire which included questions about household food sufficiency, food security, food expenditures, food program participation, and several other related items. The structure of the food security supplement used in these surveys was as follows:

1. Food expenditures during the prior week
2. Participation in food assistance programs (food stamps, meal programs for the elderly, school meal programs, and the Special Supplemental Food Program for Women, Infants, and Children)
3. Food insufficiency during the prior 12 months and ways of coping with that insufficiency
4. Food security and hunger indicator questions for the prior 12 months and the prior 30 days

Not all households were asked the full set of questions in the supplement. To minimize respondent burden, a set of preliminary screening questions was used in some cases to determine whether there was evidence that a household might have experienced food insecurity. If there was no such evidence, the subsequent food security questions were skipped. This preliminary screen has been applied to higher-income households in every year and was applied to lower-income households as well in one year (1996). Across the three CPS samples in 1995, 1996, and 1997, different screening procedures were used. Beginning in 1998, the structure of the food security questionnaire was redesigned in a more fundamental way (but with content unchanged), including a basic change in screening procedure, now expected to remain constant.

Table II. 2 shows unweighted sample sizes for the three CPS samples used here. The initial sample size for the April 1995 CPS was 53,665 households. Budget cuts in January 1996 resulted in reduced sample sizes, which are reflected in the initial sample sizes for September 1996 and April 1997 shown in Table II.2. The initial sample for September 1996 was 47,795; for April 1997, it was

TABLE II. 2
CPS SAMPLE SIZES
Unweighted Number of Households

|  | April <br> 1995 | Sept. <br> 1996 | April <br> 1997 |
| :--- | ---: | ---: | ---: |
| Full CPS | 53,665 | 47,795 | 47,306 |
| Households in Supplement |  |  |  |
|  | 44,730 | 41,811 | 41,146 |
| Households Tracked Into Food Security Module |  |  |  |
| Answered all key questions asked $^{\mathrm{a}}$ | 18,453 | 10,957 | 11,175 |
| Answered at least half of key questions asked, but not all $^{\mathrm{a}}$ | 18,179 | 10,685 | 10,937 |
| Answered fewer than half of key questions asked $^{\mathrm{a}}$ | 195 | 203 | 171 |

${ }^{\text {a }}$ There are 18 key questions for households with children and 10 for those without children.

47,306. In all three samples, roughly 85 percent of the core households entered the Food Security Supplement. ${ }^{3}$ Of those households, about 40 percent of the April 1995 sample passed the screening questions and were asked the balance of the Food Security Supplement. For the 1996 and 1997 surveys, tighter screening procedures resulted in only about 26 percent of households being asked the rest of questions in the Food Security Supplement. In all cases, there is a presumption that households failing to pass the screen are food secure. The differences in the screening procedures used across these three samples have important implications for the consistent measurement of food insecurity during the 1995-1997 period, a necessary prerequisite for measuring the changes in food insecurity across those years. These issues are discussed in more detail in Appendix B.

Most of the research reported below is based on 18 key questions (items) that are used for the measurement of household food insecurity. Households with one or more children are asked all these questions; childless households are asked only the 10 items that do not pertain to children. Once the differences between households with and without children were taken into account, there was little item nonresponse. In all three samples, more than 97 percent of the households that passed the initial screen responded to all the items used to measure food insecurity that they were asked. However, the fact that childless households responded to only 10 of the 18 items asked of households with children presents an additional complication, as discussed below.

## C. THE ONE-PARAMETER LOGISTIC ITEM RESPONSE THEORY (RASCH) MODEL

The Food Security Supplement to the April 1995 CPS contained more than 65 questions, of which about 50 were potential indicators of food insufficiency, insecurity, or actual hunger and thus

[^3]were candidates for inclusion in the measurement scale. Of the entire question set, four questions were used as a preliminary screen to identify households that showed no indication of food insecurity during the prior 12 months, and thus were not to be burdened with additional questions. Of the remaining items, 18 are used directly to measure households' food insecurity levels over the prior 12 months. (Of the questions not used in the measurement scale, some apply only to the 30 days prior to the interviews; others were found during preliminary analysis not useful in developing the full food insecurity scale.)

An important objective achieved in the first round of research on the 1995 data was to develop a method for combining answers on the 18 items into a single scale measuring the level of severity of household food insecurity as experienced within the U.S. population. In doing this, it was necessary to take the following factors into account:

- Not all questions apply to all households; in particular, eight of the questions are not relevant to households that do not have children
- The data included some item nonresponse involving households that did not answer all the questions asked.

In developing the desired food security scale, the researchers involved drew heavily on a rich body of procedures developed originally in the educational testing literature called "Item Response Theory" (IRT). IRT methods have been widely used in educational contexts, such as the Scholastic Aptitude Tests (SAT) and the National Assessment of Educational Progress, to measure student attributes (such as math ability), using tests which, for test security reasons and other factors, are not identical. ${ }^{4}$ In applying IRT methods to the food security measurement context discussed in the
${ }^{4}$ For summaries of IRT theory, see Hambleton (1993) and Wright and Masters (1982).
present report, the attribute being examined is food insecurity and the test items are the individual food security questions on the CPS supplement.

The methods used in the original analysis of the 1995 food security data, which are also applied in the present report, involve a closely related technique called "Rasch modeling." ${ }^{5}$ The salient characteristic of the Rasch model is that the model involves estimating only a single parameter, often called the "severity" level, which is used to characterize each question on the scale. Other versions of IRT theory estimate either two or three parameters per question.

Appendix A provides a more detailed summary of the Rasch model. We conclude this section by noting certain salient properties of Rasch models that are relevant to the following discussion:

- The scale measure for households with complete data can be calculated based only on the number of questions about food insecurity which they answer affirmatively.
- The scale measure determined by the model is unique up to a linear transformation; once a scale is developed, any linear transformation of the scale conveys the same information.
- In a Rasch model, each household's level of food security, and each item's level of severity, are determined simultaneously within the model.


## D. RELATIONSHIP BETWEEN IRT MODELING AND THE FOOD INSECURITY PREVALENCE ESTIMATES

As noted above, this report covers a number of different topics in food security measurement, some of which are related principally to the IRT-based model estimation and some of which relate more directly to food security prevalence estimates. As a context for these discussions, it is useful to provide an overview of the relationship between the model and the prevalence estimates. The

[^4]IRT-based model played several important roles in the development of the food security scale currently being used. First, it helped establish the order of severity for different individual questions in the CPS module, thus facilitating the work done by the initial research team in identifying ranges of responses corresponding to various food security levels. Second, the IRT-based model provided a formal way to calibrate food security levels for households without children on the same scale as households with children, even though some of the CPS questions were not applicable to the former group.

Third, the model provided a mechanism for dealing with item non-response on the survey (though this was of limited importance, given that, as discussed later in the report, item nonresponse was quite low). Fourth, the IRT model has provided a way of partially testing the stability of the index over time and across subgroups. In addition, the IRT model may be useful in the future for continuing to examine the stability of the processes which determine levels of food security.

These uses of the IRT-based modeling having been stated, it may be noted that repeated application of the measurement model is not needed for the ongoing measurement and monitoring of food security, or for the measurement of food security on the same basis in local areas. As summarized earlier, now that the model estimation work has been done, and the underlying measurement scale created the actual scale measure for a specific household with complete data on the food security questions can be calculated simply by counting the number of affirmative answers to the questions; no further estimation work is needed. This greatly simplifies the ongoing work of making annual food insecurity prevalence estimates.

## E. CONVENTIONS USED IN THIS REPORT

Because the scaling of parameters can be changed using any linear transformation without losing meaning, a convention is needed as to scaling when comparing parameters estimated with different data sets. The convention used in this report, unless otherwise noted, is that the mean of the item parameters is set to 7 and the slopes of the item response curves at their inflection points (see Appendix A) are equal to 1.

Also, because different sample screening procedures were used in different years of the CPS, a convention is needed as to how to make the data sets comparable across years. Except when otherwise noted, the convention used is to define the sample using the "common screen" as described in Appendix B.

## III. STABILITY OF ESTIMATED MODEL PARAMETERS OVER TIME

An important issue in examining the validity of the Rasch modeling approach is whether the model parameter estimates are stable over time. The underlying theory on which the Rasch model is based posits that, if the wording of an item does not change, its estimated severity level should not change. For example, even if food insecurity became more prevalent over time, a household at a given level of severity this year is assumed to answer each item the same way as a household at that same level of severity did a year earlier. Due to sampling variability and other factors, such as minor wording changes, we do not expect estimated model parameters to remain exactly the same over time; but a finding of major changes over time would call into question the validity of the model. Particularly problematic would be a finding of important changes in the ordering of the items by severity.

Table III. 1 shows estimated item parameters based on separate estimation for each of the three years. For each year, only data from that year are used in the estimation. The estimates use the common screen as described in Appendix B. Each set of item parameters has been scaled in the standard way used in this report, with the mean set at 7 and the slopes of the item characteristics curves at their inflection points set at 1 .

Of particular interest in our analysis was whether the severity ranking of the items remained reasonably constant. This was of interest, both as a general indication of parameter stability and because the item order had been drawn on substantially in the original Hamilton et al. (1997) work in designing the algorithm that translates the continuous Rasch model into the designated categorical household food security levels. (See Chapter V.)
TABLE III. 1

| Item |  | 1995 |  | 1996 |  | 1997 |  | Change (1995 to 1997) |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Standard |  | Standard |  | Standard |  | Standar |  |  |
|  |  | Parameter | Error | Parameter | Error | Parameter | Error | Parameter | Error | t | $\operatorname{Pr}(\|T\|>\|t\|)$ |
| 50 | Child not eat whole day | 11.463 | 0.220 | 11.167 | 0.192 | 11.284 | 0.245 | -0.18 | 0.330 | -0.543 | 0.587 |
| 44 | Child skipped meal, 3+ months | 10.206 | 0.106 | 10.168 | 0.208 | 9.873 | 0.187 | -0.33 | 0.215 | -1.550 | 0.121 |
| 43 | Child skipped meal | 9.662 | 0.104 | 9.687 | 0.155 | 9.293 | 0.170 | -0.37 | 0.200 | -1.849 | 0.065 |
| 29 | Adult not eat whole day, 3+ months | 8.999 | 0.086 | 9.043 | 0.059 | 8.999 | 0.073 | 0.00 | 0.113 | 0.004 | 0.997 |
| 47 | Child hungry | 8.641 | 0.083 | 8.995 | 0.069 | 8.957 | 0.111 | 0.32 | 0.139 | 2.274 | 0.023 |
| 28 | Adult not eat whole day | 8.448 | 0.074 | 8.452 | 0.070 | 8.516 | 0.072 | 0.07 | 0.103 | 0.660 | 0.509 |
| 40 | Cut size of child's meals | 8.386 | 0.116 | 8.587 | 0.102 | 8.564 | 0.087 | 0.18 | 0.145 | 1.231 | 0.219 |
| 38 | Adult lost weight | 8.294 | 0.049 | 8.158 | 0.095 | 8.168 | 0.087 | -0.13 | 0.100 | -1.267 | 0.206 |
| 35 | Adult hungry but didn't eat | 7.227 | 0.042 | 7.227 | 0.079 | 7.297 | 0.060 | 0.070 | 0.074 | 0.952 | 0.341 |
| 57 | Child not eating enough | 6.954 | 0.096 | 7.022 | 0.045 | 6.915 | 0.094 | -0.04 | 0.134 | -0.288 | 0.774 |
| 25 | Adult cut size or skipped meals, in 3+ months | 6.388 | 0.042 | 6.416 | 0.036 | 6.463 | 0.054 | 0.08 | 0.068 | 1.099 | 0.272 |
| 32 | Adult eat less than felt they should | 5.643 | 0.040 | 5.614 | 0.053 | 5.663 | 0.036 | 0.02 | 0.054 | 0.365 | 0.715 |
| 56 | Couldn't feed child balanced meals | 5.610 | 0.050 | 5.613 | 0.056 | 5.818 | 0.075 | 0.21 | 0.090 | 2.310 | 0.021 |
| 24 | Adult cut size or skipped meals | 5.590 | 0.046 | 5.505 | 0.040 | 5.586 | 0.047 | 0.00 | 0.065 | -0.063 | 0.950 |
| 58 | Adult fed child few low cost foods | 4.334 | 0.059 | 4.361 | 0.056 | 4.522 | 0.057 | 0.19 | 0.081 | 2.309 | 0.021 |
| 55 | Adult not eat balanced meals | 3.977 | 0.042 | 3.940 | 0.055 | 4.079 | 0.066 | 0.10 | 0.078 | 1.300 | 0.194 |
| 54 | Food bought didn't last | 3.719 | 0.048 | 3.686 | 0.039 | 3.700 | 0.070 | -0.02 | 0.085 | -0.226 | 0.821 |
| 53 Worried food would run out |  | 2.461 | 0.081 | 2.363 | 0.052 | 2.306 | 0.087 | -0.16 | 0.119 | -1.307 | 0.192 |
| Mean |  | 7.000 |  | 7.000 |  | 7.000 |  | 0.000 |  |  |  |
| Std Dev |  | 2.467 |  | 2.474 |  | 2.413 |  | 0.181 |  |  |  |
| Sample size |  | 6,507 |  | 6,203 |  | 5,429 |  |  |  |  |  |

The items in Table III. 1 are listed in their 1995 estimated order of severity. Inspection of Table III. 1 reveals one "inversion" of adjacent items in the 1995 ordering when the model is estimated with 1996 data and two in the 1997 results. In the 1996 estimates, the estimated severity level of Item 40 rises considerably, while the level of Item 28 does not, causing Item 40 to move ahead of Item 28 in the severity ranking. The same inversion happens in 1997, as well as one between Items 32 and 56, where Item 56 is estimated to be less severe than Item 32 in 1995, but somewhat more severe in 1997. In assessing the significance of these changes in item ordering within each of the two pairs of items, it is important to note that each pair of items began (in the 1995 analysis) very close together. The distance on the severity scale separating Items 28 and 40 was less than .10 in 1995 ( 8.448 versus 8.386 ), on a scale that spans approximately 9 units. This is smaller than the standard error of Item 40, by itself. Items 32 and 56 were even closer together in 1995 (5.643 versus 5.610). Thus, the inversions that occurred could easily be due to statistical sampling error.

Furthermore, as shown in the table, not only is the item order quite stable, the magnitudes of the estimated item parameters do not change substantially. While there is some change over time, the degree of variation is small, relative to the overall range of the scale. For instance, the severity of the most severe (and least precisely estimated) item, Item 50, fluctuates only relatively slightly. From a value of 11.463 in 1995, it drops to 11.167 in 1996, then rises to 11.284 in 1997. Fluctuations are greater for some other items. The largest difference over time is for Item 43, which changes by .37 units 2.65 to 2.29 between 1995 and 1997 .

While the fluctuations in parameters are relatively small, three are statistically significant with a 95 percent level test. They are Q47, Q56, and Q58, each of which has associated " $t$ " statistics of about 2.3. This is due in part to the very large sample sizes in the CPS.

It appears that these differences are not large enough to materially affect the conclusions of the analysis. However, to further assess the issue of whether changes in the item parameters of the magnitudes observed in Table III. 1 should be judged large enough to be pertinent to the substantive analysis, we sought the advice of two experts who have worked with the Rasch model extensively and who are generally regarded as leading national authorities on its use: Dr. Benjamin Wright, professor of education at the University of Chicago, and Dr. Robert Mislevy, principal scientist at Educational Testing Service. Both experts view the food security results as remarkably stable, given their overall experience with similar models.

## IV. STANDARD BENCHMARK METHODOLOGY

As an aid to those doing research related to food security measurement, this chapter summarizes the steps involved in starting with raw survey data and producing household food security scores and category assignments. This material is consistent with, and provides a summary of, the material in the food security measurement guide revised by USDA in 2000 (Bickel et al. 2000). It should be noted that this material applies directly only to data collected using the 1995-1997 CPS data collection instrument. ${ }^{1}$ Certain changes were made in the 1998 and subsequent surveys, particularly in the screening questions. For a discussion of how these changes can be treated, see Bickel et al. 2000.

In Section A, we describe estimation procedures to be used if a researcher wishes to score raw survey data, using all the conventions and calibrations that are part of the official government method used in producing national food security statistics-as presented, for instance, in Andrews et al. (2000). In this situation, use of the Rasch model is not necessary; ${ }^{2}$ the required steps are described in Section A. In Section B, we describe the processing steps needed if a researcher wishes to reestimate the underlying model as part of the analysis. This might be done, for example, if a researcher wishes to check whether there have been any substantial changes in the underlying population parameters (that is, the "severity levels" of individual survey items) since the basic model was calibrated, or if the parameters for a sample group display distinctive characteristics different from the national patterns.

[^5]
## A. THE BASIC GOVERNMENT METHOD (NO RECALIBRATION OF THE MODEL)

The following steps are needed in order to apply the conventions used in this report to a rawdata set.

## 1. Check and Edit the Data

The first step is to examine the data to make sure that all relevant variables are within the ranges permitted by the survey and that the skip logic is consistent. Any problems should be resolved before proceeding to the next step.

## 2. Construct Binary Variables from the Raw Data for Each of the 18 Food Security Variables

All the variables used to calculate scores and assign households to categories have just two values-an affirmative and a negative (ignoring, for now, missing value codes). However, some of the survey questions have more than two possible responses. For instance, there are several questions in the food security module with response categories of "often true," "sometimes true," and "never true." These variables must be recoded to binary form. In the preceding example, for instance, both the "often true," and "sometimes true" categories are coded to "1" (that is, affirmatively), with the third category coded to " 0 ." The rules for doing this recoding are presented in Bickel et al. (2000, p. 28), which lists the wording of each question and its answer categories.

## 3. Screening $^{3}$

Apply the following criteria to each household:
${ }^{3}$ These screening procedures apply to data collected with the CPS instruments used in 19951997. They may have to be adapted for instruments with different screeners. See, for example, Bickel et al. 2000.

3a. Determine whether the household is "poor," as defined by Table IV.1.
3b. Now apply the screens in Exhibit IV.1, as follows:
If the household is "poor" and if the household passed Screen 2 from Exhibit IV.1, then retain the household in the sample for the detailed food security analysis.

If the household is non-poor, and if the household passed Screen 1 from the Exhibit IV.1, then retain the household in the sample for the detailed food security analysis.

3c. If neither of the conditions in Step 3b applies, then classify the household as food secure, and do not compute a score for the household on the continuous food security scale.

## 4. Impute Missing Data

If a case has at least one valid answer for the 18 food security variables-but if there are missing data for some of the relevant variables-the next step is to impute values for the missing values. This is done using the algorithm presented in Exhibit IV.2.

## 5. For Each Household, Count the Number of Affirmative Answers to the $\mathbf{1 8}$ Food Security Variables

Next, for each household, compute the total number of affirmative answers.

## 6. Determine the Continuous Scale Score and Food Security Classification

The scale score and food security classification for a household can then be determined from the number of affirmative answers, using Table IV.2. The scale score for a household with all negatives (18 for a household with children, 10 for a household without children) is technically undefined because such a household is outside the range that can be measured using the CPS food security supplement; however this score is set to 0 as an approximation. The score assigned in the table to households with all affirmatives is also not derived from the Rasch model but, rather, reflects the judgment of a reasonable score that reflects the situation of the small number of households who respond to the question in this way.

TABLE IV. 1
THRESHOLDS USED TO CLASSIFY HOUSEHOLDS AS "POOR," AND DHHS POVERTY GUIDELINES AT TIME OF SURVEY

|  |  | Threshold as a Percent of Poverty Guideline |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
|  | "Poor" if Family <br> Income Below <br> (in Dollars): | 1994 | 1995 | 1996 | 1997 |
| 1 | 15,000 | 204 | 201 | 194 | 190 |
| 2 | 20,000 | 203 | 199 | 193 | 189 |
| 3 | 25,000 | 203 | 199 | 193 | 188 |
| 4 | 30,000 | 203 | 198 | 192 | 187 |
| 5 | 35,000 | 203 | 198 | 192 | 186 |
| 6 | 40,000 | 202 | 197 | 192 | 186 |
| 7 | 50,000 | 225 | 219 | 213 | 207 |
| 8 | 50,000 | 202 | 197 | 192 | 186 |
| 9 | 60,000 | 221 | 215 | 209 | 202 |
| $10+$ | 75,000 | 253 | 246 | 239 | 232 |
| Average |  | 212 | 207 | 201 | 195 |
| Average for Households |  | 203 | 199 | 193 | 188 |
| With 6 or Fewer Persons |  |  |  |  |  |

## EXHIBIT IV. 1

## SCREENS USED IN THE CPS SURVEYS

## Screen 1

Households reporting any one of the following passed the screen:

- Sometimes or often not enough to eat; or
- Ran out of food in last 12 months; or
- Not always the right kind of food and ran short of money.


## Screen 2

Households reporting any one of the following passed the screen:

- Sometimes or often not enough to eat; or
- Ran out of food in last 12 months; or
- Not always the right kind of food; or
- Ran short of money.


## EXHIBIT IV. 2

## ALGORITHM FOR IMPUTING MISSING FOOD SECURITY DATA

1. Items are ranked by increasing severity (see below) ${ }^{\text {a }}$
2. Only cases with $1+$ nonmissing responses are processed. For each of those cases, do the following:
3. Find the most severe "yes" response.
4. Find the least severe "no" response.
5. Identify the less severe of steps \#3 and \#4, above.
6. For all missing items less severe than step \#5, impute a "yes."
7. For all missing items more severe than step \#5, impute a "no."
8. Do not change any nonmissing items.

The item ranking used in implementing the above is:
Q53
Q54
Q55
Q58
Q24
Q56
Q32
Q25
Q57
Q35
Q38
Q40
Q28
Q47
Q29
Q43
Q44
Q50
${ }^{a}$ The items that refer to children are only used for households with children.

TABLE IV. 2

## CONTINUOUS SCALE SCORES BASED ON THE NUMBER OF AFFIRMATIVE ANSWERS

$\begin{array}{ccc}\hline \text { Number of Yes Responses } & & \\$\cline { 1 - 2 } \(\left.$$
\begin{array}{c}\text { Household with } \\
\text { Child }\end{array}
$$ \& $$
\begin{array}{c}\text { Household with } \\
\text { No Child }\end{array}
$$ \& $$
\begin{array}{c}\text { Scale Value for Standard } \\
\text { Computational Metric }\end{array}
$$ <br>
\hline 0 \& 0 \& 0.0 <br>
1 \& \& 1.4 <br>
Food Security Status <br>

Category\end{array}\right]\)| Food secure |
| :---: |
| 2 |

[^6]
## B. CHECKING THE CALIBRATION OF THE MODEL ON THE DATA SETS

As noted earlier, reestimation of the item parameters of the model is not necessary for scoring a new data set using the standard government method. It is sometimes useful, however, to run the model on new data sets as they become available, in order to assess whether any substantial changes have occurred in item calibration that might signal changes in the underlying phenomena of food insecurity and hunger being examined. Also, some researchers may wish to reestimate the model as part of a strategy to use different imputation methods than those used in this report. Below, we describe the steps involved in computing food insecurity scores when reestimation of the model is desired.

## 1. Edit the Data

Same as under Section A above.

## 2. Construct Binary Variables from the Raw Data for Each of the $\mathbf{1 8}$ Food Security Variables

Same as under Section A above.
3. Screening

Same as under Section A above.

## 4. Estimate the Rasch Model Parameters

The next step in the process is to estimate both household parameters and item parameters for the Rasch model. This can be accomplished using one of several software programs currently
available. ${ }^{4}$ Following are some guidelines for doing this:
C Some software packages allow estimation of IRT models involving more than one parameter. However, if a Rasch model comparable to that which has been used in this report is desired, a one-parameter logistic IRT model should be specified.

C As noted earlier, a Rasch scale is unique only up to a linear transformation. Each software package has its own default normalization for fixing this scale. The conventions used for this purpose in this report are: (a) the mean of the item parameters is set equal to 7 ; (b) the slope of the item response functions at their inflection points are set equal to 1 . This normalization facilitates the assessment of parameter values in comparison to the national benchmark values by utilizing the same metric (unit of measure).

C The software produces both item parameters and household parameters. The item parameters are the question-severity levels; the household parameters are the household scores on the continuous household scale.
${ }^{4}$ For instance: BIGSTEPS Rasch-Model Computer Program, MESA Press, 5835 Kimbark Ave., Chicago IL. 60637-1609. BILOG Scientific Software International, 7833 N. Lincoln Ave., Suite 100, Lincolnwood, IL 60712-1704.

## V. DEFINING "CUTPOINTS" OVER TIME

The Rasch model places each household in a continuous numeric food security scale. For purposes of policy analysis, it is also useful to assign households to a small number of categories corresponding to designated ranges of severity on the underlying continuous scale, that can summarize their food security status. To create this categorical measure, Hamilton et al. (1997a) specified four categories: food secure, food insecure without hunger, food insecure with moderate hunger, and food insecure with severe hunger. More recently, the latter two categories have usually been collapsed to form a single category.

In this chapter, we first summarize the work done by Hamilton et al. in categorizing households based on the output of the Rasch model. An essential step in that work was the specification of "cutpoints," which partition the continuous food security scale into the designated categories. After describing how Hamilton et al. established cutpoints, we highlight several key issues that arise in defining cutpoints, and then present alternative approaches to dealing with these issues.

## 1. Earlier Work

Hamilton et al. (1997a) established numerical "cutpoints" on the continuous food security scale and placed households in categories based on their numerical scores in relation to these cutpoints. For instance, on a 0 to 10 scaling of the continuous variable as used in that work, households with scores less than approximately 2.1 were ranked as food secure, those with scores between approximately 2.1 and 4.5 were ranked as food insecure without hunger, and so on.

These cutpoints were set by a group of experts that included members of the Abt Associates research team, academic consultants, and staff from the USDA Food and Nutrition Service (FNS), the Economic Research Service (ERS), and other Federal agencies. The process they went through
was based on a detailed analysis of the content of the food security module questions, as that content related to the conceptual basis of the food security severity-range categories being developed.

Members of the team focused largely on "modal" households-namely, households that answered negatively to all the more severe questions, once they had answered any single question negatively. (Such a response pattern may not be modal in the statistical sense of "most frequent.") Through careful analysis of the content of the question sequence, the working group assessed how modal households answering affirmatively to varying numbers of questions in the sequence should be classified.

In locating these cutpoints, the team was concerned to identify the threshold level of credible evidence needed to identify the particular conditions of food insecurity described by each designated category, or range of severity, on the underlying empirically determined scale. For example, in locating the threshold degree of severity for classifying a household as "food-insecure," it was judged that affirmative responses to at least three independent indicators of food insecurity was sufficient to identify the condition of food insecurity. (Most food-insecure households show more than three of the scaled indicators of the condition, but all households classified as food insecure must show at least three.) ${ }^{1}$

Following the same logic, the technical team next identified the "hunger threshold" for households within the sequence of these scale items, ordered by measured severity level, that capture reported reductions in quantity of food intakes for adults in the household. The same procedural rule

[^7]was adopted as for the food-insecurity threshold—that is, multiple separate indications of reduced food intake for adults, including a repeated pattern of cutting or skipping meals in three or more months, were judged sufficient for defining the measurement category and for classifying households to the category, "food insecure with hunger evident."

In the context of Rasch measurement of the severity of a phenomenon, these operational rules become extremely simple for households that answer completely all questions in the scale sequence (that is, item nonresponse of zero).

The substantive analysis described above led to decisions on approximate values as to where the numerical scale cutpoints should fall. Final exact values of the cutpoints were then determined by Abt technical staff. ${ }^{2}$

## 2. Generalization of Cutpoints to Other Years

During the current contract, as the research was extended to the 1996 and 1997 data, the issue arose as to whether it was appropriate to use the same 1995 numerical cutpoint values for the later years. The conclusion reached was that, if the Rasch model were recalibrated on the new data sets, use of the same numerical cutpoints might not necessarily be feasible. The reasons for this, and the alternative procedures considered, are described below.

Background. A key basis for the following discussion is that, for any given Rasch model calibration, households with complete data that give the same number of affirmative responses are assigned the same score, independent of which questions are answered affirmatively. This characteristic means that, ignoring cases with missing data for now, all households with children

[^8]were assigned to 1 of only 17 points on the continuous scale, corresponding to the possible range of affirmative answers from 1 to $17 .{ }^{3}$ Similarly, households without children were assigned to one of nine locations on the scale. Thus, the complete-data households, which are nearly all the households, are not distributed smoothly across the continuous scale; rather, the distribution is "lumpy," with 26 clusters of households.

The only households not assigned to 1 of these 26 clusters are the relatively few with missing items on the food security data items. Because scores assigned to cases with missing data depend both on how many items they are missing and which items they are missing, their scores tend to be distributed across the scale, without significant clustering.

Why Constant Numerical Cutpoints May Not Be Feasible. In principle, if a Rasch model is correctly specified, the estimated item-severity levels should not change when the model is estimated on different years' data. In practice, however, some changes between years in these parameters may be inevitable, due to statistical sampling error and, probably, to at least some modest changes in the phenomenon being studied. This, in turn, implies that there will be at least slight changes between years in the points on the continuous scale to which households with any given numbers of affirmative answers are assigned.

This leads to the following analytic possibility: if a cutpoint happened to be located on the food security scale very near the score for a cluster of households, then it is possible that a slight movement in the cluster's score from one year to the next, perhaps caused entirely by sampling error, could lead to all the households in the cluster being placed in a different category. The lumpiness
${ }^{3}$ Households at the extremes, with 0 or 18 affirmative answers, cannot be assigned a specific place on the scale. This does not affect the argument in the text; thus, for ease of exposition, we ignore it.
in the distribution of scores could cause food security prevalence estimates to be more unstable than the real underlying characteristics being measured.

A generalization of the concern is as follows:

If the model is reestimated on a different data set, but the cutpoints used in assigning households to categories are not changed, there is a risk that households with the same answer patterns could be assigned to different categories in the two versions of the model, because of changes in the model's estimated item parameters.

In practice, it is not clear exactly how serious the risk is in the current application. Because of the high degree of model stability observed from one year to the next (see Chapter III), the numeric scale values associated with any given response pattern change relatively little. Indeed, experiments conducted with the 1995-through-1997 data suggest that it would have been possible to use the 1995 cutpoints developed in Hamilton et al. (1997) with models based on each of the three years of data, without encountering the problem of large clusters of households with the same answer pattern being assigned to different categories in different years. However, there is nothing in the Rasch modeling process to ensure that this would not happen in the future, if cutpoints values are held constant.

## 3. Alternative Approaches to Setting Cutpoints

As noted above, the algorithm for translating from the Food Security Module question responses to the food security categories was originally developed in terms of the content of the question sequence. In light of this, it would appear that a highly desirable objective of the food security measurement process should be to develop a method which ensures that, at least for cases with complete data, the same pattern of responses is consistently translated into the same food security category classifications. This feature is particularly desirable for households with complete data, since, as described above, these cases are clustered at only 26 points on the numerical scale, and,
therefore, a change in food security classification for one of these clusters of households could have substantial implications for estimates of food security prevalence levels.

As noted above, reestimating the model each year, but keeping the same cutpoints, does not guarantee that the objective of a constant relationship of question response patterns to assigned food security levels is met. How can this objective be achieved? Four possible approaches for doing this are described here:

1. Hold the model calibration and cutpoints constant. As long as the model parameters used to assign numerical food security scores to households are kept constant over time, this will guarantee that the same response patterns are equated to the same numerical scores. If the cutpoints are held constant, this, in turn, ensures that the same answer patterns are equated to the same household classifications. In the current application, for instance, if the 1995 parameters are consistently used, the cutpoints established with the 1995 data will always achieve the desired objectives. Under this approach, a simple algorithm can be specified for complete-data households that maps the number of affirmative responses into a category assignment. For households with missing data, however, Rasch model software can be used to obtain numerical scores for comparison with specified cutpoints.
2. Define the classifications of household response patterns into food security levels directly, without going through the intermediate step of a numerical score. This essentially defines away the cutpoint problem. However, by itself, it does not provide a method for dealing with cases with missing data. In the Rasch model, the score assigned to a household with missing data depends on the exact number of missing items and their severity levels.
3. Impute responses for the missing items, then define the classification of households as in the previous alternative. This method, which is part of the government's benchmark methodology described in Chapter IV, involves developing a relatively simple algorithm for imputing missing data and then translating from the number of affirmative answers (including any imputed affirmatives) to a food security classification. It is similar to Alternative 2, but includes a solution to the problem of dealing with missing data cases.
4. Recalibrate the model and the cutpoints with each new application. The fourth alternative is to set new cutpoints with each new application of the model. One way of doing this would be to place the cutpoints between the scale scores assigned to households that "bound" the edges of the food security categories. For instance, under the procedures used in the Hamilton et al. analysis and in this report, a complete-data household with seven affirmatives is categorized as food insecure without hunger, and
one with eight affirmatives is categorized as food insecure with hunger. Therefore, for any given Rasch model calibration, the numerical cutpoint that will ensure these households are categorized as desired, must be between the numerical score assigned in that new model calibration to households with seven affirmatives and the score assigned to households and with eight affirmatives. Where within this range to assign the cutpoint is arbitrary; however, the midpoint has some attraction as a practical solution, because it is likely to reduce the sensitivity of classifications of missing data households.

Summary. All but the second of the four methods outlined above represent reasonable approaches to the issue of how to ensure that households with the same response patterns are grouped into the same food security categories from one year to another. The first and third are the same, except in the treatment of missing data. From among these two, the third is recommended, because it is simpler and more intuitive and because it makes the method more accessible to other researchers by not requiring use of Rasch model software.

The disadvantage of the second approach is that, by itself, it does not provide a solution for dealing with missing data cases. Finally, the fourth alternative, like the first, has the disadvantage of requiring Rasch model software for implementation. This alternative also involves the cutpoints changing somewhat over time, which could potentially be confusing to some users of the analysis. An advantage of the last approach, however, is that it captures the effect of changes over time in item parameters.

## VI. CONCLUSION: REFLECTING ON THE STRENGTHS AND LIMITATION OF THE FOOD SECURITY METHODOLOGY

In this concluding chapter, we draw upon the results of our research, together with related work, to discuss the strengths and limitations of the use of the Rasch model as a basis for food security measurement. Possible directions for future research are also noted.

## A. STRENGTHS AND LIMITATIONS OF THE METHODOLOGY

As indicated earlier, the food security scale reflects more than 10 years of methodological development by both government and private groups. Use of the IRT methodology has made it possible to guide the development of the food security estimates with a thoroughly studied model that has well-understood statistical properties. The approach has undergone extensive review by experts in both the public and the private sector. In general, these experts consider the model an appropriate application of the IRT methodology, and they have viewed the analysis results as reasonable.

Another important strength, which has been revealed by the current project, is that the estimated item parameters of the IRT model are robust across time and population subgroups. The values obtained from the 1996 and 1997 data are essentially the same as the original 1995 values. ${ }^{1}$ In addition to stability over time, there is stability across subgroups defined by such characteristics as race/ethnicity, household composition, and region of the country. (See Appendix E.)

Tempering these strengths are several limitations that should be recognized (as noted earlier on pp. xx-xxi). Perhaps the most important of these limitations occurs as a result of the process used to translate the continuous food security scale score into food security categories (that is, food secure,

[^9]food insecure without hunger, and food insecure with hunger). The process of developing rules for translating the values on the continuous scale into these categories was inherently much more judgmental than the process of deriving continuous scale scores. To be sure, the location of each boundary was carefully set by a broad-based group of experts to reflect the consensus of their judgment about how the words in the conceptual basis of the designated severity range categories should be interpreted, defined operationally, based on the observed sequence of indicator items available in the CPS data. However, room for disagreement remains as to what patterns of response to the questions should be construed as reflecting appropriately the naming conventions used in designating the severity-range categories.

This problem is exacerbated by the fact that many households do not exhibit response patterns that are exactly consistent with the item hierarchy. While such "nonmodal" answers are to be expected, given the underlying probabilistic model (in the same way that a student has a chance of getting a hard math problem right after getting an easier problem wrong), nonmodal answers sometimes create patterns that defy easy classification into the prevalence groups. An example of this type of difficulty is that it is possible for a household to be classified as "food insecure with hunger," even though the respondent may have answered negatively the question, "In the past 12 months, .....were you ever hungry but didn't eat because you couldn't afford enough food?" For this to happen, the respondent either must have replied "yes" to a series of increasingly severe indicators of food insufficiency, including at least three items indicating reduced food intake for themselves and/or other adult members of the household, one of which establishes a repeated pattern of such reduced intakes over the year; or they must have replied "yes" to most of the foregoing, plus one or more of the items that are more severe than the explicit hunger question. The categorical measure (and its naming convention) reflects the judgment that, on the balance of this evidence, one or more
adult members of the household has, with high probability, experienced resource-constrained hunger sometime during the year. Conversely, the opposite case also can occur: the household can be classier "food insecure without hunger," based on its overall pattern of response and the resulting scale score, even though the respondent has answered "yes" to the explicit hunger question as such. They have argued that probabilistic labels for the categories would be more accurate, and certainly more consistent with the underlying probabilistic model.

An additional limitation of the food security data available for the years 1995-1997 stems from the use in the CPS data collection of a preliminary screen for high-income households which was changed slightly in each of the first three years of data collection, as well as a (somewhat looser) screen applied to low-income households in one year (1996). The need to use a common screen for the 1995-1997 period to compensate for differences in the survey screens tested during these years requires imposing the assumption that certain households are food secure, even though they gave survey responses indicating the contrary. While this convention is necessary to ensure strict comparability in analysis findings across this time, the use of the screens involves discarding information, and the method used of retroactively applying a common screen, results in estimated prevalence rates that are lower than they would otherwise be. Tests of the sensitivity of the findings to differences in various assumptions and procedures found that the screening procedure used, especially in 1996, tend to have a relatively large impact. This problem is exacerbated by weaknesses in the income data present on the relevant CPS files, since some households probably are being misclassified as to income level, on which the application of the preliminary screen is based.

## B. DIRECTIONS FOR FURTHER RESEARCH

Several potential directions for future research are suggested by these analyses. One important activity will be to continue to estimate food security prevalence rates for future years, in order to monitor trends in U.S. and state-level food security. Associated with this, we recommend continuing to estimate the IRT model item parameters each year, as data become available. In contrast to the approach adopted in this report, in which the prevalence rates are estimated without reestimating the model, the reestimation process is, in our judgment, important for monitoring whether underlying structural changes are occurring that would warrant reassessing the model itself.

An additional line of research that is of considerable interest, and that has already begun (Nord and Bickel 1999), is to develop a measure for estimating the prevalence of hunger among children. It is possible that adults respond differently to food insecurity, depending on whether there are children (especially young children) in the household, and separate scales based on adults and childspecific indicators could reflect these differences.

A more substantial extension of the current work, which would be valuable, is to develop food security models for individuals as well as households as the units of observation. There is considerable interest in the policy community in obtaining estimates of particular categories of adults and (especially) of children who are food insecure. The work done so far-based on survey questions asked mostly about all household members collectively—provides only a likely upper bound on such estimates. For instance, while we can estimate the number of children living in households that are food insecure with hunger, we cannot yet directly estimate the number of children who themselves have been hungry due to lack of household resources. The latter number may be significantly smaller than the former, for instance, if parents to shield some of their children
from hunger more than others, or if older children shield younger children, even at the cost of experiencing greater hunger themselves.

Making substantial headway on the individual scales will require a different data collection instrument than the current CPS Food Security Supplement, since most questions on that supplement are asked at the household level. Experimental work in this direction was undertaken with one rotation group in that survey. More extensive work developing individual scales can be carried out with data from the Fourth National Health and Nutrition Examination Survey (NHANES4) and the most recent Continuing Survey of Food Intakes by Individuals. Both of these data sources are collecting key food security items at the personal as well as household level.

## REFERENCES

Anderson, J.J., C.A. Wehler, and R.I. Scott. "Scaling and Indexing to Measure the Severity of Food Insecurity and Hunger." In conference on Food Security Measurement Research: Papers and Proceedings, January 21-22, 1994. Alexandria, VA: U.S. Department of Agriculture, Food and Consumer Service, 1995.

Andrews, Margaret, Mark Nord, Gary Bickel, and Steven Carlson. Household Food Security in the United States, 1999. Washington, DC: U.S. Government Printing Office, 2000.

Bickel, Gary, Mark Nord, Cristofer Price, William Hamilton, and John Cook. Guide to Measuring Household Food Security. Washington, DC: U.S. Department of Agriculture, Food and Nutrition Service, March 2000.

Hambleton, Ronald H. "Principles and Selected Applications of Item Response Theory." In Educational Measurement, Third Edition, edited by Robert Linn. New York: Collier MacMillan, 1993.

Hamilton, William L., John Cook, William L. Thompson, Lawrence F. Buron, Edward A. Frongillo, Jr., Christine M. Olson, and Cheryl A. Wehler. Household Food Security in the United States in 1995: Summary Report of the Food Security Measurement Project. Cambridge, MA: Abt Associates, 1997a.

Hamilton, William L., John Cook, William L. Thompson, Lawrence F. Buron, Edward A. Frongillo, Jr., Christine M. Olson, and Cheryl A. Wehler. Household Food Security in the United States in 1995: Technical Report of the Food Security Measurement Project. Cambridge, MA: Abt Associates, 1997b.

Nord, Mark, and Gary Bickel. "The Prevalence of Children’s Hunger in U.S. Households, 19951999." Washington, DC: U.S. Department of Agriculture, Economic Research Service, August 9, 2000.

Ohls, Jim, Allen Schirm, Abhijay Prakash, and Larry Radbill. "Methodogical Findings and Early Conclusions Based on the 1995, 1996, and 1997 Food Security Data." Paper presented at the Second Food Security Measurement and Research Conference sponsored by the Economic Research Service and Food and Nutrition Service, United States Department of Agriculture, Alexandria, VA, February 1999.

Price, Cristofer, W. Hamilton, and J. Cook. Guide to Implementing the Core Food Security Model. Cambridge, MA: Abt Associates, 1997.

Radimer, K., C. Olson, J. Greene, C. Campbell, and J. Habicht. "Understanding Hunger and Developing Indicators to Assess It in Women and Children." Journal of Nutrition Education, vol. 24, supplement, January/February 1992, pp. 35S-45S.
U.S. Department of Agriculture, Food and Nutrition Service. "Measuring Food Security in the United States. Household Food Security in the United States 1995-1998 (Advance Report)." Washington, DC: USDA, July 1999.

Wehler, C., R. Scott, J. Anderson, and L. Parker. Community Childhood Hunger Identification Project: A Survey of Childhood Hunger in the United States. Washington, DC: Food Research and Action Center, 1991.

Wright, Benjamin D., and Geofferey N. Masters. Rating Scale Analysis. Rasch Measurement. Chicago: Mesa Press, 1982.

## APPENDIX A

SUMMARY OF SALIENT PARTS OF ITEM RESPONSE THEORY AND RELATED MEASUREMENT ISSUES

## SUMMARY OF SALIENT PARTS OF ITEM RESPONSE THEORY AND RELATED MEASUREMENT ISSUES

Item response theory (IRT) has played an important role in the development of the food insecurity scale and forms the basis of the continuous form of the scale. In the original analysis of the 1995 CPS data, Hamilton et al. (1997) used IRT modeling, in conjunction with other analytic techniques, to identify a core set of items that could be used to construct a scale to measure the relative degrees of food insecurity of different households. IRT models were also used in that work to "calibrate" the model and perform initial analyses of the stability of the index across several key demographic groups.

In the analysis of the 1996 and 1997 data, we have relied on IRT methods to assess the temporal stability of the models and to assess the sensitivity of the index to different screening procedures used in the April 1995, September 1996, and April 1997 CPS food security supplements. We have also used IRT to analyze the stability of the index across demographic groups.

This appendix presents an introduction to IRT, since it may not be familiar to some readers. Our presentation emphasizes those aspects of the theory and models that are most pertinent to the use of IRT models in measuring food insecurity. ${ }^{1}$

The balance of this appendix is organized as follows. Section 1 provides an overview of IRT methods. Section 2 describes in greater detail the variant of IRT methods that has been used in the food security research. Section 3 discusses methods for estimating parameters.

[^10]
## 1. Overview of IRT

IRT models are a subset of a broader group of latent variable models that includes LISREL models and other factor analytic models used in the social sciences. IRT begins by assuming the existence of some characteristic that cannot be measured directly but can be assessed indirectly. The objective is to measure this latent trait. For example, IRT methods have been used most extensively in educational testing to measure student ability levels. They are used, for instance, by Educational Testing Service (ETS) for scoring and test development work with the Scholastic Aptitude Tests (SAT). IRT models are also used to measure trends in cognitive achievement among U.S. children in the National Assessment of Educational Progress (NAEP). We will be focusing on a specific IRT model: the one-parameter logistic model, also known as the Rasch model (named for Georg Rasch, who introduced the model in the 1950s) (Wright and Masters 1982).

The IRT model is based on the assumption that the latent trait is a unidimensional continuum. ${ }^{2}$ In traditional educational testing applications, the characteristic is some type of ability or aptitude (such as the quantitative skills tested in parts of the SAT exams). In our application, the characteristic is a household's relative level of "food insecurity."

It is also assumed that the characteristic being measured is associated with respondents—students in the traditional testing applications, but households in our application-and that, in principle, it is possible to determine where every respondent falls on the underlying continuum relative to every other respondent. The continuum is measured as an interval level scale, which means that we cannot determine the absolute quantity of the characteristic that each respondent has, but we can determine the relative differences between respondents. In other words,

[^11]we can compare the positions of respondents $\mathbf{A}, \mathbf{B}$, and $\mathbf{C}$ and then determine how much closer (and in which direction) $\mathbf{A}$ is to $\mathbf{C}$ than $\mathbf{B}$ is to $\mathbf{C}$. For example, $\mathbf{A}$ might be twice as far below $\mathbf{C}$ as $\mathbf{B}$ is above $\mathbf{C}$ (as would be the case, for instance, if $\mathbf{A}=-1, \mathbf{B}=2$, and $\mathbf{C}=1$ ).

While the characteristic itself cannot be measured directly, there are indirect indicators that can be observed and that can provide information about the characteristic. In traditional applications, these indicators are responses to test questions; in the current application, they are responses to questions on a household survey. Just as respondents have locations on the underlying scale, these indicators also have locations on the same underlying scale as the respondents. In traditional testing applications, we think of this in terms of the relative difficulty of different test items. In our application, some survey questions ask about behaviors associated with more severe levels of food insecurity than are asked about by other questions.

In the current application, the Rasch model has two important advantages over other latent variable approaches, such as LISREL models. First, because the "ability" of students and the "difficulty" of test items can be estimated independently of each other, this IRT model provides a framework for assessing how test items function, across groups, over time, and relative to each other. $^{3}$ For example, with IRT models, it is possible to address each of the following questions:

- Do all respondents of the same ability find a given item equally difficult?
- Do this year's respondents find an item more or less difficult than respondents of the same ability found it last year?

[^12]- If item $\mathbf{A}$ was more difficult than item $\mathbf{B}$ last year, is it still more difficult this year? Has its difficulty relative to Item B changed?

The key in all three examples is the assumption that the evaluation of the item is independent of the ability of the students taking the test. If one group of students have much higher ability than a second group, the first group will be more likely than the second group to answer each of the test questions correctly. The Rasch model provides a way to determine whether the test items are more (or less) difficult for one group, relative to another group, after accounting for their different abilities. In the current application, the model provides the means by which survey questions can be evaluated across groups with different degrees of food insecurity, and over time, even when the prevalence of food insecurity changes. In similar fashion, the evaluation of respondent ability is independent of the difficulty of the items used to construct the test. ${ }^{4}$

A second advantage of the IRT models is that they allow for "test equating." In traditional applications, students may be given different sets of test questions to discourage cheating. Computerized "adaptive" tests, under which each student receives a test that is tailored to his or her own ability level, are increasingly being used to achieve greater testing efficiency. In these applications, there is a need to score all students on the same scale, despite their having been given different tests. In the current application, childless households were not asked the eight food security questions that refer specifically to children. Despite this, there is a need to assess the level of food insecurity of all households, regardless of whether children are present. The IRT model used here

[^13]provides a way to equate the 10 -item test administered to childless households with the 18 -item test administered to households with children. ${ }^{5}$

## 2. Formal Properties of the One-Parameter Logistic IRT Model

## Formal Statement of Model

The IRT model used here is built from the following assumptions:

- There is a unidimensional variable that is not observed directly (for example, quantitative aptitude or food insecurity).
- Each respondent (student or household) has a position on that latent variable, denoted by $b_{j}$, where the subscript $j$ identifies each respondent. The $b$ are referred to as the respondent parameters in the model, and a value for $b$ is estimated for each respondent. In educational testing, $b$ is the respondent's ability.
- Each test item has a position on the latent variable, denoted by $d_{i}$, where the subscript $i$ identifies each item. The $d_{i}$ are referred to as the item parameters in the model, and a value for $d$ is estimated for each item. In educational testing, $d$ is the item's difficulty.
- The probability that a given respondent "correctly" answers a specific test item is only a function of the difference between that respondent's ability $(b)$ and that test item's difficulty $(d) .{ }^{6}$ Specifically,

$$
\begin{equation*}
\text { Probability }\left(\text { Correct Response } \mid b_{j}\right)=\frac{e^{\left(b_{j}-d_{i}\right)}}{1+e^{\left(b_{j}-d_{i}\right)}} \text {, } \tag{1}
\end{equation*}
$$

or, equivalently,
${ }^{5}$ In earlier work on food security, Hamilton et al. (1997) also used IRT models, in conjunction with other models, to identify the specific questionnaire items to be used for constructing the food insecurity index.
${ }^{6}$ There are many different values for $b_{j}$ and $d_{i}$ that yield the same probability. All such pairs of values are linear functions of each other. In order to estimate unique values for $b_{j}$ and $d_{i}$, it is necessary to impose two constraints. One commonly used constraint is to reposition all the $d_{i}$ so that they have a mean of 0 . A second constraint that is commonly used fixes the slope of the response function shown in Equation 1, evaluated at its inflection point (where $b_{j}$ equals $d_{i}$ ), to be equal to 1 . In the next section, we discuss in greater detail how $b_{j}$ and $d_{i}$ are estimated.
(2) $P\left(\Theta_{j}\right)=\frac{e^{\left(\Theta_{j}-d_{i}\right)}}{1+e^{\left(\Theta_{j}-d_{i}\right)}}$

Note that the probability of a correct response to any single item is only a function of the respondent's ability $(b)$ and the item's difficulty $(d)$. The probability of a correct response does not depend on any of the other test items. This property-known as conditional local independence-is an important assumption of this IRT model. ${ }^{7}$ Also, the probability of a correct response does not depend on any characteristics of the respondent other than ability $(b) .{ }^{8}$

## Item and Test Characteristic Curves

Equation 1 is the formula used to compute the probability that a specific respondent (respondent $j$ ) will correctly answer a specific test item (item $i$ ). Figure A. 1 shows a plot of Equation 1 for five equally spaced hypothetical test items. These plots are called "item characteristic curves." The horizontal axis is the latent variable, and the vertical axis is the probability of a correct answer to a test item. The difficulty of an item is the value on the horizontal axis where this probability is .5 . The first item has a difficulty $(d)$ of -2 , because when the value on the latent variable is -2 (read on the horizontal axis), the probability of a correct answer is 0.5 (read on the vertical axis). The other four items have difficulties of $-1,0,1$, and 2. Figure A. 2 shows item characteristic curves for a

[^14]
Note: As explained in the text, the exact scaling of the horizontal axis is arbitrary. This figure is centered at 0 for ease of exposition.

Note: As explained in the text, the exact scaling of the horizontal axis is arbitrary. This figure is centered at 0 for ease of exposition.
different set of five hypothetical items. In this example, the items are not equally spaced: while the first and last items still have difficulties of -2 and 2, the middle three items are clustered close to 0 : they have difficulties of $-0.1,0.0$, and 0.1 .

Just as multiple items can be combined to form a test, multiple-item characteristic curves can be combined to produce a test characteristic curve. Raw test scores are computed by summing the number of correct answers, and the test characteristic curve is created by summing the characteristic curves of the items included in the test. Figure A. 3 shows the test characteristic curve that results when the items shown in Figure A. 1 constitute the test. Figure A. 4 shows a similar curve for the test that results when the five items from Figure A. 2 are combined. The interpretation of the vertical axis of the test characteristic curve is slightly different from its interpretation in the context of item characteristic curves. Rather than represent the probability of a correct response to a single item, the vertical axis represents the expected number of correct responses to all items included in the testthe expected value of the respondent's raw score on the test. ${ }^{9}$

Just as an item characteristic curve relates the probability of a correct (or "positive" in the food security context) answer to a respondent's estimated location on the latent variable, a test characteristic curve relates the number of correct answers on a test to a respondent's location on the latent variable. For example, in Figure A.3, respondents who answer one of the five items correctly (read on the vertical axis) are assigned a scale score of -1.9 (on the horizontal axis). Respondents with raw scores of 2,3 , and 4 are assigned scale scores of $-0.5,0.6$, and 2.0 , respectively. (While

[^15]
Note: As explained in the text, the exact scaling of the horizontal axis is arbitrary. This figure is centered at 0 for ease of exposition.

Note: As explained in the text, the exact scaling of the horizontal axis is arbitrary. This figure is centered at 0 for ease of exposition.
it may not be intuitively obvious, it can be shown mathematically that these are maximum likelihood estimates of the respondents' latent trait values.)

## Impossibility of Scoring "Extreme" Response Patterns

Several important points should be noted. First, using this test and this IRT model, respondents with raw scores of 0 or 5 cannot be assigned scale scores. In the case of no correct responses, all we know is that the respondent's scale score ability is below the range that can be measured by this test. The problem is, the model is completely uninformative about how far below the range of this test the respondent's score lies. In the case of a perfect score (a correct response for each of the five items), all we know is that the respondent's scale score is above the range that can be measured by this test. Again, however, the model provides no information about how far above.

In the current food security application, it is clear that the IRT model cannot provide scale scores to households that answered "Yes" to every item they were asked. Because there are very few such households, this presents no real problem. However, the model also cannot provide scale scores for households that answer "No" to all items (households with raw scores of zero). This is a limitation, because nearly 90 percent of the sample falls in this category. It is clear that such households should be classified as "food secure," and they are so identified in the ordinal variable that classifies households into one of the three food-insecurity categories. However, the IRT results are uninformative about how much more secure they are than the households that answered "Yes" to only one item. The model is also uninformative about whether all such households are equally secure. Although none of the households reported any behaviors that are symptomatic of food insecurity, some are undoubtedly at greater risk than others.

## Lack of a Direct Relationship Between Item Scores and Respondent Scores

Another important feature to note is that the scale scores assigned to respondents do not have any simple relationship to the difficulties of the items. In Figure A.1, the five items had difficulties of $-2,-1,0,1$, and 2 ; but the scores assigned to respondents in Figure A. 3 are $-1.9,-0.5,0.6$, and 2.0 (and notice that the score of 2.0 is assigned to students with 4 positive answers, not 5). The lack of an obvious relationship between the item parameters and the respondent parameters can be seen even more clearly by comparing Figures A. 2 and A.4. While the items in Figure A. 2 had difficulties of $-2,-0.1,0,0.1$, and 2 , the scale scores assigned to respondents are $-1.8,-0.5,0.5$, and 1.8 . Respondents with two positive answers are assigned a scale score ( -0.5 ) that is actually below the difficulty level of the second item in the test (-0.1). Partly because of the lack of a simple relationship between item difficulties and respondent scale scores, it is not useful to think of test items as directly measuring a specific level of ability. ${ }^{10}$

## 5. The Scaling of the Latent Variable Is Arbitrary

Scaling of the IRT parameters is arbitrary. We noted above that the latent variable is measured as an interval-level scale. We also noted that many pairs of values for $b$ and $d$ produce the same probability, and those pairs are each linear functions of each other. The three hypothetical IRT scales shown in Figure A. 5 are all equivalent: in every case, the difference between the first and fourth numbers is twice the difference between the first and third numbers.

[^16]FIGURE A. 5

| THREE EQUIVALENT IRT SCALES |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
| - -1 | 5 | 8 | 17 |  |
| - 11 | 17 | 20 | 29 | (added 12 to line 1) |
| - 10 | 70 | 100 | 190 | (added 2 to line 1, then multiplied by 10 ) |

In most of this report, we use a different scaling convention than Hamilton et al. (1997) used in their work. Therefore, some of the IRT parameters presented in this report appear to be substantially different from those derived in the earlier work, when, in fact, our estimates are close to those reported by Hamilton et al.

## Maximizing Information Content of Items

Although each test item provides information about respondents with a range of abilities, the amount of information depends on how close a respondent's ability is to a test item's difficulty. The amount of information provided by an item reaches its maximum for respondents with ability levels equal to the item's difficulty. This relationship can be formalized, and is shown in Equation 3:
(3) $I_{i}(\Theta)=\frac{\left[P_{i}^{\prime}(\Theta)\right]^{2}}{P_{i}(\Theta)\left[1-P_{i}(\Theta)\right]}$,
where $I_{i}(\Theta)$ is the information provided by item $i$ at the point $\Theta$ on the latent variable,
$P_{i}^{\prime}(\Theta)$ is the slope of the item response function (Equation 1) evaluated at $\Theta$,
and $P_{i}(\Theta)$ is the probability of a correct response to item $i$ by a respondent with ability $\Theta .{ }^{11}$

In the Rasch model, Equation 3 simplifies to Equation 4:

[^17](4) $I_{i}(\Theta)=P_{i}(\Theta)\left[1-P_{i}(\Theta)\right]$,
which reaches a maximum at $P_{i}(\Theta)=.5$, the point on the graph where the respondents' ability equals the items' difficulty.

Figure A. 6 illustrates the form of the information function for the five hypothetical items originally presented in Figure A.1, and Figure A. 7 shows the information curves for the five hypothetical items from Figure A.2. In each case, note that an item's information reaches a maximum at the item's location-that is, its estimated difficulty. But also note that each item provides information about a range of values along the latent variable.

Just as item characteristic curves are summed to arrive at a test characteristic curve, the item information curves are summed to arrive at a test information curve:
(5) $I_{T}(\Theta)=\sum_{i=1}^{n} I_{i}(\Theta)$.

In Equation 5, $I_{T}(\Theta)$ is the total test information evaluated at $\Theta$. To put this in more familiar terms, the test information is inversely related to the standard error of the estimated value of $\Theta$ for respondent $j$ :
(6) $\operatorname{SE}\left(\Theta_{j}\right)=\frac{1}{\sqrt{I_{T}\left(\Theta_{j}\right)}}$.

Note: As explained in the text, the exact scaling of the horizontal axis is arbitrary. This figure is centered at 0 for ease of exposition.

Note: As explained in the text, the exact scaling of the horizontal axis is arbitrary. This figure is centered at 0 for ease of exposition.

In other words, when a test provides more information about a respondent's ability, that respondent's ability (scale score) is estimated more precisely.

Figures A. 8 and A. 9 show the test information curves for the hypothetical tests depicted in Figures A. 6 and A.7. The curve shown in Figure A. 9 has a higher maximum than the curve shown in Figure A. 8 ( 0.96 for the curve shown in Figure A.9, compared to 0.85 for the Figure A. 8 curve). This is a result of having the three middle test items much more tightly clustered in the Figure A. 9 example than in the Figure A. 8 example: the Figure A. 9 test provides more information about that part of the ability scale where the test items are clustered.

In both examples, it is clear that the amount of information a test provides about a respondent is a function of that respondent's ability relative to the difficulty of the test items. For example, the test information curve shown in Figure A. 8 illustrates how much more information the five-item test provides about a respondent with an ability of 0.1 (where the test information is close to its maximum value of 0.85 ) than a respondent with an ability of 3.6 (where the test information is only 0.25). In general, the standard error of the estimate of a respondent's scale score is inversely related to the number of test items (longer tests yield more precise estimates than shorter tests) and how well the difficulty of the test items match the ability of the respondent (questions from a second-grade reading test would yield highly imprecise estimates of sixth-grade reading ability).

The test information curves for the two versions of the food security test are shown in Figure A.10. The information curve for the 18 -item test administered to households with children is above the curve for the 10 -item test over the entire range of the food insecurity scale: the 18 -item test yields estimates of household food insecurity that are substantially more precise than the 10 -item test. The curves also show that the 18 -item test attains its maximum information at a more extreme level of food insecurity than does the 10 -item test. This happens because the items that ask about


Note: As explained in the text, the exact scaling of the horizontal axis is arbitrary. This figure is centered at 0 for ease of exposition. 0
FIGUREA. 9
TEST INFORMATION CURVE
Sum of Item Information Curves
5 items, 3 Very Close Together
(
Note: As explained in the text, the exact scaling of the horizontal axis is arbitrary. This figure is centered at 0 for ease of exposition.


Note: As explained in the text, the exact scaling of the horizontal axis is arbitrary. This figure is centered at 0 for ease of exposition.
children have severity levels that are higher, on average, than the items that do not ask specifically about children.

## Mapping Raw Scores to Scale Scores

Another aspect of the Rasch model to note is that, except for respondents who did not answer some of the questions, all respondents with the same raw score are assigned the same scale score, regardless of which items they answered positively. ${ }^{12}$ This one-to-one mapping between raw scores and respondent scale scores with the Rasch model means that, when all respondents take the same test and all answer the same questions, the number of correct responses completely determines the scale score assigned to a respondent. Even though the scale scores appear to be drawn from a continuous, interval-level variable, only a small number of actual scores are assigned when the test has a small number of items. If a test has only five items, all respondents answering all the questions (except for those at the extremes with either zero or five correct answers) will be assigned one of four different scale scores because model-based scale scores can be assigned only for those respondents with raw scores from $1,2,3$, or $4 .{ }^{13}$

## Equating Scores from Different Tests

The use of test characteristic curves is the key to scoring respondents on a common scale when they take different tests. Once the difficulty of each test item, relative to every other item, is known,

[^18]test characteristic curves can be constructed for any combination of those items. Each test characteristic curve translates raw scores-the number of current responses-into scale scores for respondents. Because all test items are calibrated on the same scale, the resulting respondent scores are also on a common scale, even when the tests taken by different respondents have few test items in common. ${ }^{14}$

In the current application, there are basically two different versions of the test: (1) an 18-item test administered to households with children, and (2) a 10-item test administered to households without children. ${ }^{15}$ Figure A. 10 shows the two resulting test characteristic curves that are used for assigning household scores based on households' responses to the food insecurity questions. Using these curves, a raw score of 9 is mapped to a scale score of about 0.0 for households with children, but is mapped to a scale score of about 2.0 for households without children. Using the upper curve in Figure A.10, households with children are assigned 1 of 17 different scale scores. Using the lower curve, childless households are assigned 1 of 9 different scale scores.

[^19]
## 3. ESTIMATION OF RESPONDENT AND ITEM PARAMETERS

Up to this point, the discussion has been about the properties of the IRT model, assuming that the model parameters (or at least the item parameters) were already known. In many applications, the model parameters-both the respondent parameters and the item parameters-need to be estimated. In this section, we briefly describe how those parameters are estimated. Specialized software for estimating these models is readily available and easily used. ${ }^{16}$

Several approaches have been devised for estimating the parameters of the IRT models used here, all of which are different variations of maximum likelihood estimators. The goal in maximum likelihood estimation is to choose the values of the model parameters that maximize the likelihood of yielding the observed sample. For example, if a trait has a uniform distribution in the population, then the maximum likelihood estimate for the population mean value is the mean value observed in the sample itself. Any other estimate of the population mean (such as the mean from some other sample, the median from the current sample, or even the mean from a census) would result in a lower likelihood of observing the current sample. One important point to notice in this example is that it begins by assuming a specific distribution for the trait in the population (in this case, the trait is assumed to be uniformly distributed). This is a strong assumption; in many applications, results are not robust to different distributional assumptions that usually cannot be fully tested. An assumption about the underlying distribution is always a necessary starting point in maximum likelihood

[^20]estimation. Maximum likelihood estimators are often not unbiased. ${ }^{17}$ However, Rasch model parameters estimated using conditional likelihood procedures are unbiased, consistent, and efficient.

The likelihood expression for the Rasch model uses Equation 1 as its starting point. Equation 1 expresses the probability of a current response by a specific individual to a single test item as a function of that item's difficulty and that person's ability. If $U_{i j}$ is the response (and equals 1 for a correct response, 0 for an incorrect response), then, from Equation 1, we have:

$$
\begin{gather*}
P\left(U_{i j}=1 \mid b_{j}\right)=\frac{e^{\left(\omega_{j} u_{i}\right)}}{1+e^{\left(b_{j}-d_{i}\right)}}=P_{i j}  \tag{7}\\
P\left(U_{i:}=0 \mid b_{i}\right)=\left(1-P_{i:}\right)=Q_{i: \cdot}
\end{gather*}
$$

Thus, we represent the probability of a correct response by respondent $j$ to item $i$ as $P_{i j}$, and the probability of an incorrect response as $Q_{i j}$. If there are five test items, person $j$ 's responses might be something like $(1,1,0,1,0)$. The associated probabilities would be $\left(P_{l j}, P_{2 j}, Q_{3 j,} P_{4 j, \text { and }} Q_{5 j}\right)$. The probability of this specific string of responses would then be the product of the five probabilities-that is, $\operatorname{Pr}(1,1,0,1,0)=P_{l j} P_{2 j} Q_{3 j} P_{4 j} Q_{5 j}$. Generalizing from this example, the likelihood of a sample is the product of the probabilities across all sample members of each person's response string. The likelihood function, then, has the following form:
(8) $L=\prod_{i} P_{i j}^{\cup_{i j}} Q_{i j}^{i{ }^{{ }^{{ }_{i j}}} \text {, }}$

[^21]where $P$ and $Q$ are defined in Equation 7.
The product over $i$ gives the probability of person $i$ 's response string. The product over $j$ multiplies the response string probabilities for everybody who took the test. Rather than maximize Equation 8 directly, the $\log$ of the likelihood function is generally used:
(9) $\ln (L)=\sum_{j i}\left[U_{i j} \ln \left(P_{i j}\right)\right]+\left[\left(1-U_{i j}\right) \ln \left(Q_{i j}\right)\right]$.

The goal, then, is to find the values of $b_{j}$ and $d_{i}$ that maximize the expression in Equation 9 . Generally, the problem is broken into two parts, so that item parameters $\left(d_{i}\right)$ are first estimated (using either marginal or conditional maximum likelihood); then, using those estimated item parameters, the respondent parameters are estimated (using a numerical procedure equivalent to translating raw scores into scale scores, using the test characteristic function). Joint estimation of the item and person parameters is also an option and may be statistically more efficient than the two- stage procedure just described. All these estimation strategies are easily implemented with the current generation of IRT software.

## APPENDIX B

METHOD FOR SCREENING HOUSEHOLDS INTO THE FOOD SECURITY ANALYSIS

## METHODS FOR SCREENING HOUSEHOLDS INTO THE FOOD SECURITY ANALYSIS

The food security supplements in the 1995-1997 CPS had two general sections. The first section gathered information about food expenditures, participation in several programs aimed at providing food to needy families (for example, food stamps and school meal programs), and the sufficiency of food eaten during the preceding 12 months. The second section gathered more-detailed information about food security and coping behaviors during the previous 12 months and prior 30 days. Not all households were asked this second set of questions, which includes the questions used to construct the food security scale. In order to minimize respondent burden, households who, on the basis of earlier questions, appeared to have a high likelihood of being food secure were excluded from the more detailed questions and were implicitly assumed to be food secure in the analysis.

The screening procedures used to determine which households were asked the second set of questions were different in each of the three CPS samples (1995-1997) analyzed for this report; this section describes the analytic techniques we have developed to put these samples on a consistent basis for analysis purposes. We begin in the next section by providing background information. Then we describe the screening procedures used in each of these three CPS samples and identify a "greatest common denominator" screen. The latter is an analysis "screen" that can be applied uniformly across all the samples, to ensure that the samples used in the analytic work are comparable. (A standard screening procedure, which as been used consistently in all subsequent years, was adopted in 1998.)

Parts of this material overlap with information provided in Chapter IV. However, the discussion here provided additional detail not presented in that earlier, summary treatment.

## 1. Background

## Measuring Income in the CPS Core

As will be described below, the food security supplement screening procedures were based partly on income. Therefore, we begin by describing the income data used.

Income is measured in the CPS core questionnaire, not the food security supplement. The income measure that is a part of the CPS core is quite limited. The only question about income is one that allows only categorical responses, is asked only of the household reference person, and is asked only when the household rotates into the CPS sample (in month 1 and in month 5 of the CPS sample rotation). For all other households, information from the most recent response is carried forward to the current interview month. Income information for replacement households ${ }^{1}$ is collected when the household first appears in the CPS sample. The income question is:

$$
\begin{aligned}
& <1>\text { Less than } \$ 5,000 \\
& <2>\$ 5,000 \text { to } \$ 7,499 \\
& <3>\$ 7,500 \text { to } \$ 9,999 \\
& <4>\$ 10,000 \text { to } \$ 12,499 \\
& <5>\$ 12,500 \text { to } \$ 14,999 \\
& <6>\$ 15,000 \text { to } \$ 19,999 \\
& <7>\$ 20,000 \text { to } \$ 24,999 \\
& <8>\$ 25,000 \text { to } \$ 29,999 \\
& <9>\$ 30,000 \text { to } \$ 34,999 \\
& <10>\$ 35,000 \text { to } \$ 39,999 \\
& <11>\$ 40,000 \text { to } \$ 49,999 \\
& <12>\$ 50,000 \text { to } \$ 59,999 \\
& <13>\$ 60,000 \text { to } \$ 74,999 \\
& <14>\$ 75,000 \text { or more }
\end{aligned}
$$

(I am going to read a list of income categories.) Which category represents (your/name of reference person/the total combined income) (total combined income during the past 12 months/of all members of this FAMILY during the past 12 months)? This includes money from jobs, net income from business, farm or rent, pensions, dividends, interest, social security payments, and any other money income received (by members of this FAMILY who are 15 years of age or older).
${ }^{1}$ The CPS is designed to visit the same addresses over time, not the same people. When the members of a household move to a new address, they leave the CPS sample (unless they happen to move into another CPS address). The new residents of the CPS address are a replacement household, and are interviewed for that address's remaining, planned interviews.

## Defining "Poor" Households in the Food Security Supplement

Screening in the 1995-1997 data depended, in part, on whether households were "poor," as defined by the criteria set up for the screening. Each version of the food security supplement applied different screening criteria to households categorized as "poor" and those categorized as "nonpoor." Here we describe the definitions used. Table B. 1 shows the thresholds used to define households as poor. All the supplements used the same set of thresholds. Table B. 1 also shows the threshold levels as a percentage of the DHHS Poverty Guidelines published in 1994, 1995, and 1996. ${ }^{2}$ Within each year, the thresholds across households of different sizes, up through six people, are nearly constant multiples of the poverty guidelines. For the small number of households with 7,9 , or $10+$ members, much higher thresholds were used.

Because thresholds set at constant nominal dollars were used over this period, the criteria used to classify households as "poor" were increasingly stringent (in real terms) over time, as shown in the decreasing percentages across rows of Table B.1. Because the categories for the CPS income variable are fixed at the same nominal levels across the three CPS samples, it is not possible to apply an income screen at a constant real value.

Even with the limitations noted, however, it is a fair approximation to describe the "poor" (or low-income) households as those reporting incomes less than about 200 percent of the poverty guideline.

## Measuring Food Sufficiency

Several different versions of the basic food sufficiency questions were administered, but the general form of them is the same. All ask about household experiences over the prior 12 months and

[^22]TABLE B. 1
THRESHOLDS USED TO CLASSIFY HOUSEHOLDS AS "POOR" AND DHHS POVERTY GUIDELINES AT TIME OF SURVEY

|  |  | Threshold as a Percent of Poverty Guideline |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
|  | "Poor" if Family <br> Income Below <br> (in Dollars): | 1994 | 1995 | 1996 | 1997 |
| 1 | 15,000 | 204 | 201 | 194 | 190 |
| 2 | 20,000 | 203 | 199 | 193 | 189 |
| 3 | 25,000 | 203 | 199 | 193 | 188 |
| 4 | 30,000 | 203 | 198 | 192 | 187 |
| 5 | 35,000 | 203 | 198 | 192 | 186 |
| 6 | 40,000 | 202 | 197 | 192 | 186 |
| 7 | 50,000 | 225 | 219 | 213 | 207 |
| 8 | 50,000 | 202 | 197 | 192 | 186 |
| 9 | 60,000 | 221 | 215 | 209 | 202 |
| $10+$ | 75,000 | 253 | 246 | 239 | 232 |
| Average |  | 212 | 207 | 201 | 195 |
| Average for Households |  | 203 | 199 | 193 | 188 |
| With 6 or Fewer Persons |  |  |  |  |  |

ask the respondent (the household reference person) to choose the statement that best describes the household's experience. The response categories are:

1. Enough of the kinds of food we want to eat
2. Enough but not always the kinds of food we want to eat
3. Sometimes not enough to eat
4. Often not enough to eat
5. Don't know
6. Refused
7. No response

In the following discussion of screening procedures, we will refer to response 2 as "Not always the kind of food we want." Responses 3 and 4 will be referred to jointly as "Sometimes or often not enough to eat."

Two other related items were also used in the screening process. One item asks:
"In the last 12 months, did you ever run short of money and try to make your food or your money go further?"

We will refer to this item as "Ran short of money." The second item asks:
"In the last 12 months, did you ever run out of the foods that you needed to make a meal and didn't have enough money to get more?"

We will refer to this item as "Ran out of food."

Across the 1995-1997 supplements, these items were combined at various times and with various groups of households to create two screens:

## Screen 1

C Sometimes or often not enough to eat; or
C Ran out of food; or
C Not always the kind of food we want and ran short of money.

## Screen 2

C Sometimes or often not enough to eat; or
C Ran out of food; or
C Not always the kind of food we want; or
C Ran short of money.

Note that the screens are identical except that, under Screen 1, it is necessary to both not eat the kind of food we want and run out of money, whereas either condition is sufficient under Screen 2. Thus, Screen 2 is more liberal, allowing more households to pass, while Screen 1 is more restrictive.

## 2. Application of the Screens

Table B. 2 summarizes the screening procedures used in each of the CPS samples:
TABLE B. 2
CRITERIA USED IN 1995-1997 SURVEYS TO SCREEN HOUSEHOLDS INTO THE FOOD SECURITY MODULE

|  | "Poor" Households | "Nonpoor" Households |
| :--- | :--- | :--- |
| April 1995 | All "poor" households entered the <br> sample. | Screen 1 |
| September <br> 1996 | Screen 2 | Rotation groups 1, 2, 3, 5, 6, 7: <br> Screen 1 |
| April 1997 | Screen 2 | Rotation groups 4 and 8: Screen 2. <br>  |
|  | Rotation groups 1, 5, and 7: Screen 1. <br> Rotation groups 2, 4, 6, and 8: Screen 2. |  |
|  | Rotation group 3 used an incorrect <br> variant of Screen 1.3 |  |

Note: See text for definitions of screens.
${ }^{3}$ Households reporting any one of the following passed the (incorrect) screen: 1. Sometimes not enough to eat (incorrectly omitting "often not enough to eat") ; or 2. Ran out of food in last 12 months; or 3. Enough to eat (instead of "Not always the right kind of food") and ran short of money.

The procedures were designed to screen out households that showed no indication of any food insufficiency. A household passing the applicable screen was asked the remaining questions in the food security supplement. In effect, to avoid the food-security questions a household had to say "No" to all screening questions. Households that had missing data for the screening items (including "Don't know" or "Refused" responses) passed the screen.

## 3. A Consistent Screen for 1995-1997

Applying consistent screening procedures across all the CPS samples is important in two different contexts. First, to the extent that estimated item parameters differ across the three CPS samples, it is important to know whether those differences reflect changes, over time, in how the items are functioning or differences in the composition of the population responding to those items. ${ }^{4}$ Second, to the extent that estimates of the prevalence of food insecurity and hunger differ across the CPS samples, it is important to know whether those changes reflect real changes, over time, in prevalence or differences in the size (and composition) of the population for which the condition was measured.

To allow consistency in the analysis, we have developed an algorithm for screening households that can be applied uniformly across the 1995-1997 CPS samples. All households failing to pass the screen are categorized as food secure and are assigned the same score as households that answered "No" to all the applicable food-security questions. The screen is the following:
"Pass into the sample if
(a) household is poor and passes Screen 2, or
${ }^{4}$ The concern here is with changes in the composition of the responding population that are an artifact of the screening procedures used, rather than those that reflect real changes in population composition.
(b) household is not poor and passes Screen 1"

## 4. A "Tighter," Consistent Screen

The common screen described above utilizes the CPS income data to determine exactly which screening criteria to apply to a case. However, as noted in Section 1, there are significant limitations in the quality of these income data. Therefore, a different, considerably "tighter" screen was considered in parts of the analysis, which consisted of applying "Screen 1" regardless of income. However, this screen has not been used in the main parts of the analysis because of the loss of sample size involved in applying it.

## 5. Impacts on Prevalence Estimates

Relative to the other assumptions, conventions, and choices of analysis procedures discussed in this report, the screening conventions described above tend to have relatively substantial impacts on estimates of food security prevalence. Table B. 3 compares the basic prevalence estimates based on the common screen presented earlier with "full sample" estimates, based on all the available information. Since, in certain cases, the imposition of the common screen essentially involves automatically coding the food security answers to be negative (that is, to indicate no evidence of food insecurity), the imposition of the common screen must, by definition, lead to higher estimated prevalences of food security. As shown in the table, estimated rates of food security are reduced by approximately 1 to 1.5 percentage point when the maximum samples are used rather than the samples defined by the common denominator.

How the differences in prevalence should be interpreted depends on one's view of the accuracy of the information contained in the cases that were screened out. Since these screened-out cases are

TABLE B. 3
FOOD SECURITY PREVALENCE ESTIMATES WITH ALTERNATIVE DEFINITIONS OF THE ANALYSIS SAMPLE

Using Alternative Household Screening Conventions (Entries are Percentages)

|  |  | Food Insecure |  |  |
| :--- | :---: | :---: | :---: | :---: |
|  | Food <br> Secure | All | Without <br> Hunger | With <br> Hunger |
| $\mathbf{1 9 9 5}$ |  |  |  |  |
| Common Denominator Sources | 89.6 | 10.3 | 6.4 | 3.9 |
| Maximum Sample | 88.1 | 11.8 | 7.6 | 4.2 |
|  |  |  |  |  |
| 1996 |  |  |  |  |
| Common Denominator Sources | 89.5 | 10.5 | 6.4 | 4.1 |
| Maximum Sample | 88.6 | 11.3 | 7.0 | 4.3 |
| 1997 |  |  |  |  |
| Common Denominator Sources | 91.2 | 8.8 | 5.7 | 3.1 |
| Maximum Sample | 90.3 | 9.6 | 6.4 | 3.2 |

households with relatively higher income and/or little evidence of food insufficiency, it is possible that some of the households that answered affirmatively to 3 or more of the detailed 18 items of food security information may, for some reason, have been confused or gave misleading answers. In this case, using the lower food security estimates, based on the common screen, may be appropriate. On the other hand, to the extent that the answers given by these screened-out households are accurate, the "full sample" estimates may provide a better picture of food security in America. Our decision to focus on the estimates based on the common screen was based principally on the importance of having a consistent measure of food security over time. It does not reflect any assessment of which of the two sets of estimates more accurately reflects the underlying phenomenon of interest.

## APPENDIX C

IMPUTING MISSING FOOD SECURITY DATA

## IMPUTING MISSING FOOD SECURITY DATA

More than 99 percent of the households in the 1995-1997 CPS data sets have complete responses to the questions used to construct the Food Security Index. ${ }^{1}$ The remaining households, however, must either be dropped from the analysis or have their missing items imputed. This section discusses two alternative approaches to dealing with item nonresponse and describes the approach used in the current study.

## 1. Alternative 1: Using the Rasch Model to Impute Missing Items

As long as at least one relevant question has been answered, the Rasch model is capable of estimating a food security level for a household with missing data, as part of its maximum likelihood estimation approach. However, a potential drawback of using an imputation approach based on the Rasch model is that, when the number of unanswered questions is large, the standard error of the estimated household food security level tends to become large; and, in some cases, the estimated household score can lack face validity, in terms of the questions on which they are based. Hamilton et al. (1997) dealt with this problem by estimating scores only for those households with at least half the relevant items answered (9 out of 18 items for households with children, and 5 out of 10 items for households without children).

Another drawback of using the Rasch model for imputation is that it fails to use all the available data. Under quite general assumptions, it is reasonable to believe that imputation accuracy can be improved by drawing both on information about the answers to the food-security questions

[^23]with valid answers and on available data on household characteristics. The Rasch approach ignores the latter source of information. ${ }^{2}$

Another issue in using the Rasch model to estimate household scores for households with missing data is that the approach requires the user to have available software for estimating the Rasch model. While such software is relatively easy to obtain, this may present an obstacle to some users who do not already have this software up and running for other purposes.

## 2. Alternative 2: An Algorithm Based on Modality Assumptions for Imputing Missing Data

An alternative approach to dealing with missing responses to food security items is to impute responses outside the Rasch model, then treat the households as complete data households when assigning scores and categories. This method--which, with one exception, noted below, is the one used in the present study--has the advantages: (a) it can be specified so as to ensure that imputations have face validity; and (b) it can be implemented without Rasch software, thus increasing its accessibility. Below, we describe the algorithm we have used, first in intuitive terms and then more formally.

The basic imputation approach relies on the notion that the most sensible response patterns are those that are "modal." That is, if the questions are ordered by their severity levels, households with any affirmative responses tend to answer affirmatively to the items with the lowest severity levels and then at some point begin answering negatively to the more severe items. In this context, our imputation method assumes that, as a general rule, the highest affirmative answer indicates the household's true level of food insecurity, and the method fills in missing data below that affirmative with other affirmatives and missing data above the highest affirmative with negatives. There is,

[^24]however, an important exception to this general rule. If there is a negative answer below the highest affirmative answer, then the imputation takes this as suggesting a lower overall level of food insecurity than that implied by the highest affirmative. In the latter case, the algorithm selects the lowest negative response and fills in affirmatives for missing data below that negative response and negatives above it.

A more formal statement of this algorithm was presented in Chapter IV of the text of the report as Exhibit IV.1.

## 3. Details of Procedures Used in Current Analysis

In general, in the work summarized in this report, we have used the imputation method outlined immediately above, in Subsection 2. This method is substantially easier to implement than the Rasch-based approach, and the differences in analysis results are negligible, given (a) the small number of cases involved, and (b) that both methods tend to yield similar classifications of households.

There is one exception, however. In the 1997 sample, roughly a quarter of the cases were incorrectly asked one of the questions. For these cases, we first implemented the imputation algorithm in Section 2 above, ignoring the question that had been asked incorrectly (Question 35; "Respondent was hungry but didn't eat"). We then used Rasch model software to calculate continuous food security scores for those households, using the standard item calibrations utilized throughout this study. These numerical scores were then translated into food security categories, based on the table relating numbers of answers to continuous scores and food security categories. ${ }^{3}$

[^25]
## APPENDIX D

HOW "MODAL" ARE HOUSEHOLD RESPONSE PATTERNS?

## HOW "MODAL" ARE HOUSEHOLD RESPONSE PATTERNS?

The Rasch model implies that many households will exhibit item response patterns that are reasonably "modal," in the sense that if a household answers "yes" to any of the items, it will tend to answer "yes" to the least severe items first, and then answer "no" to all of the more severe items. A household that exhibits this pattern exactly—a string of all "yes" answers followed by a string of all "no" answers--is said to be a "modal" household. There is nothing in the Rasch theory that predicts that all households will be modal; (indeed, the probabilistic Rasch model cannot be estimated if all households are exactly modal). Still, it is of interest in understanding the data to examine the degree of modality that is present. A large number of strongly nonmodal response patterns could call into question the validity of the model.

## A. NUMBER OF RESPONSE CHANGES TO MAKE HOUSEHOLD MODEL.

One approach to examining the degree to which households exhibit modal answers is to calculate for each household the minimum number of answers that would have to be different in order to make the household responses be modal. Of course, if the household's answers are already perfectly modal, then the number of answers that would have to be changed is zero. However, consider, as an example, a household with the following response pattern: three "yes" answers, then a "no," then a "yes," then all "no's." For such a household, only one item (the first "no" or the last "yes") would have had to be changed to make the response pattern modal. Similarly, to take a second example, suppose a household has two "no" answers, then four "yes" answers, then all "no," answers. It would require at least two changes (the first two "no" answers) to make the household modal.

Table D. 1 tabulates the minimum number of answer changes required to make the households in the 1997 sample modal. It shows that 39 percent of the households in the 1997 data are perfectly modal. For another 36 percent there is only one discrepancy between their scores and a modal pattern. Fifteen percent have two such discrepancies, while 9 percent have three or more. Overall, this suggests substantial adherence to modal response patterns. ${ }^{1}$

Households without children appear to exhibit greater modality than those with children. However, this may be due to the fact that there are many fewer questions applicable to the group without children (10, rather than 18 , for households with children), and hence fewer opportunities for nonmodality.

Table D. 2 presents a more detailed look at these issues, focusing on households with children. The central section of the table shows, for each possible number of "yes" answers, the frequency distribution of the highest item (in terms of severity) to which the nonmodal households with children answered "yes." For instance, the fifth row shows that, when one considers the 233 nonmodal households that gave 5 "yes" answers, for 57 of them, the highest "yes" answer was on Item 6, while for another 44, the highest "yes" answer was Item 7, and so on.

The shading in the table reflects the fact that certain cell entries are logically impossible-if, for instance, there are five "yes" answers, the highest nonmodal item with a "yes" answer cannot come before the sixth item. To the extent that the nonmodal households are "almost" modal, we would expect households to be clustered just to the right of the shaded area. For instance, using a
${ }^{1}$ Of course, for a household with just one nonmodal answer, that answer could be severely nonmodal (a "no" at the beginning of a long string of "yes" answers, or a "yes" many items after the previous "yes"). We examine the severity of nonmodality shortly.

TABLE D. 1

## PERCENT OF HOUSEHOLDS BY THE MINIMUM NUMBER OF NONMODAL RESPONSES TO THE FOOD SECURITY ITEMS <br> (Percentages; 1997 Data)

| Number of Nonmodal Responses ${ }^{\text {a }}$ | All Households | Households Without Children | Households With Children |
| :---: | :---: | :---: | :---: |
| 0 | 39 | 50 | 30 |
| 1 | 36 | 38 | 33 |
| 2 | 15 | 9 | 21 |
| 3 | 6 | 2 | 9 |
| 4 | 2 | 0 | 4 |
| 5 | 1 | 0 | 2 |
| $6+$ | 0 | 0 | 1 |
| Total | 100 | 100 | 100 |

${ }^{a}$ Minimum number of responses that would have to be changed for the household to be modal.
TABLE D. 2
ANALYSIS OF MODALITY BY NUMBER OF YES RESPONSES TO FOOD SECURITY ITEMS

|  | ALL |  | MODAL |  | NONMODAL - HIGHEST YES RESPONSE ITEM |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | NONMODAL |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\begin{aligned} & \text { Number } \\ & \text { of Yes } \end{aligned}$ | HH | \% | HH | \% | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13 | 14 | 15 | 16 | 17 | 18 | HH | \% |
| 1 | 600 | 20.4\% | 367 | 42.5\% | 56 | 69 | 82 | 7 | 7 | 11 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 233 | 11.2\% |
| 2 | 485 | 16.5\% | 204 | 23.6\% |  | 77 | 111 | 22 | 16 | 38 | 4 | 4 | 4 | 1 | 3 | 1 | 0 | 0 | 0 | 0 | 0 | 281 | 13.5\% |
| 3 | 372 | 12.6\% | 104 | 12.0\% |  |  | 117 | 30 | 29 | 54 | 15 | 11 | 6 | 2 | 2 | 1 | 1 | 0 | 0 | 0 | 0 | 268 | 12.9\% |
| 4 | 320 | 10.9\% | 72 | 8.3\% |  |  |  | 18 | 71 | 75 | 42 | 15 | 9 | 5 | 7 | 2 | 2 | 0 | 0 | 1 | 1 | 248 | 11.9\% |
| 5 | 241 | 8.2\% | 8 | 0.9\% |  |  |  |  | 57 | 44 | 48 | 32 | 20 | 12 | 6 | 10 | 0 | 2 | 0 | 2 | 0 | 233 | 11.2\% |
| 6 | 236 | 8.0\% | 8 | 0.9\% |  |  |  |  |  | 35 | 48 | 98 | 14 | 9 | 9 | 3 | 3 | 6 | 1 | 2 | 0 | 228 | 11.0\% |
| 7 | 178 | 6.0\% | 10 | 1.2\% |  |  |  |  |  |  | 39 | 49 | 30 | 22 | 6 | 8 | 5 | 5 | 3 | 1 | 0 | 168 | 8.1\% |
| 8 | 132 | 4.5\% | 33 | 3.8\% |  |  |  |  |  |  |  | 20 | 22 | 24 | 10 | 10 | 4 | 7 | 0 | 2 | 0 | 99 | 4.8\% |
| 9 | 129 | 4.4\% | 32 | 3.7\% |  |  |  |  |  |  |  |  | 21 | 17 | 15 | 4 | 12 | 19 | 4 | 4 | 1 | 97 | 4.7\% |
| 10 | 68 | 2.3\% | 10 | 1.2\% |  |  |  |  |  |  |  |  |  | 4 | 14 | 2 | 15 | 16 | 2 | 5 | 0 | 58 | 2.8\% |
| 11 | 62 | 2.1\% | 7 | 0.8\% |  |  |  |  |  |  |  |  |  |  | 8 | 2 | 14 | 15 | 9 | 5 | 2 | 55 | 2.6\% |
| 12 | 44 | 1.5\% | 1 | 0.1\% |  |  |  |  |  |  |  |  |  |  |  | 1 | 10 | 16 | 5 | 11 | 0 | 43 | 2.1\% |
| 13 | 28 | 1.0\% | 0 | 0.0\% |  |  |  |  |  |  |  |  |  |  |  |  | 5 | 12 | 1 | 8 | 2 | 28 | 1.3\% |
| 14 | 22 | 0.7\% | 0 | 0.0\% |  |  |  |  |  |  |  |  |  |  |  |  |  | 7 | 2 | 10 | 3 | 22 | 1.1\% |
| 15 | 14 | 0.5\% | 5 | 0.6\% |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 1 | 5 | 3 | 9 | 0.4\% |
| 16 | 8 | 0.3\% | 0 | 0.0\% |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 4 | 4 | 8 | 0.4\% |
| 17 | 2 | 0.1\% |  | 0.0\% |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 2 | 2 | 0.1\% |
| 18 | 3 | 0.1\% | 3 | 0.3\% |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 0 | 0.0\% |
| TOTAL | 2,944 | 100.0\% | 864 | 100.0\% | 56 | 146 | 310 | 77 | 180 | 257 | 196 | 229 | 127 | 96 | 80 | 44 | 71 | 105 | 28 | 60 | 18 | 2,080 | 100.0\% |
|  |  |  |  |  |  |  |  |  |  |  | 9.4\% |  |  | 4.6\% | 3.8\% | 2.1\% | 3.4\% |  | 1.3\% |  | 0.9\% |  |  |

previous example, a household with five "yes" answers that has the sixth item as its highest "yes" answer has only one nonmodal answer in its overall string of answers. For most rows in the table, nonmodal households do cluster near the shaded diagonal, suggesting that the nonmodal response patterns are not severely nonmodal. About 47 percent of nonmodal households are in the first two off-diagonal cells, and an additional 22 percent are in the third cell.

## B. BOUNDS ON THE EFFECTS OF NONMODALITY ON PREVALENCE ESTIMATES.

A useful way to understand the implications of nonmodal response patterns is to assess their effects on prevalence estimates. Accordingly, in this section, we calculate for each household the minimum and maximum food-insecurity levels that can be obtained by making the household's response pattern modal (Table D.3). To obtain the minimum insecurity level, we classify a household based on the items before its first "no" answer. This effectively converts all higher "yes" answers to "no" answers, giving a modal pattern. To obtain the maximum insecurity level, we classify a household based on the items up to and including its last "yes" answer. This effectively converts all lower "no" answers to "yes" answers, again giving a modal pattern.

Consider, as an example, a household with children that answers "yes" to the first two items, "no" to the third and fourth items, "yes" to the next four items, and "no" to the last ten items (for six "yes" answers in all). For this household, the minimum food insecurity level is based on the modal pattern of "yes" to the first 2 items and "no" to the last 16 items. The maximum food insecurity level is based on the modal pattern of "yes" to the first 8 items and "no" to the last 10 items. ${ }^{2}$ Likewise,

[^26]TABLE D. 3

> MINIMUM AND MAXIMUM FOOD INSECURITY PREVALENCE ESTIMATES
> (Percent of Households)

| Food Security Status | Base <br> Estimate | Estimate Involving <br> Minimum Estimate of Food Insecurity | Estimate Involving Maximum Estimate of Food Insecurity |
| :---: | :---: | :---: | :---: |
| 1995 |  |  |  |
| Food Secure | 89.6 | 90.47 | 87.30 |
| Food Insecure, Hunger Not Evident | 6.4 | 6.67 | 6.03 |
| Food Insecure, Hunger Evident | 3.9 | 2.66 | 6.66 |
| Food Security Status Not Determined | 0.1 | 0.20 |  |
| 1996 |  |  |  |
| Food Secure | 89.5 | 90.50 | 87.22 |
| Food Insecure, Hunger Not Evident | 6.4 | 6.55 | 6.14 |
| Food Insecure, Hunger Evident | 4.1 | 2.69 | 6.64 |
| Food Security Status Not Determined | 0.1 | 0.26 |  |
| 1997 |  |  |  |
| Food Secure | 91.2 | 92.00 | 88.21 |
| Food Insecure, Hunger Not Evident | 5.7 | 5.75 | 4.01 |
| Food Insecure, Hunger Evident | 3.1 | 2.09 | 7.79 |
| Food Security Status Not Determined | 0.1 | 0.16 |  |

NOTES: To compute the maximum estimate of food insecurity, households were classified based on the most severe item with "yes" response.

To compute the minimum estimate of food insecurity, households were classified based on the most severe "yes" item preceding the least severe "no" response.

Percentages may not add to 100 due to rounding.
when every household is classified at its maximum food insecurity level, the overall prevalence of insecurity is at a maximum. In each of the three years, going to the "minimum insecurity" scenario tends to raise the proportion of households that are classified as food secure by about 1 percentage point compared to the "base" estimates, and there is a decrease in the proportion classified as experiencing hunger of between 1 and 2 percentage points. Going to the scenario with "maximum insecurity" raises the proportion with hunger by between 2 and 5 percentage points, depending on the year.

Interestingly, the scenario that involves the minimum food insecurity category causes the proportion of households in the middle category--food insecure without hunger--to go up in each of the three years. This is because more households move into this category from the most severe category than leave it to enter the food secure category.

## APPENDIX E

## ITEM SEVERITY LEVELS BY SUBGROUPS

## ITEM- SEVERITY LEVELS, BY SUBGROUPS

In addition to examining the stability of the model parameters over time, it is important to examine model stability for key subgroups in the population, in order to test whether the data support the aggregation of various components of the population into a single model, as has been done so far in the discussion. Here, we report findings from estimating models for subsets of households classified by (1) race/ethnicity; (2) household composition, including the presence of children; (3) metropolitan versus nonmetropolitan locations; and (4) region.

Our overall conclusions are that the results obtained from the Rasch model are reasonably robust when examined by subgroup. The next section presents our results for the race/ethnicity classification in some detail; we then more briefly review the findings from the other groupings. We focus on the race/ethnicity classification because of its intrinsic importance and because we anticipated prior to the analysis that the differences between groups might be most pronounced with this classification.

## 1. Findings When Households Are Classified by Race

Table E. 1 presents the raw item scores for 1997 for the 18 questions used in the food security scale—first, for all households, then for each of four race/ethnicity subgroups. For purposes of comparison, these scores have been normalized, using the standard procedures summarized in Chapter II of the main report. They are estimated based on the 1997 CPS data, using the common screen.

Table E. 2 summarizes the order of severity in which each item enters the model, both for the sample as a whole and for each of the subgroups, defined by ethnicity. Table E. 3 shows the

TABLE E. 1

THRESHOLD PARAMETER ESTIMATES FOR 18-ITEM 12-MONTH FOOD SECURITY SCALE BY RACE/ETHNIC GROUP

April 1997 CPS Data, Common Denominator Screen

| Item | Model |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Constrained |  | Unconstrained |  |  |  |  |  |  |  |
|  | All Households |  | White |  | Black |  | Hispanic |  | Other |  |
|  | Estimate | Std Err | Estimate | Std Err | Estimate | Std Err | Estimate | Std Err | Estimate | Std Err |
| 50 | 11.284 | 0.245 | 11.599 | 0.561 | 10.805 | 0.684 | 11.527 | 0.815 | 12.372 | 2.244 |
| 44 | 9.873 | 0.187 | 10.301 | 0.232 | 9.833 | 0.287 | 9.301 | 0.431 | 10.000 | 0.725 |
| 43 | 9.293 | 0.170 | 9.647 | 0.183 | 9.075 | 0.186 | 8.890 | 0.392 | 10.000 | 0.725 |
| 29 | 8.999 | 0.073 | 8.735 | 0.106 | 9.363 | 0.137 | 9.196 | 0.259 | 9.013 | 0.466 |
| 47 | 8.957 | 0.111 | 8.870 | 0.139 | 9.025 | 0.237 | 9.186 | 0.233 | 8.346 | 0.350 |
| 28 | 8.516 | 0.072 | 8.286 | 0.115 | 8.716 | 0.106 | 8.899 | 0.165 | 8.182 | 0.342 |
| 40 | 8.564 | 0.087 | 8.754 | 0.133 | 8.526 | 0.232 | 8.125 | 0.120 | 9.470 | 0.724 |
| 38 | 8.168 | 0.087 | 7.852 | 0.108 | 8.474 | 0.201 | 8.688 | 0.232 | 7.958 | 0.394 |
| 35 | 7.297 | 0.060 | 6.859 | 0.097 | 7.871 | 0.140 | 7.911 | 0.155 | 6.955 | 0.244 |
| 57 | 6.915 | 0.094 | 7.354 | 0.093 | 6.701 | 0.168 | 6.213 | 0.235 | 6.818 | 0.327 |
| 25 | 6.463 | 0.054 | 6.279 | 0.079 | 6.577 | 0.104 | 6.692 | 0.129 | 6.223 | 0.252 |
| 32 | 5.663 | 0.036 | 5.437 | 0.075 | 5.804 | 0.108 | 5.943 | 0.144 | 5.542 | 0.258 |
| 56 | 5.818 | 0.075 | 5.941 | 0.102 | 5.763 | 0.163 | 5.478 | 0.228 | 5.478 | 0.304 |
| 24 | 5.586 | 0.047 | 5.316 | 0.070 | 5.756 | 0.107 | 5.996 | 0.133 | 5.400 | 0.303 |
| 58 | 4.522 | 0.057 | 4.548 | 0.066 | 4.419 | 0.139 | 4.316 | 0.257 | 4.664 | 0.346 |
| 55 | 4.079 | 0.066 | 4.059 | 0.097 | 4.064 | 0.152 | 3.875 | 0.172 | 3.700 | 0.404 |
| 54 | 3.700 | 0.070 | 3.847 | 0.075 | 3.277 | 0.124 | 3.365 | 0.294 | 3.791 | 0.430 |
| 53 | 2.306 | 0.087 | 2.316 | 0.112 | 1.952 | 0.170 | 2.400 | 0.244 | 2.092 | 0.699 |
| Mean | 7.000 |  | 7.000 |  | 7.000 |  | 7.000 |  | 7.000 |  |
| Std Dev | 2.413 |  | 2.481 |  | 2.475 |  | 2.473 |  | 2.651 |  |
| Sample size | 5,429 |  | 3,347 |  | 1,079 |  | 779 |  | 225 |  |

TABLE E. 2
ITEM ORDERING FOR RACE/ETHNIC GROUPS AND FULL POPULATION

|  | Question | Full <br> Population | White | Black | Hispanic | Other |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| q50 | Child not eat whole day | 1 | 1 | 1 | 1 | 1 |
| q44 | Child skipped meal, 3+ months | 2 | 2 | 2 | 2 | 2 |
| q43 | Child skipped meal | 3 | 3 | 4 | 6 | 2 |
| q29 | Adult not eat whole day, 3+ months | 4 | 6 | 3 | 3 | 5 |
| q47 | Child hungry | 5 | 4 | 5 | 4 | 6 |
| q28 | Adult not eat whole day | 7 | 7 | 6 | 5 | 7 |
| q40 | Cut size of child's meals | 6 | 5 | 7 | 8 | 4 |
| q38 | Adult lost weight | 8 | 8 | 8 | 7 | 8 |
| q35 | Adult hungry but didn't eat | 9 | 10 | 9 | 9 | 9 |
| q57 | Child not eating enough | 10 | 9 | 10 | 11 | 10 |
| q25 | Adult cut size or skipped meals, in 3+ months | 11 | 11 | 11 | 10 | 11 |
| q32 | Adult eat less than felt they should | 13 | 13 | 12 | 13 | 12 |
| q56 | Couldn't feed child balanced meals | 12 | 12 | 13 | 14 | 13 |
| q24 | Adult cut size or skipped meals | 14 | 14 | 14 | 12 | 14 |
| q58 | Adult fed child few low cost foods | 15 | 15 | 15 | 15 | 15 |
| q55 | Adult not eat balanced meals | 16 | 16 | 16 | 16 | 17 |
| q54 | Food bought didn't last | 17 | 17 | 17 | 17 | 16 |
| q53 | Worried food would run out | 18 | 18 | 18 | 18 | 18 |

Based on weighted 1997 data using BilogMG calibrations.

TABLE E. 3

## CROSS-GROUP DIFFERENCES IN THRESHOLD PARAMETER ESTIMATES BY RACE/ETHNIC GROUP

April 1997 CPS Data, Common Denominator Screen

| Variable | Black - White |  |  | Hispanic - White |  |  | Other - White |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Estimate | Std Err | $\operatorname{Pr}(\|T\|>\|t\|)$ | Estimate | Std Err | $\operatorname{Pr}(\|\mathrm{T}\|>\mid \mathrm{t})$ | Estimate | Std Err | $\operatorname{Pr}(\|\mathrm{T}\|>\|\mathrm{t}\|)$ |
| 50 | 0.794 | 0.731 | 0.278 | 0.072 | 1.023 | 0.944 | -0.773 | 2.391 | 0.747 |
| 44 | 0.468 | 0.354 | 0.187 | 1.000 | 0.503 | 0.047 | 0.302 | 0.744 | 0.685 |
| 43 | 0.573 | 0.209 | 0.006 | 0.757 | 0.396 | 0.056 | -0.352 | 0.693 | 0.611 |
| 29 | -0.628 | 0.188 | 0.001 | -0.460 | 0.283 | 0.104 | -0.278 | 0.424 | 0.512 |
| 47 | -0.155 | 0.290 | 0.594 | -0.316 | 0.214 | 0.141 | 0.524 | 0.366 | 0.153 |
| 28 | -0.430 | 0.161 | 0.008 | -0.613 | 0.213 | 0.004 | 0.104 | 0.325 | 0.750 |
| 40 | 0.229 | 0.246 | 0.352 | 0.629 | 0.204 | 0.002 | -0.715 | 0.763 | 0.349 |
| 38 | -0.622 | 0.221 | 0.005 | -0.837 | 0.243 | 0.001 | -0.107 | 0.430 | 0.804 |
| 35 | -1.012 | 0.179 | 0.000 | -1.052 | 0.189 | 0.000 | -0.096 | 0.272 | 0.725 |
| 57 | 0.654 | 0.183 | 0.000 | 1.141 | 0.195 | 0.000 | 0.536 | 0.329 | 0.104 |
| 25 | -0.298 | 0.114 | 0.009 | -0.413 | 0.147 | 0.005 | 0.056 | 0.290 | 0.846 |
| 32 | -0.367 | 0.141 | 0.009 | -0.506 | 0.183 | 0.006 | -0.105 | 0.290 | 0.718 |
| 56 | 0.178 | 0.172 | 0.300 | 0.463 | 0.261 | 0.077 | 0.464 | 0.325 | 0.154 |
| 24 | -0.441 | 0.130 | 0.001 | -0.680 | 0.132 | 0.000 | -0.084 | 0.305 | 0.783 |
| 58 | 0.129 | 0.147 | 0.381 | 0.232 | 0.252 | 0.357 | -0.116 | 0.356 | 0.745 |
| 55 | -0.005 | 0.153 | 0.975 | 0.185 | 0.195 | 0.344 | 0.360 | 0.424 | 0.397 |
| 54 | 0.569 | 0.129 | 0.000 | 0.482 | 0.269 | 0.074 | 0.056 | 0.424 | 0.895 |
| 53 | 0.364 | 0.188 | 0.053 | -0.084 | 0.233 | 0.717 | 0.224 | 0.686 | 0.744 |

difference between estimated parameters, as well as estimates of the statistical significance of the differences.

When the item orderings for Black households are compared to those for Whites, half of the items remain in exactly the same locations. Most of the remaining items differ in order between the two groups by only one or two places. (For instance, Item 40 is ordered at position 15 in the analysis for the White sample and position 17 for the Black sample.) Only one item, Item 29, changes by more than two places. Similar patterns holds when comparisons are made between the White sample and the Hispanic sample.

As shown in Table E.3, about half the Black-White differences and more than half of the Hispanic-White differences are statistically significant. However, in general, the differences are relatively modest.

Figure E. 1 provides a visual representation of the calibration of each item for the full sample, as plotted against each subgroup calibration. For instance, the lowest point designated with a triangle indicates that, for the least severe item ( Item 53) the Black subgroup parameter for that item is slightly below the calibration of the full group or the other subgroups. ${ }^{1}$ The closeness with which the lines in the figure track each other provides an additional way of confirming that the separately estimated item parameters for the different racial/ethnic groups are very similar to each another.

Despite the heuristic appeal of both the tabular and the graphic comparisons of the item calibrations across groups, what really matters, we believe, is whether differences of the type and size we are observing are likely to have any substantial effects on the overall food security measurement process. Given the complexity of the issues underlying this question, there appears

[^27]FIGURE E. 1
ITEM CALIBRATIONS FOR 3 RACE/ETHNIC GROUPS
vs. Full Population Item Calibrations
April 1997
CPS, 12-Month Scale

$\rightarrow$ All - White $\rightarrow$ Black $\rightarrow$ Hispanic
not to be a definitive, formal statistical test available to answer it. Therefore, we have appealed to the judgment of the two members of the project Advisory Committee, Benjamin Wright and Robert Mislevy, who have extensive experience using this type of model in a broad range of applications. Both have expressed the view that differences across subgroups of the magnitude shown in the above tables are not large enough to seriously interfere with the measurement process that we are pursuing.

Therefore, we have concluded that, on balance, the results of our analysis by race/ethnicity subgroup tend to support the overall modeling exercise being undertaken.

## 2. Other Findings

The rest of the tables in this section present comparable analysis results for several other subgroup classifications. Three tables and a figure are presented for each. In general, the observed differences across subgroups are either similar to, or smaller than, the differences for race/ethnicity.

Household composition. When a similar analysis is done, based on household composition, the item order for the subgroups track very closely with each other, with no item changing its order by more than one place (Tables E.4, E.5, E.6, and Figure E.2). In general, as shown in Table E.4, the estimated absolute values of the items for households without children tends to be higher than those for households with children. However, this arises from the arbitrary scaling convention of ensuring a mean of 7 for the items, coupled with the fact that the 10 items applicable to households without children happen, on average to be of lower than average severity. ${ }^{2}$

Metropolitan status. When the analysis is disaggregated according the metropolitan versus nonmetropolitan residence, ten of the items stay in the same positions across the two relevant

[^28]TABLE E. 4
THRESHOLD PARAMETER ESTIMATES FOR 12-MONTH FOOD SECURITY SCALE BY HOUSEHOLD TYPE

April 1997 CPS Data, Common Denominator Screen

| Item | Model |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Constrained |  | Unconstrained |  |  |  |  |  |
|  | All Households |  | Kids |  | Elderly |  | Others |  |
|  | Estimate | Std Err | Estimate | Std Err | Estimate | Std Err | Estimate | Std Err |
| 50 | 11.284 | 0.245 | 11.211 | 0.246 |  |  |  |  |
| 44 | 9.873 | 0.187 | 9.819 | 0.184 |  |  |  |  |
| 43 | 9.293 | 0.170 | 9.247 | 0.168 |  |  |  |  |
| 29 | 8.999 | 0.073 | 8.926 | 0.099 | 10.103 | 0.161 | 10.070 | 0.132 |
| 47 | 8.957 | 0.111 | 8.915 | 0.111 |  |  |  |  |
| 28 | 8.516 | 0.072 | 8.478 | 0.096 | 9.753 | 0.160 | 9.481 | 0.113 |
| 40 | 8.564 | 0.087 | 8.528 | 0.086 |  |  |  |  |
| 38 | 8.168 | 0.087 | 8.328 | 0.145 | 9.401 | 0.168 | 8.861 | 0.090 |
| 35 | 7.297 | 0.060 | 7.233 | 0.082 | 8.763 | 0.135 | 8.157 | 0.077 |
| 57 | 6.915 | 0.094 | 6.904 | 0.096 |  |  |  |  |
| 25 | 6.463 | 0.054 | 6.598 | 0.066 | 7.188 | 0.122 | 7.273 | 0.062 |
| 32 | 5.663 | 0.036 | 5.720 | 0.046 | 6.416 | 0.070 | 6.560 | 0.104 |
| 56 | 5.818 | 0.075 | 5.818 | 0.078 |  |  |  |  |
| 24 | 5.586 | 0.047 | 5.690 | 0.062 | 6.649 | 0.111 | 6.250 | 0.066 |
| 58 | 4.522 | 0.057 | 4.530 | 0.061 |  |  |  |  |
| 55 | 4.079 | 0.066 | 4.323 | 0.070 | 4.244 | 0.191 | 4.860 | 0.088 |
| 54 | 3.700 | 0.070 | 3.715 | 0.094 | 4.390 | 0.144 | 4.675 | 0.095 |
| 53 | 2.306 | 0.087 | 2.017 | 0.102 | 3.092 | 0.164 | 3.813 | 0.108 |
| Mean | 7.000 |  | 7.000 |  | 7.000 |  | 7.000 |  |
| Std Dev | 2.413 |  | 2.404 |  | 2.500 |  | 2.144 |  |
| Sample size | 5,429 |  | 2,944 |  | 874 |  | 1,611 |  |

Notes: Standard errors have been estimated using a balanced replication procedure to account for the complex sample design of the CPS .

TABLE E. 5
ITEM ORDERING FOR THREE HOUSEHOLD TYPES AND FULL POPULATION
April 1997 CPS, 12-Month Scale

|  | Question | Full <br> Population | Kids | Elderly | Others |
| :---: | :---: | :---: | :---: | :---: | :---: |
| q50 | Child not eat whole day | 1 | 1 |  |  |
| q44 | Child skipped meal, 3+ months | 2 | 2 |  |  |
| q43 | Child skipped meal | 3 | 3 |  |  |
| q29 | Adult not eat whole day, 3+ months | 4 | 4 | 1 | 1 |
| q47 | Child hungry | 5 | 5 |  |  |
| q28 | Adult not eat whole day | 7 | 7 | 2 | 2 |
| q40 | Cut size of child's meals | 6 | 6 |  |  |
| q38 | Adult lost weight | 8 | 8 | 3 | 3 |
| q35 | Adult hungry but didn't eat | 9 | 9 | 4 | 4 |
| q57 | Child not eating enough | 10 | 10 |  |  |
| q25 | Adult cut size or skipped meals, in 3+ months | 11 | 11 | 5 | 5 |
| q32 | Adult eat less than felt they should | 13 | 13 | 7 | 6 |
| q56 | Couldn't feed child balanced meals | 12 | 12 |  |  |
| q24 | Adult cut size or skipped meals | 14 | 14 | 6 | 7 |
| q58 | Adult fed child few low cost foods | 15 | 15 |  |  |
| q55 | Adult not eat balanced meals | 16 | 16 | 9 | 8 |
| q54 | Food bought didn't last | 17 | 17 | 8 | 9 |
| q53 | Worried food would run out | 18 | 18 | 10 | 10 |

Based on weighted 1997 data using BilogMG calibrations.

TABLE E. 6

## CROSS-GROUP DIFFERENCES IN THRESHOLD PARAMETER ESTIMATES BY HOUSEHOLD TYPE

April 1997 CPS Data, Common Denominator Screen

| Variable | Kids-Elderly |  |  | Kids-Others |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Estimate | Std Err | $\operatorname{Pr}(\|\mathrm{T}\|>\|\mathrm{t}\|)$ | Estimate | Std Err | $\operatorname{Pr}(\|\mathrm{T}\|>\|\mathrm{t}\|)$ |
| 50 |  |  |  |  |  |  |
| 44 |  |  |  |  |  |  |
| 43 |  |  |  |  |  |  |
| 29 | -0.033 | 0.193 | 0.865 | 0.000 | 0.171 | 0.999 |
| 47 |  |  |  |  |  |  |
| 28 | -0.168 | 0.161 | 0.298 | 0.103 | 0.155 | 0.507 |
| 40 |  |  |  |  |  |  |
| 38 | -0.004 | 0.205 | 0.986 | 0.537 | 0.153 | 0.000 |
| 35 | -0.565 | 0.154 | 0.000 | 0.041 | 0.112 | 0.713 |
| 57 |  |  |  |  |  |  |
| 25 | 0.328 | 0.126 | 0.009 | 0.242 | 0.094 | 0.010 |
| 32 | 0.150 | 0.082 | 0.068 | 0.007 | 0.108 | 0.949 |
| 56 |  |  |  |  |  |  |
| 24 | -0.114 | 0.111 | 0.304 | 0.285 | 0.091 | 0.002 |
| 58 |  |  |  |  |  |  |
| 55 | 0.832 | 0.178 | 0.000 | 0.215 | 0.121 | 0.076 |
| 54 | 0.039 | 0.191 | 0.839 | -0.244 | 0.099 | 0.013 |
| 53 | -0.466 | 0.175 | 0.008 | -1.187 | 0.121 | 0.000 |

FIGURE E. 2
ITEM CALIBRATIONS FOR HOUSEHOLD TYPE GROUPS
vs. Full Population Item Calibrations
April 1997
CPS, 12-Month Scale

subgroups, and no change in order is more than two positions (Tables E.7, E.8, and E.9, and Figure E.3).

Region of Country. When the analysis is done by region of the country, no item switches by more than a one position (Tables E.10, E.11, E.12, and Figure E.4).

Summary. Overall, these findings suggest that there is considerable robustness in the model, when it is estimated for various population subgroups.

TABLE E. 7

THRESHOLD PARAMETER ESTIMATES FOR 18-ITEM 12-MONTH FOOD SECURITY SCALE BY METROPOLITAN STATUS

April 1997 CPS Data, Common Denominator Screen

| Item | Model |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Constrained |  | Unconstrained |  |  |  |
|  |  |  | Metro |  | NonMetro |  |
|  | Estimate | Std Err | Estimate | Std Err | Estimate | Std Err |
| 50 | 11.284 | 0.245 | 11.087 | 0.369 | 11.751 | 1.017 |
| 44 | 9.873 | 0.187 | 9.742 | 0.241 | 10.721 | 0.611 |
| 43 | 9.293 | 0.170 | 9.145 | 0.190 | 10.087 | 0.478 |
| 29 | 8.999 | 0.073 | 9.123 | 0.124 | 8.772 | 0.257 |
| 47 | 8.957 | 0.111 | 8.954 | 0.166 | 9.074 | 0.272 |
| 28 | 8.516 | 0.072 | 8.612 | 0.097 | 8.227 | 0.220 |
| 40 | 8.564 | 0.087 | 8.523 | 0.101 | 8.961 | 0.284 |
| 38 | 8.168 | 0.087 | 8.113 | 0.120 | 8.355 | 0.176 |
| 35 | 7.297 | 0.060 | 7.342 | 0.080 | 7.036 | 0.203 |
| 57 | 6.915 | 0.094 | 6.858 | 0.129 | 6.813 | 0.129 |
| 25 | 6.463 | 0.054 | 6.430 | 0.070 | 6.441 | 0.125 |
| 32 | 5.663 | 0.036 | 5.724 | 0.053 | 5.486 | 0.192 |
| 56 | 5.818 | 0.075 | 5.800 | 0.099 | 5.533 | 0.179 |
| 24 | 5.586 | 0.047 | 5.541 | 0.060 | 5.733 | 0.157 |
| 58 | 4.522 | 0.057 | 4.553 | 0.078 | 4.188 | 0.273 |
| 55 | 4.079 | 0.066 | 4.141 | 0.079 | 3.721 | 0.302 |
| 54 | 3.700 | 0.070 | 3.766 | 0.097 | 3.415 | 0.306 |
| 53 | 2.306 | 0.087 | 2.549 | 0.109 | 1.687 | 0.443 |
| Mean | 7.000 |  | 7.000 |  | 7.000 |  |
| Std Dev | 2.413 |  | 2.344 |  | 2.730 |  |
| Sample size | 5,429 |  | 3,140 |  | 1,493 |  |

Notes: Standard errors have been estimated using a balanced replication procedure to account fo the complex sample design of the CPS.

TABLE E. 8
ITEM ORDERING BY METROPOLITAN STATUS AND FULL POPULATION
April 1997 CPS, 12-Month Scale

|  |  |  |  |  |
| :--- | :--- | :---: | :---: | :---: |
|  | Question | Full <br> Population | Metro | NonMetro |
| q50 | Child not eat whole day | 1 | 1 | 1 |
| q44 | Child skipped meal, 3+ months | 2 | 2 | 2 |
| q43 | Child skipped meal | 3 | 3 | 3 |
| q29 | Adult not eat whole day, 3+ months | 4 | 4 | 6 |
| q47 | Child hungry | 5 | 5 | 4 |
| q28 | Adult not eat whole day | 7 | 6 | 8 |
| q40 | Cut size of child's meals | 6 | 7 | 5 |
| q38 | Adult lost weight | 8 | 8 | 7 |
| q35 | Adult hungry but didn't eat | 9 | 9 | 9 |
| q57 | Child not eating enough | 10 | 10 | 10 |
| q25 | Adult cut size or skipped meals, in 3+ months | 11 | 11 | 11 |
| q32 | Adult eat less than felt they should | 13 | 13 | 14 |
| q56 | Couldn't feed child balanced meals | 12 | 12 | 13 |
| q24 | Adult cut size or skipped meals | 14 | 14 | 12 |
| q58 | Adult fed child few low cost foods | 15 | 15 | 15 |
| q55 | Adult not eat balanced meals | 16 | 16 | 16 |
| q54 | Food bought didn't last | 17 | 17 | 17 |
| q53 | Worried food would run out | 18 | 18 | 18 |
|  |  |  |  |  |

Based on weighted 1997 data using BilogMG calibrations.

TABLE E. 9

## CROSS-GROUP DIFFERENCES IN THRESHOLD PARAMETER ESTIMATES BY METROPOLITAN STATUS

April 1997 CPS Data, Common Denominator Screen

|  | NonMetro-Metro |  |  |
| :---: | :---: | :---: | :---: |
| Variable | Estimate | Std Err | $\operatorname{Pr}(\|\mathrm{T}\|>\|\mathrm{t}\|)$ |
| 50 | -0.664 | 1.057 | 0.530 |
| 44 | -0.979 | 0.646 | 0.130 |
| 43 | -0.943 | 0.508 | 0.064 |
| 29 | 0.352 | 0.324 | 0.279 |
| 47 | -0.120 | 0.336 | 0.721 |
| 28 | 0.384 | 0.245 | 0.118 |
| 40 | -0.438 | 0.302 | 0.147 |
| 38 | -0.242 | 0.256 | 0.344 |
| 35 | 0.306 | 0.237 | 0.197 |
| 57 | 0.045 | 0.175 | 0.795 |
| 25 | -0.011 | 0.139 | 0.936 |
| 32 | 0.238 | 0.213 | 0.265 |
| 56 | 0.267 | 0.212 | 0.208 |
| 24 | -0.192 | 0.164 | 0.241 |
| 58 | 0.365 | 0.290 | 0.209 |
| 55 | 0.420 | 0.309 | 0.174 |
| 54 | 0.350 | 0.310 | 0.259 |
| 53 | 0.861 | 0.446 | 0.054 |

FIGURE E. 3
ITEM CALIBRATIONS BY METROPOLITAN STATUS vs. Full Population Item Calibrations
April 1997 CPS, 12-Month Scale

122

TABLE E. 10
THRESHOLD PARAMETER ESTIMATES FOR 18-ITEM 12-MONTH FOOD SECURITY SCALE BY REGION

April 1997 CPS Data, Common Denominator Screen

| Item | Model |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Constrained |  | Unconstrained |  |  |  |  |  |  |  |
|  | All Households |  | Northeast |  | Midwest |  | South |  | West |  |
|  | Estimate | Std Err | Estimate | Std Err | Estimate | Std Err | Estimate | Std Err | Estimate | Std Err |
| 50 | 11.284 | 0.245 | 10.666 | 0.522 | 12.397 | 1.662 | 11.197 | 0.466 | 11.670 | 0.285 |
| 44 | 9.873 | 0.187 | 9.955 | 0.381 | 9.484 | 0.265 | 10.326 | 0.222 | 9.538 | 0.621 |
| 43 | 9.293 | 0.170 | 9.619 | 0.443 | 9.305 | 0.237 | 9.407 | 0.153 | 8.888 | 0.501 |
| 29 | 8.999 | 0.073 | 8.892 | 0.176 | 9.066 | 0.205 | 8.988 | 0.156 | 9.017 | 0.073 |
| 47 | 8.957 | 0.111 | 9.018 | 0.357 | 8.710 | 0.185 | 9.128 | 0.158 | 8.850 | 0.294 |
| 28 | 8.516 | 0.072 | 8.532 | 0.210 | 8.580 | 0.215 | 8.429 | 0.118 | 8.548 | 0.097 |
| 40 | 8.564 | 0.087 | 8.801 | 0.216 | 8.514 | 0.251 | 8.698 | 0.143 | 8.226 | 0.148 |
| 38 | 8.168 | 0.087 | 7.895 | 0.254 | 8.369 | 0.130 | 8.285 | 0.166 | 8.000 | 0.086 |
| 35 | 7.297 | 0.060 | 7.282 | 0.146 | 7.060 | 0.163 | 7.463 | 0.106 | 7.220 | 0.109 |
| 57 | 6.915 | 0.094 | 6.922 | 0.199 | 7.000 | 0.147 | 6.767 | 0.066 | 7.003 | 0.333 |
| 25 | 6.463 | 0.054 | 6.309 | 0.106 | 6.449 | 0.126 | 6.499 | 0.095 | 6.475 | 0.148 |
| 32 | 5.663 | 0.036 | 5.505 | 0.070 | 5.579 | 0.213 | 5.711 | 0.069 | 5.707 | 0.080 |
| 56 | 5.818 | 0.075 | 5.887 | 0.163 | 5.675 | 0.211 | 5.769 | 0.134 | 5.895 | 0.152 |
| 24 | 5.586 | 0.047 | 5.512 | 0.084 | 5.462 | 0.210 | 5.668 | 0.076 | 5.551 | 0.142 |
| 58 | 4.522 | 0.057 | 4.499 | 0.146 | 4.375 | 0.334 | 4.460 | 0.102 | 4.668 | 0.116 |
| 55 | 4.079 | 0.066 | 4.361 | 0.174 | 3.992 | 0.375 | 3.899 | 0.071 | 4.132 | 0.168 |
| 54 | 3.700 | 0.070 | 3.822 | 0.095 | 3.695 | 0.389 | 3.437 | 0.065 | 3.918 | 0.221 |
| 53 | 2.306 | 0.087 | 2.523 | 0.171 | 2.289 | 0.509 | 1.871 | 0.106 | 2.695 | 0.242 |
| Mean | 7.000 |  | 7.000 |  | 7.000 |  | 7.000 |  | 7.000 |  |
| Std Dev | 2.413 |  | 2.335 |  | 2.543 |  | 2.544 |  | 2.316 |  |
| Sample size | 5,429 |  | 1,023 |  | 1,136 |  | 1,821 |  | 1,449 |  |

Notes:
Standard errors have been estimated using a balanced replication procedure to account for the complex sample design of the CPS.

TABLE E. 11
ITEM ORDERING FOR 4 REGIONS AND FULL POPULATION April 1997 CPS, 12-Month Scale

|  |  |  |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | Question | Full <br> Population | Northeast | Midwest | South | West |
| q50 | Child not eat whole day | 1 | 1 | 1 | 1 | 1 |
| q44 | Child skipped meal, 3+ months | 2 | 2 | 2 | 2 | 2 |
| q43 | Child skipped meal | 3 | 3 | 3 | 3 | 4 |
| q29 | Adult not eat whole day, 3+ months | 4 | 5 | 4 | 5 | 3 |
| q47 | Child hungry | 5 | 4 | 5 | 4 | 5 |
| q28 | Adult not eat whole day | 7 | 7 | 6 | 7 | 6 |
| q40 | Cut size of child's meals | 6 | 6 | 7 | 6 | 7 |
| q38 | Adult lost weight | 8 | 8 | 8 | 8 | 8 |
| q35 | Adult hungry but didn't eat | 9 | 9 | 9 | 9 | 9 |
| q57 | Child not eating enough | 10 | 10 | 10 | 10 | 10 |
| q25 | Adult cut size or skipped meals, in 3+ months | 11 | 11 | 11 | 11 | 11 |
| q32 | Adult eat less than felt they should | 13 | 14 | 13 | 13 | 13 |
| q56 | Couldn't feed child balanced meals | 12 | 12 | 12 | 12 | 12 |
| q24 | Adult cut size or skipped meals | 14 | 13 | 14 | 14 | 14 |
| q58 | Adult fed child few low cost foods | 15 | 15 | 15 | 15 | 15 |
| q55 | Adult not eat balanced meals | 16 | 16 | 16 | 16 | 16 |
| q54 | Food bought didn't last | 17 | 17 | 17 | 17 | 17 |
| q53 | Worried food would run out | 18 | 18 | 18 | 18 | 18 |
|  |  |  |  |  |  |  |

Based on weighted 1997 data using BilogMG calibrations.

TABLE E. 12

## CROSS-GROUP DIFFERENCES IN THRESHOLD PARAMETER ESTIMATES BY REGION

Y REGION
April 1997 CPS Data, Common Denominator Screen

| Variable | Northeast-Midwest |  |  | Northeast-South |  |  | Northeast-West |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Estimate | Std Err | $\operatorname{Pr}(\|\mathrm{T}\|>\|\mathrm{t}\|)$ | Estimate | Std Err | $\operatorname{Pr}(\|\mathrm{T}\|>\|\mathrm{t}\|)$ | Estimate | Std Err | $\operatorname{Pr}(\|\mathrm{T}\|>\mid \mathrm{t})$ |
| 50 | -1.731 | 1.763 | 0.326 | -0.531 | 0.688 | 0.440 | -1.004 | 0.577 | 0.082 |
| 44 | 0.471 | 0.457 | 0.303 | -0.370 | 0.431 | 0.391 | 0.417 | 0.737 | 0.572 |
| 43 | 0.314 | 0.491 | 0.523 | 0.211 | 0.469 | 0.652 | 0.730 | 0.675 | 0.279 |
| 29 | -0.174 | 0.276 | 0.528 | -0.096 | 0.238 | 0.687 | -0.125 | 0.192 | 0.517 |
| 47 | 0.308 | 0.397 | 0.438 | -0.110 | 0.394 | 0.780 | 0.168 | 0.457 | 0.713 |
| 28 | -0.048 | 0.306 | 0.875 | 0.104 | 0.242 | 0.669 | -0.016 | 0.233 | 0.945 |
| 40 | 0.287 | 0.331 | 0.386 | 0.103 | 0.260 | 0.691 | 0.575 | 0.258 | 0.026 |
| 38 | -0.474 | 0.279 | 0.090 | -0.390 | 0.306 | 0.203 | -0.105 | 0.262 | 0.689 |
| 35 | 0.222 | 0.223 | 0.320 | -0.181 | 0.179 | 0.312 | 0.062 | 0.175 | 0.723 |
| 57 | -0.078 | 0.254 | 0.760 | 0.156 | 0.209 | 0.456 | -0.080 | 0.390 | 0.837 |
| 25 | -0.140 | 0.170 | 0.412 | -0.189 | 0.143 | 0.184 | -0.166 | 0.181 | 0.360 |
| 32 | -0.074 | 0.231 | 0.750 | -0.205 | 0.095 | 0.031 | -0.202 | 0.097 | 0.038 |
| 56 | 0.212 | 0.272 | 0.436 | 0.118 | 0.210 | 0.575 | -0.008 | 0.225 | 0.972 |
| 24 | 0.050 | 0.234 | 0.831 | -0.156 | 0.111 | 0.160 | -0.039 | 0.162 | 0.811 |
| 58 | 0.124 | 0.371 | 0.739 | 0.039 | 0.177 | 0.827 | -0.169 | 0.185 | 0.361 |
| 55 | 0.369 | 0.423 | 0.384 | 0.462 | 0.185 | 0.013 | 0.228 | 0.240 | 0.342 |
| 54 | 0.128 | 0.419 | 0.761 | 0.385 | 0.110 | 0.001 | -0.096 | 0.240 | 0.690 |
| 53 | 0.235 | 0.561 | 0.676 | 0.653 | 0.199 | 0.001 | -0.172 | 0.294 | 0.559 |



## APPENDIX F

PREVALENCE ESTIMATES AND THEIR STANDARD ERRORS

This appendix presents detailed food security prevalence tables and the corresponding standard error estimates for the 1995-1997 period. The estimates are disaggregated by year and income group. The exact prevalence estimates may, in some instances, vary slightly from those issued in the "Advance Report" which was issued by USDA in July 1999, with preliminary findings. These differences are due to additional editing and are not substantial.

Totals of food secure and food insecure cases do not quite add to population totals, because there were a very small number of households for which food insecurity status could not be calculated. All percentages are based on using all cases, including those with missing food security data, as the denominator.

Tabulations based on metropolitan status do not add to national totals, because for many households inside metropolitan areas, the CPS data do not indicate whether or not the household is in the central city.

The data for 1995-1997 are adjusted to reflect the screening adopted in 1998 and thereafter. Essentially, this means that certain households who answered "no" to all of the less severe food security questions are coded as having missing values for the more severe items. This is analogue to what would have been their treatment in the 1998 and subsequent years, when such households were skipped out of the later questions, based on their negative answers to the earlier ones. ${ }^{1}$
${ }^{1}$ The 1995-1997 files used in the current report, like the corresponding public use files available on the CPS website, have one further adjustment which is designed to make them more comparable to the 1998 data. In the 1998 survey, under certain circumstances, households were not asked relatively severe food security questions if they had already consistently answered "no" to blocks of less severe questions. The 1995-1997 data were edited to emulate this screening by replacing answers to the more severe questions with missing value codes for questions which would have been skipped in 1998. See: www.bls.census.gov/cps/foodsecu/1997/agnote.htm ( "Food Security, Scales and Screener Variables").

The following shows the table numbers for the various tables:

|  | Year | Prevalence Table | Standard Error Table |
| :---: | :---: | :---: | :---: |
| All Income Groups | 1995 | F. 1 | F. 22 |
|  | 1996 | F. 2 | F. 23 |
|  | 1997 | F. 3 | F. 24 |
| Under 100 Percent of Poverty |  |  |  |
|  | 1995 | F. 4 | F. 25 |
|  | 1996 | F. 5 | F. 26 |
|  | 1997 | F. 6 | F. 27 |
| Over 100 Percent of Poverty |  |  |  |
|  | 1995 | F. 7 | F. 28 |
|  | 1996 | F. 8 | F. 29 |
|  | 1997 | F. 9 | F. 30 |
| Under 130 Percent of Poverty |  |  |  |
|  | 1995 | F. 10 | F. 31 |
|  | 1996 | F. 11 | F. 32 |
|  | 1997 | F. 12 | F. 33 |
| Over 130 Percent of Poverty |  |  |  |
|  | 1995 | F. 13 | F. 34 |
|  | 1996 | F. 14 | F. 35 |
|  | 1997 | F. 15 | F. 36 |
| Under 185 Percent of Poverty |  |  |  |
|  | 1995 | F. 16 | F. 37 |
|  | 1996 | F. 17 | F. 38 |
|  | 1997 | F. 18 | F. 39 |
| Over 185 Percent of Poverty |  |  |  |
|  | 1995 | F. 19 | F. 40 |
|  | 1996 | F. 20 | F. 41 |
|  | 1997 | F. 21 | F. 42 |

TABLE F. 1

## PREVALENCE OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER BY SELECTED CHARACTERSITICS OF HOUSEHOLDS, 1995 USING THE COMMON DENOMINATOR SCREEN

|  | Total (000) | Food Secure |  | Food Insecure |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  | Without | Hunger | With | nger |
|  |  | (000) | Percent | (000) | Percent | (000) | Percent | (000) | Percent |
| All Households | 100,445 | 89,969 | 89.6 | 10,392 | 10.3 | 6,437 | 6.4 | 3,955 | 3.9 |
| Household Composition |  |  |  |  |  |  |  |  |  |
| With Children < 6 | 18,353 | 15,199 | 82.8 | 3,116 | 17.0 | 2,177 | 11.9 | 939 | 5.1 |
| With Children < 18 | 38,181 | 32,230 | 84.4 | 5,911 | 15.5 | 3,993 | 10.5 | 1,918 | 5.0 |
| Married Couple Families | 27,222 | 24,460 | 89.9 | 2,743 | 10.1 | 1,992 | 7.3 | 751 | 2.8 |
| Female Head, No Spouse | 8,578 | 5,808 | 67.7 | 2,764 | 32.2 | 1,751 | 20.4 | 1,013 | 11.8 |
| Male Head, No Spouse | 1,836 | 1,535 | 83.6 | 288 | 15.8 | 201 | 11.0 | 87 | 4.8 |
| Other Households with Children | 545 | 428 | 78.7 | 116 | 21.3 | 49 | 9.0 | 67 | 12.3 |
| With No Children < 18 | 62,264 | 57,739 | 92.7 | 4,480 | 7.2 | 2,444 | 3.9 | 2,036 | 3.3 |
| More Than One Adult | 36,777 | 34,826 | 94.7 | 1,924 | 5.2 | 1,148 | 3.1 | 776 | 2.1 |
| Women Living Alone | 15,161 | 13,708 | 90.4 | 1,447 | 9.6 | 811 | 5.4 | 636 | 4.2 |
| Men Living Alone | 10,325 | 9,204 | 89.1 | 1,108 | 10.7 | 484 | 4.7 | 624 | 6.0 |
| Households with Elderly | 23,776 | 22,570 | 94.9 | 1,192 | 5.0 | 742 | 3.1 | 450 | 1.9 |
| Elderly Living Alone | 10,069 | 9,489 | 94.2 | 576 | 5.8 | 327 | 3.3 | 249 | 2.5 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |  |  |  |
| White Non-Hispanic | 78,100 | 72,266 | 92.5 | 5,799 | 7.4 | 3,618 | 4.6 | 2,181 | 2.8 |
| Black Non-Hispanic | 11,906 | 9,321 | 78.3 | 2,569 | 21.6 | 1,521 | 12.8 | 1,048 | 8.8 |
| Hispanic | 7,753 | 5,971 | 77.0 | 1,749 | 22.6 | 1,136 | 14.7 | 613 | 7.9 |
| Other Non-Hispanic | 2,686 | 2,411 | 89.8 | 273 | 10.2 | 161 | 6.0 | 112 | 4.2 |

TABLE F. 1 (continued)

|  | $\begin{aligned} & \text { Total } \\ & (000) \\ & \hline \end{aligned}$ | Food Secure |  | Food Insecure |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | All |  | Without Hunger |  | With Hunger |  |
|  |  | (000) | Percent | (000) | Percent | (000) | Percent | (000) | Percent |
| Household Income-to-Poverty Ratio |  |  |  |  |  |  |  |  |  |
| Under 0.50 | 5,603 | 3,467 | 61.9 | 2,113 | 37.7 | 1,177 | 21.0 | 936 | 16.7 |
| Under 1.00 | 15,924 | 10,717 | 67.3 | 5,154 | 32.3 | 3,081 | 19.3 | 2,073 | 13.0 |
| Under 1.30 | 21,953 | 15,480 | 70.5 | 6,413 | 29.2 | 3,823 | 17.4 | 2,590 | 11.8 |
| Under 1.85 | 32,182 | 24,030 | 74.7 | 8,080 | 25.1 | 4,935 | 15.3 | 3,145 | 9.8 |
| 1.85 and Over | 58,400 | 56,615 | 96.9 | 1,777 | 3.0 | 1,176 | 2.0 | 601 | 1.0 |
| Income Not Known | 9,863 | 9,324 | 94.5 | 534 | 5.4 | 325 | 3.3 | 209 | 2.1 |
| Area of Residence |  |  |  |  |  |  |  |  |  |
| Inside Metropolitan Area |  |  |  |  |  |  |  |  |  |
| In Central City | 24,701 | 21,168 | 85.7 | 3,525 | 14.3 | 2,149 | 8.7 | 1,376 | 5.6 |
| Not in Central City | 37,059 | 34,092 | 92.0 | 2,935 | 7.9 | 1,757 | 4.7 | 1,178 | 3.2 |
| Outside Metropolitan Area | 22,658 | 20,173 | 89.0 | 2,460 | 10.8 | 1,590 | 7.0 | 870 | 3.8 |
| Census Geographical Region |  |  |  |  |  |  |  |  |  |
| Northeast | 19,491 | 17,776 | 91.2 | 1,694 | 8.7 | 1,056 | 5.4 | 638 | 3.3 |
| Midwest | 23,656 | 21,442 | 90.6 | 2,212 | 9.3 | 1,375 | 5.8 | 837 | 3.5 |
| South | 35,891 | 31,906 | 88.9 | 3,966 | 11.1 | 2,498 | 7.0 | 1,468 | 4.1 |
| West | 21,407 | 18,845 | 88.0 | 2521 | 11.7 | 1,508 | 7.0 | 1,013 | 4.7 |

TABLE F. 2
PREVALENCE OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER BY SELECTED CHARACTERSITICS OF HOUSEHOLDS, 1996
USING THE COMMON DENOMINATOR SCREEN

|  | $\begin{array}{r} \text { Total } \\ (000) \\ \hline \end{array}$ | Food Secure |  | Food Insecure |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | All |  | Without Hunger |  | With Hunger |  |
|  |  | (000) | Percent | (000) | Percent | (000) | Percent | (000) | Percent |
| All Households | 101,508 | 90,809 | 89.5 | 10,592 | 10.5 | 6,462 | 6.4 | 4,130 | 4.1 |
| Household Composition |  |  |  |  |  |  |  |  |  |
| With Children < 6 | 17,871 | 14,703 | 82.3 | 3,127 | 17.5 | 2,176 | 12.2 | 951 | 5.3 |
| With Children < 18 | 38,196 | 32,205 | 84.3 | 5,922 | 15.5 | 3,974 | 10.4 | 1,948 | 5.1 |
| Married Couple Families | 27,187 | 24,326 | 89.5 | 2,817 | 10.4 | 1,985 | 7.3 | 832 | 3.1 |
| Female Head, No Spouse | 8,488 | 5,793 | 68.2 | 2,669 | 31.4 | 1,699 | 20.0 | 970 | 11.4 |
| Male Head, No Spouse | 1,936 | 1,604 | 82.9 | 332 | 17.1 | 219 | 11.3 | 113 | 5.8 |
| Other Households with Children | 585 | 481 | 82.3 | 104 | 17.7 | 71 | 12.1 | 33 | 5.6 |
| With No Children < 18 | 63,312 | 58,604 | 92.6 | 4,670 | 7.3 | 2,488 | 3.9 | 2,182 | 3.4 |
| More Than One Adult | 37,806 | 35,743 | 94.5 | 2,044 | 5.4 | 1,153 | 3.0 | 891 | 2.4 |
| Women Living Alone | 14,961 | 13,457 | 89.9 | 1,496 | 10.0 | 795 | 5.3 | 701 | 4.7 |
| Men Living Alone | 10,545 | 9,404 | 89.2 | 1,131 | 10.7 | 540 | 5.1 | 591 | 5.6 |
| Households with Elderly | 24,087 | 22,849 | 94.9 | 1,219 | 5.0 | 779 | 3.2 | 440 | 1.8 |
| Elderly Living Alone | 9,760 | 9,185 | 94.1 | 574 | 5.9 | 353 | 3.6 | 221 | 2.3 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |  |  |  |
| White Non-Hispanic | 77,830 | 72,056 | 92.6 | 5,737 | 7.3 | 3,460 | 4.4 | 2,277 | 2.9 |
| Black Non-Hispanic | 12,083 | 9,332 | 77.2 | 2,709 | 22.4 | 1,684 | 13.9 | 1,025 | 8.5 |
| Hispanic | 8,165 | 6,340 | 77.6 | 1,805 | 22.1 | 1,101 | 13.5 | 704 | 8.6 |
| Other Non-Hispanic | 3,430 | 3,082 | 89.8 | 342 | 9.9 | 217 | 6.3 | 125 | 3.6 |

TABLE F. 2 (continued)

|  | $\begin{aligned} & \text { Total } \\ & (000) \\ & \hline \end{aligned}$ | Food Secure |  | Food Insecure |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | All |  | Without Hunger |  | With Hunger |  |
|  |  | (000) | Percent | (000) | Percent | (000) | Percent | (000) | Percent |
| Household Income-to-Poverty Ratio |  |  |  |  |  |  |  |  |  |
| Under 0.50 | 6,442 | 3,932 | 61.0 | 2,480 | 38.5 | 1,342 | 20.8 | 1,138 | 17.7 |
| Under 1.00 | 14,548 | 9,390 | 64.5 | 5,106 | 35.1 | 2,951 | 20.3 | 2,155 | 14.8 |
| Under 1.30 | 20,282 | 13,868 | 68.4 | 6,353 | 31.4 | 3,726 | 18.4 | 2,627 | 13.0 |
| Under 1.85 | 32,814 | 24,449 | 74.5 | 8,290 | 25.3 | 5,038 | 15.4 | 3,252 | 9.9 |
| 1.85 and Over | 57,123 | 55,379 | 96.9 | 1,743 | 3.1 | 1,082 | 1.9 | 661 | 1.2 |
| Income Not Known | 11,571 | 10,982 | 94.9 | 558 | 4.8 | 341 | 2.9 | 217 | 1.9 |
| Area of Residence |  |  |  |  |  |  |  |  |  |
| Inside Metropolitan Area |  |  |  |  |  |  |  |  |  |
| In Central City | 26,262 | 22,611 | 86.1 | 3,614 | 13.7 | 2,139 | 8.1 | 1,475 | 5.6 |
| Not in Central City | 40,956 | 37,604 | 91.8 | 3,291 | 8.0 | 1,968 | 4.8 | 1,323 | 3.2 |
| Outside Metropolitan Area | 19,958 | 17,698 | 88.7 | 2,256 | 11.3 | 1,461 | 7.3 | 795 | 4.0 |
| Census Geographical Region |  |  |  |  |  |  |  |  |  |
| Northeast | 19,740 | 18,169 | 92.0 | 1,548 | 7.9 | 938 | 4.8 | 610 | 3.1 |
| Midwest | 24,132 | 22,029 | 91.3 | 2,092 | 8.7 | 1,281 | 5.3 | 811 | 3.4 |
| South | 35,792 | 31,460 | 87.9 | 4,274 | 12.0 | 2,644 | 7.4 | 1,630 | 4.6 |
| West | 21,844 | 19,151 | 87.7 | 2,677 | 12.2 | 1,598 | 7.3 | 1,079 | 4.9 |

TABLE F. 3
PREVALENCE OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER BY SELECTED CHARACTERSITICS OF HOUSEHOLDS, 1997
USING THE COMMON DENOMINATOR SCREEN

|  | $\begin{aligned} & \text { Total } \\ & (000) \\ & \hline \end{aligned}$ | Food Secure |  | Food Insecure |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | All |  | Without Hunger |  | With Hunger |  |
|  |  | (000) | Percent | (000) | Percent | (000) | Percent | (000) | Percent |
| All Households | 102,373 | 93,319 | 91.2 | 8,963 | 8.8 | 5,831 | 5.7 | 3,132 | 3.1 |
| Household Composition |  |  |  |  |  |  |  |  |  |
| With Children < 6 | 17,538 | 15,133 | 86.3 | 2,396 | 13.6 | 1,813 | 10.3 | 583 | 3.3 |
| With Children < 18 | 38,195 | 33,237 | 87.0 | 4,920 | 12.9 | 3,538 | 9.3 | 1,382 | 3.6 |
| Married Couple Families | 26,859 | 24,751 | 92.2 | 2,095 | 7.8 | 1,585 | 5.9 | 510 | 1.9 |
| Female Head, No Spouse | 8,715 | 6,264 | 71.9 | 2,433 | 27.9 | 1,653 | 19.0 | 780 | 8.9 |
| Male Head, No Spouse | 2,105 | 1,793 | 85.2 | 312 | 14.8 | 248 | 11.8 | 64 | 3.0 |
| Other Households with Children | 516 | 429 | 83.2 | 81 | 15.7 | 52 | 10.1 | 29 | 5.6 |
| With No Children < 18 | 64,178 | 60,082 | 93.6 | 4,042 | 6.3 | 2,293 | 3.6 | 1,749 | 2.7 |
| More Than One Adult | 37,788 | 36,034 | 95.4 | 1,739 | 4.6 | 1,086 | 2.9 | 653 | 1.7 |
| Women Living Alone | 15,411 | 14,144 | 91.8 | 1,245 | 8.1 | 690 | 4.5 | 555 | 3.6 |
| Men Living Alone | 10,980 | 9,904 | 90.2 | 1,059 | 9.6 | 517 | 4.7 | 542 | 4.9 |
| Households with Elderly | 24,420 | 23,278 | 95.3 | 1,125 | 4.6 | 775 | 3.2 | 350 | 1.4 |
| Elderly Living Alone | 10,323 | 9,795 | 94.9 | 517 | 5.0 | 335 | 3.2 | 182 | 1.8 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |  |  |  |
| White Non-Hispanic | 78,021 | 73,234 | 93.9 | 4,749 | 6.1 | 2,967 | 3.8 | 1,782 | 2.3 |
| Black Non-Hispanic | 12,289 | 9,966 | 81.1 | 2,296 | 18.7 | 1,500 | 12.2 | 796 | 6.5 |
| Hispanic | 8,383 | 6,737 | 80.4 | 1,624 | 19.4 | 1,174 | 14.0 | 450 | 5.4 |
| Other Non-Hispanic | 3,680 | 3,382 | 91.9 | 294 | 8.0 | 190 | 5.2 | 104 | 2.8 |

TABLE F. 3 (continued)

TABLE F. 4
PREVALENCE OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER
BY SELECTED CHARACTERSITICS OF HOUSEHOLDS, 1995
FOR THOSE FAMILIES UNDER 100 PERCENT OF POVERTY THRESHHOLD
USING THE COMMON DENOMINATOR SCREEN
Food Secure

|  | $\begin{aligned} & \text { Total } \\ & (000) \\ & \hline \end{aligned}$ | Food Secure |  | All |  | Without Hunger |  | With Hunger |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | (000) | Percent | (000) | Percent | (000) | Percent | (000) | Percent |
| All Households | 15,924 | 10,717 | 67.3 | 5,154 | 32.3 | 3,081 | 19.3 | 2,073 | 13.0 |
| Household Composition |  |  |  |  |  |  |  |  |  |
| With Children < 6 | 4,648 | 2,752 | 59.2 | 1,866 | 40.2 | 1,198 | 25.8 | 668 | 14.4 |
| With Children < 18 | 7,856 | 4,653 | 59.2 | 3,173 | 40.4 | 2,001 | 25.5 | 1,172 | 14.9 |
| Married Couple Families | 3,081 | 1,963 | 63.7 | 1,108 | 35.9 | 746 | 24.2 | 362 | 11.7 |
| Female Head, No Spouse | 4,150 | 2,312 | 55.7 | 1,831 | 44.1 | 1,125 | 27.1 | 706 | 17.0 |
| Male Head, No Spouse | 463 | 277 | 59.8 | 174 | 37.5 | 108 | 23.3 | 66 | 14.2 |
| Other Households with Children | 162 | 101 | 62.6 | 61 | 37.4 | 22 | 13.5 | 39 | 23.9 |
| With No Children < 18 | 8,067 | 6,064 | 75.2 | 1,981 | 24.6 | 1,080 | 13.4 | 901 | 11.2 |
| More Than One Adult | 3,185 | 2,436 | 76.5 | 737 | 23.2 | 432 | 13.6 | 305 | 9.6 |
| Women Living Alone | 3,274 | 2,498 | 76.3 | 773 | 23.6 | 458 | 14.0 | 315 | 9.6 |
| Men Living Alone | 1,608 | 1,130 | 70.3 | 471 | 29.3 | 190 | 11.8 | 281 | 17.5 |
| Households with Elderly | 4,094 | 3,430 | 83.8 | 663 | 16.2 | 425 | 10.4 | 238 | 5.8 |
| Elderly Living Alone | 2,432 | 2,096 | 86.2 | 335 | 13.8 | 191 | 7.9 | 144 | 5.9 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |  |  |  |
| White Non-Hispanic | 8,633 | 6,183 | 71.6 | 2,433 | 28.2 | 1,457 | 16.9 | 976 | 11.3 |
| Black Non-Hispanic | 4,178 | 2,653 | 63.5 | 1,516 | 36.3 | 857 | 20.5 | 659 | 15.8 |
| Hispanic | 2,592 | 1,500 | 57.9 | 1,068 | 41.2 | 692 | 26.7 | 376 | 14.5 |
| Other Non-Hispanic | 522 | 382 | 73.3 | 138 | 26.3 | 75 | 14.3 | 63 | 12.0 |

TABLE F. 4 (continued)

TABLE F. 5
PREVALENCE OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER
BY SELECTED CHARACTERSITICS OF HOUSEHOLDS, 1996
FOR THOSE FAMILIES UNDER 100 PERCENT OF POVERTY THRESHHOLD
USING THE COMMON DENOMINATOR SCREEN

|  | $\begin{aligned} & \text { Total } \\ & (000) \end{aligned}$ | Food Secure |  | Food Insecure |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | All |  | Without Hunger |  | With Hunger |  |
|  |  | (000) | Percent | (000) | Percent | (000) | Percent | (000) | Percent |
| All Households | 14,548 | 9,390 | 64.5 | 5,106 | 35.1 | 2,951 | 20.3 | 2,155 | 14.8 |
| Household Composition |  |  |  |  |  |  |  |  |  |
| With Children < 6 | 4,035 | 2,300 | 57.0 | 1,721 | 42.7 | 1,118 | 27.7 | 603 | 15.0 |
| With Children < 18 | 7,127 | 4,034 | 56.6 | 3,061 | 43.0 | 1,902 | 26.7 | 1,159 | 16.3 |
| Married Couple Families | 2,866 | 1,748 | 61.0 | 1,098 | 38.3 | 711 | 24.8 | 387 | 13.5 |
| Female Head, No Spouse | 3,697 | 1,959 | 53.0 | 1,726 | 46.7 | 1,034 | 28.0 | 692 | 18.7 |
| Male Head, No Spouse | 428 | 242 | 56.6 | 185 | 43.4 | 124 | 29.1 | 61 | 14.3 |
| Other Households with Children | 136 | 85 | 62.2 | 51 | 37.8 | 32 | 23.8 | 19 | 14.0 |
| With No Children < 18 | 7,422 | 5,356 | 72.2 | 2,045 | 27.5 | 1,049 | 14.1 | 996 | 13.4 |
| More Than One Adult | 3,021 | 2,277 | 75.4 | 740 | 24.4 | 394 | 13.0 | 346 | 11.4 |
| Women Living Alone | 2,943 | 2,147 | 72.9 | 790 | 26.8 | 434 | 14.7 | 356 | 12.1 |
| Men Living Alone | 1,457 | 932 | 64.0 | 515 | 35.4 | 221 | 15.2 | 294 | 20.2 |
| Households with Elderly | 3,608 | 2,931 | 81.2 | 672 | 18.6 | 416 | 11.5 | 256 | 7.1 |
| Elderly Living Alone | 2,066 | 1,738 | 84.1 | 328 | 15.9 | 206 | 10.0 | 122 | 5.9 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |  |  |  |
| White Non-Hispanic | 7,635 | 5,334 | 69.9 | 2,284 | 29.9 | 1,274 | 16.7 | 1,010 | 13.2 |
| Black Non-Hispanic | 3,772 | 2,154 | 57.1 | 1,593 | 42.3 | 938 | 24.9 | 655 | 17.4 |
| Hispanic | 2,516 | 1,457 | 57.9 | 1,052 | 41.8 | 630 | 25.0 | 422 | 16.8 |
| Other Non-Hispanic | 625 | 446 | 71.4 | 177 | 28.3 | 109 | 17.4 | 68 | 10.9 |

TABLE F. 5 (continued)

TABLE F. 6
PREVALENCE OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER
BY SELECTED CHARACTERSITICS OF HOUSEHOLDS, 1997
FOR THOSE FAMILIES UNDER 100 PERCENT OF POVERTY THRESHHOLD
USING THE COMMON DENOMINATOR SCREEN

|  | $\begin{aligned} & \text { Total } \\ & (000) \\ & \hline \end{aligned}$ | Food Secure |  | All |  | Without Hunger |  | With Hunger |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | (000) | Percent | (000) | Percent | (000) | Percent | (000) | Percent |
| All Households | 14,157 | 9,828 | 69.4 | 4,297 | 30.4 | 2,676 | 18.9 | 1,621 | 11.5 |
| Household Composition |  |  |  |  |  |  |  |  |  |
| With Children < 6 | 4,058 | 2,675 | 65.9 | 1,378 | 33.9 | 996 | 24.5 | 382 | 9.4 |
| With Children < 18 | 7,036 | 4,435 | 63.0 | 2,590 | 36.8 | 1,790 | 25.4 | 800 | 11.4 |
| Married Couple Families | 2,632 | 1,850 | 70.3 | 778 | 29.5 | 572 | 21.7 | 206 | 7.8 |
| Female Head, No Spouse | 3,819 | 2,209 | 57.8 | 1,607 | 42.0 | 1,059 | 27.7 | 548 | 14.3 |
| Male Head, No Spouse | 463 | 303 | 65.3 | 160 | 34.7 | 125 | 27.1 | 35 | 7.6 |
| Other Households with Children | 122 | 74 | 60.4 | 45 | 37.2 | 34 | 28.3 | 11 | 8.9 |
| With No Children < 18 | 7,121 | 5,392 | 75.7 | 1,708 | 23.9 | 886 | 12.4 | 822 | 11.5 |
| More Than One Adult | 2,793 | 2,147 | 76.9 | 639 | 22.9 | 348 | 12.5 | 291 | 10.4 |
| Women Living Alone | 2,759 | 2,131 | 77.2 | 620 | 22.5 | 339 | 12.3 | 281 | 10.2 |
| Men Living Alone | 1,569 | 1,114 | 71.0 | 448 | 28.5 | 198 | 12.6 | 250 | 15.9 |
| Households with Elderly | 3,326 | 2,768 | 83.2 | 556 | 16.7 | 385 | 11.6 | 171 | 5.1 |
| Elderly Living Alone | 2,023 | 1,746 | 86.3 | 276 | 13.7 | 175 | 8.7 | 101 | 5.0 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |  |  |  |
| White Non-Hispanic | 7,026 | 5,154 | 73.3 | 1,863 | 26.5 | 1,061 | 15.1 | 802 | 11.4 |
| Black Non-Hispanic | 3,810 | 2,461 | 64.6 | 1,340 | 35.2 | 872 | 22.9 | 468 | 12.3 |
| Hispanic | 2,684 | 1,736 | 64.7 | 937 | 34.9 | 641 | 23.9 | 296 | 11.0 |
| Other Non-Hispanic | 637 | 477 | 74.9 | 157 | 24.7 | 101 | 15.9 | 56 | 8.8 |

TABLE F. 6 (continued)

TABLE F. 7


|  | Total (000) | Food Secure |  | Food Insecure |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | All |  | Without Hunger |  | With Hunger |  |
|  |  | (000) | Percent | (000) | Percent | (000) | Percent | (000) | Percent |
| All Households | 74,658 | 69,928 | 93.7 | 4,704 | 6.3 | 3,031 | 4.1 | 1,673 | 2.2 |
| Household Composition |  |  |  |  |  |  |  |  |  |
| With Children < 6 | 12,557 | 11,412 | 90.9 | 1,136 | 9.1 | 886 | 7.1 | 250 | 2.0 |
| With Children < 18 | 27,630 | 25,162 | 91.1 | 2,459 | 8.9 | 1,778 | 6.4 | 681 | 2.5 |
| Married Couple Families | 22,106 | 20,622 | 93.3 | 1,476 | 6.7 | 1,120 | 5.1 | 356 | 1.6 |
| Female Head, No Spouse | 3,961 | 3,124 | 78.9 | 837 | 21.1 | 551 | 13.9 | 286 | 7.2 |
| Male Head, No Spouse | 1,227 | 1,126 | 91.7 | 101 | 8.3 | 84 | 6.9 | 17 | 1.4 |
| Other Households with Children | 335 | 291 | 86.7 | 45 | 13.3 | 23 | 6.9 | 22 | 6.4 |
| With No Children < 18 | 47,028 | 44,766 | 95.2 | 2,244 | 4.8 | 1,252 | 2.7 | 992 | 2.1 |
| More Than One Adult | 29,217 | 28,141 | 96.3 | 1,067 | 3.6 | 654 | 2.2 | 413 | 1.4 |
| Women Living Alone | 9,885 | 9,282 | 93.9 | 600 | 6.0 | 321 | 3.2 | 279 | 2.8 |
| Men Living Alone | 7,926 | 7,343 | 92.6 | 579 | 7.3 | 278 | 3.5 | 301 | 3.8 |
| Households with Elderly | 16,038 | 15,559 | 97.0 | 471 | 3.0 | 284 | 1.8 | 187 | 1.2 |
| Elderly Living Alone | 6,075 | 5,859 | 96.4 | 211 | 3.5 | 117 | 1.9 | 94 | 1.6 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |  |  |  |
| White Non-Hispanic | 61,416 | 58,293 | 94.9 | 3,111 | 5.1 | 2,006 | 3.3 | 1,105 | 1.8 |
| Black Non-Hispanic | 6,870 | 5,937 | 86.4 | 928 | 13.5 | 599 | 8.7 | 329 | 4.8 |
| Hispanic | 4,512 | 3,948 | 87.5 | 556 | 12.4 | 355 | 7.9 | 201 | 4.5 |
| Other Non-Hispanic | 1,860 | 1,750 | 94.1 | 110 | 5.9 | 71 | 3.8 | 39 | 2.1 |

TABLE F. 7 (continued)

TABLE F. 8


|  | $\begin{aligned} & \text { Total } \\ & (000) \\ & \hline \end{aligned}$ | Food Secure |  | All |  | Without Hunger |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | With Hunger |  |  |
|  |  | (000) | Percent |  |  | (000) | Percent | (000) | Percent | (000) | Percent |
| All Households | 75,388 | 70,438 | 93.4 | 4,928 | 6.5 | 3,169 | 4.2 | 1,759 | 2.3 |
| Household Composition |  |  |  |  |  |  |  |  |  |
| With Children < 6 | 12,467 | 11,151 | 89.4 | 1,303 | 10.5 | 980 | 7.9 | 323 | 2.6 |
| With Children < 18 | 27,836 | 25,206 | 90.6 | 2,612 | 9.4 | 1,883 | 6.8 | 729 | 2.6 |
| Married Couple Families | 22,082 | 20,460 | 92.7 | 1,611 | 7.3 | 1,188 | 5.4 | 423 | 1.9 |
| Female Head, No Spouse | 4,066 | 3,226 | 79.3 | 832 | 20.5 | 581 | 14.3 | 251 | 6.2 |
| Male Head, No Spouse | 1,297 | 1,171 | 90.3 | 125 | 9.7 | 84 | 6.5 | 41 | 3.2 |
| Other Households with Children | 392 | 349 | 89.0 | 43 | 11.0 | 29 | 7.5 | 14 | 3.5 |
| With No Children < 18 | 47,552 | 45,232 | 95.1 | 2,317 | 4.9 | 1,287 | 2.7 | 1,030 | 2.2 |
| More Than One Adult | 29,709 | 28,575 | 96.2 | 1,131 | 3.8 | 673 | 2.3 | 458 | 1.5 |
| Women Living Alone | 9,976 | 9,351 | 93.7 | 624 | 6.3 | 318 | 3.2 | 306 | 3.1 |
| Men Living Alone | 7,867 | 7,305 | 92.9 | 561 | 7.2 | 296 | 3.8 | 265 | 3.4 |
| Households with Elderly | 16,292 | 15,820 | 97.1 | 470 | 2.9 | 313 | 1.9 | 157 | 1.0 |
| Elderly Living Alone | 5,941 | 5,736 | 96.6 | 205 | 3.4 | 121 | 2.0 | 84 | 1.4 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |  |  |  |
| White Non-Hispanic | 60,975 | 57,828 | 94.8 | 3,139 | 5.2 | 1,988 | 3.3 | 1,151 | 1.9 |
| Black Non-Hispanic | 6,901 | 5,930 | 85.9 | 960 | 13.9 | 651 | 9.4 | 309 | 4.5 |
| Hispanic | 5,005 | 4,313 | 86.2 | 687 | 13.7 | 442 | 8.8 | 245 | 4.9 |
| Other Non-Hispanic | 2,508 | 2,366 | 94.4 | 141 | 5.6 | 88 | 3.5 | 53 | 2.1 |

TABLE F. 8 (continued)

TABLE F. 9
PREVALENCE OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER
BY SELECTED CHARACIERSITICS OF HOUSEHOLDS, 1997
FOR THOSE FAMILIES OVER 100 PERCENT OF POVERTY THRESHHOLD
USING THE COMMON DENOMINATOR SCREEN

|  | Total (000) | Food Secure |  | Food Insecure |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | All |  | Without Hunger |  | With Hunger |  |
|  |  | (000) | Percent | (000) | Percent | (000) | Percent | (000) | Percent |
| All Households | 76,383 | 72,268 | 94.6 | 4,084 | 5.4 | 2,790 | 3.7 | 1,294 | 1.7 |
| Household Composition |  |  |  |  |  |  |  |  |  |
| With Children < 6 | 12,153 | 11,237 | 92.5 | 913 | 7.5 | 729 | 6.0 | 184 | 1.5 |
| With Children < 18 | 27,964 | 25,902 | 92.6 | 2,056 | 7.4 | 1,557 | 5.6 | 499 | 1.8 |
| Married Couple Families | 21,892 | 20,730 | 94.7 | 1,157 | 5.3 | 897 | 4.1 | 260 | 1.2 |
| Female Head, No Spouse | 4,253 | 3,525 | 82.9 | 727 | 17.1 | 531 | 12.5 | 196 | 4.6 |
| Male Head, No Spouse | 1,499 | 1,362 | 90.9 | 136 | 9.1 | 111 | 7.4 | 25 | 1.7 |
| Other Households with Children | 320 | 285 | 88.9 | 35 | 11.1 | 17 | 5.5 | 18 | 5.6 |
| With No Children < 18 | 48,420 | 46,366 | 95.8 | 2,028 | 4.1 | 1,233 | 2.5 | 795 | 1.6 |
| More Than One Adult | 29,845 | 28,915 | 96.9 | 923 | 3.1 | 621 | 2.1 | 302 | 1.0 |
| Women Living Alone | 10,295 | 9,728 | 94.5 | 557 | 5.4 | 324 | 3.1 | 233 | 2.3 |
| Men Living Alone | 8,279 | 7,723 | 93.3 | 547 | 6.6 | 288 | 3.5 | 259 | 3.1 |
| Households with Elderly | 16,742 | 16,262 | 97.1 | 473 | 2.9 | 331 | 2.0 | 142 | 0.9 |
| Elderly Living Alone | 6,374 | 6,146 | 96.4 | 220 | 3.5 | 144 | 2.3 | 76 | 1.2 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |  |  |  |
| White Non-Hispanic | 61,383 | 58,802 | 95.8 | 2,570 | 4.2 | 1,728 | 2.8 | 842 | 1.4 |
| Black Non-Hispanic | 7,331 | 6,510 | 88.8 | 810 | 11.0 | 530 | 7.2 | 280 | 3.8 |
| Hispanic | 5,036 | 4,443 | 88.2 | 585 | 11.6 | 457 | 9.1 | 128 | 2.5 |
| Other Non-Hispanic | 2,634 | 2,513 | 95.4 | 120 | 4.6 | 76 | 2.9 | 44 | 1.7 |

TABLE F. 9 (continued)

|  | $\begin{aligned} & \text { Total } \\ & (000) \\ & \hline \end{aligned}$ | Food Secure |  | Food Insecure |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | All |  | Without Hunger |  | With Hunger |  |
|  |  | (000) | Percent | (000) | Percent | (000) | Percent | (000) | Percent |
| Household Income-to-Poverty Ratio |  |  |  |  |  |  |  |  |  |
| Under 0.50 | - | - | 0.0 | - | 0.0 | - | 0.0 | - | 0.0 |
| Under 1.00 | - | - | 0.0 | - | 0.0 | - | 0.0 | - | 0.0 |
| Under 1.30 | 6,001 | 4,925 | 82.1 | 1,068 | 17.8 | 711 | 11.8 | 357 | 6.0 |
| Under 1.85 | 17,403 | 14,801 | 85.0 | 2,580 | 14.8 | 1,795 | 10.3 | 785 | 4.5 |
| 1.85 and Over | 58,980 | 57,467 | 97.4 | 1,504 | 2.6 | 995 | 1.7 | 509 | 0.9 |
| Income Not Known | - | - | 0.0 | - | 0.0 | - | 0.0 | - | 0.0 |
| Area of Residence |  |  |  |  |  |  |  |  |  |
| Inside Metropolitan Area |  |  |  |  |  |  |  |  |  |
| In Central City | 18,241 | 17,011 | 93.3 | 1,216 | 6.7 | 762 | 4.2 | 454 | 2.5 |
| Not in Central City | 32,660 | 31,185 | 95.5 | 1,466 | 4.5 | 998 | 3.1 | 468 | 1.4 |
| Outside Metropolitan Area | 14,368 | 13,543 | 94.3 | 821 | 5.7 | 613 | 4.3 | 208 | 1.4 |
| Census Geographical Region |  |  |  |  |  |  |  |  |  |
| Northeast | 14,434 | 13,796 | 95.6 | 626 | 4.4 | 412 | 2.9 | 214 | 1.5 |
| Midwest | 18,784 | 17,855 | 95.1 | 929 | 5.0 | 594 | 3.2 | 335 | 1.8 |
| South | 26,406 | 24,857 | 94.1 | 1,539 | 5.8 | 1,079 | 4.1 | 460 | 1.7 |
| West | 16,759 | 15,760 | 94.0 | 990 | 5.9 | 705 | 4.2 | 285 | 1.7 |

TABLE F. 10

[^29]|  | $\begin{aligned} & \text { Total } \\ & (000) \\ & \hline \end{aligned}$ | Food Secure |  | Food Insecure |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | All |  | Without Hunger |  | With Hunger |  |
|  |  | (000) | Percent | (000) | Percent | (000) | Percent | (000) | Percent |
| All Households | 21,953 | 15,480 | 70.5 | 6,413 | 29.2 | 3,823 | 17.4 | 2,590 | 11.8 |
| Household Composition |  |  |  |  |  |  |  |  |  |
| With Children < 6 | 5,823 | 3,602 | 61.8 | 2,190 | 37.6 | 1,430 | 24.6 | 760 | 13.0 |
| With Children < 18 | 10,064 | 6,202 | 61.6 | 3,830 | 38.0 | 2,419 | 24.0 | 1,411 | 14.0 |
| Married Couple Families | 4,446 | 2,969 | 66.8 | 1,463 | 32.9 | 986 | 22.2 | 477 | 10.7 |
| Female Head, No Spouse | 4,799 | 2,701 | 56.3 | 2,090 | 43.5 | 1,272 | 26.5 | 818 | 17.0 |
| Male Head, No Spouse | 601 | 395 | 65.8 | 193 | 32.1 | 126 | 21.0 | 67 | 11.1 |
| Other Households with Children | 219 | 136 | 62.2 | 82 | 37.7 | 34 | 15.7 | 48 | 22.0 |
| With No Children < 18 | 11,889 | 9,278 | 78.0 | 2,584 | 21.7 | 1,404 | 11.8 | 1,180 | 9.9 |
| More Than One Adult | 4,771 | 3,778 | 79.2 | 978 | 20.5 | 590 | 12.4 | 388 | 8.1 |
| Women Living Alone | 4,801 | 3,837 | 79.9 | 959 | 20.0 | 553 | 11.5 | 406 | 8.5 |
| Men Living Alone | 2,317 | 1,662 | 71.8 | 647 | 27.9 | 261 | 11.3 | 386 | 16.6 |
| Households with Elderly | 6,495 | 5,619 | 86.5 | 872 | 13.4 | 538 | 8.3 | 334 | 5.1 |
| Elderly Living Alone | 3,878 | 3,413 | 88.0 | 465 | 12.0 | 257 | 6.6 | 208 | 5.4 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |  |  |  |
| White Non-Hispanic | 12,902 | 9,669 | 74.9 | 3,209 | 24.9 | 1,906 | 14.8 | 1,303 | 10.1 |
| Black Non-Hispanic | 5,076 | 3,289 | 64.8 | 1,776 | 35.0 | 1,022 | 20.1 | 754 | 14.9 |
| Hispanic | 3,295 | 2,021 | 61.3 | 1,250 | 37.9 | 792 | 24.0 | 458 | 13.9 |
| Other Non-Hispanic | 681 | 501 | 73.6 | 177 | 26.0 | 102 | 15.0 | 75 | 11.0 |

TABLE F. 10 (continued)

|  | $\begin{aligned} & \text { Total } \\ & (000) \\ & \hline \end{aligned}$ | Food Secure |  | Food Insecure |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | All |  | Without Hunger |  | With Hunger |  |
|  |  | (000) | Percent | (000) | Percent | (000) | Percent | (000) | Percent |
| Household Income-to-Poverty Ratio |  |  |  |  |  |  |  |  |  |
| Under 0.50 | 5,603 | 3,467 | 61.9 | 2,113 | 37.7 | 1,177 | 21.0 | 936 | 16.7 |
| Under 1.00 | 15,924 | 10,717 | 67.3 | 5,154 | 32.3 | 3,081 | 19.3 | 2073 | 13 |
| Under 1.30 | 21,953 | 15,480 | 70.5 | 6,413 | 29.2 | 3,823 | 17.4 | 2590 | 11.8 |
| Under 1.85 | 21,953 | 15,480 | 70.5 | 6,413 | 29.2 | 3,823 | 17.4 | 2590 | 11.8 |
| 1.85 and Over | - | - | 0.0 | - | 0.0 | - | 0.0 | 0 | 0 |
| Income Not Known | - | - | 0.0 | - | 0.0 | - | 0.0 | 0 | 0 |
| Area of Residence |  |  |  |  |  |  |  |  |  |
| Inside Metropolitan Area |  |  |  |  |  |  |  |  |  |
| In Central City | 6,967 | 4,644 | 66.7 | 2,320 | 33.3 | 1,386 | 19.9 | 934 | 13.4 |
| Not in Central City | 5,247 | 3,754 | 71.5 | 1,471 | 28.1 | 817 | 15.6 | 654 | 12.5 |
| Outside Metropolitan Area | 6,410 | 4,666 | 72.8 | 1,727 | 26.9 | 1,077 | 16.8 | 650 | 10.1 |
| Census Geographical Region |  |  |  |  |  |  |  |  |  |
| Northeast | 3,692 | 2,642 | 71.6 | 1,038 | 28.1 | 617 | 16.7 | 421 | 11.4 |
| Midwest | 4,801 | 3,489 | 72.7 | 1,311 | 27.4 | 790 | 16.5 | 521 | 10.9 |
| South | 8,888 | 6,291 | 70.8 | 2,585 | 29.1 | 1,549 | 17.4 | 1036 | 11.7 |
| West | 4,572 | 3,059 | 66.9 | 1,479 | 32.4 | 867 | 19.0 | 612 | 13.4 |

Table F. 11

[^30]|  | Total (000) | Food Secure |  | Food Insecure |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | All |  | Without Hunger |  | With Hunger |  |
|  |  | (000) | Percent | (000) | Percent | (000) | Percent | (000) | Percent |
| All Households | 20,282 | 13,868 | 68.4 | 6,353 | 31.4 | 3,726 | 18.4 | 2,627 | 13.0 |
| Household Composition |  |  |  |  |  |  |  |  |  |
| With Children < 6 | 5,345 | 3,253 | 60.9 | 2,073 | 38.7 | 1,360 | 25.4 | 713 | 13.3 |
| With Children < 18 | 9,434 | 5,675 | 60.2 | 3,720 | 39.5 | 2,355 | 25.0 | 1,365 | 14.5 |
| Married Couple Families | 4,275 | 2,769 | 64.8 | 1,482 | 34.7 | 975 | 22.8 | 507 | 11.9 |
| Female Head, No Spouse | 4,403 | 2,439 | 55.4 | 1,947 | 44.2 | 1,186 | 26.9 | 761 | 17.3 |
| Male Head, No Spouse | 568 | 336 | 59.2 | 231 | 40.8 | 157 | 27.7 | 74 | 13.1 |
| Other Households with Children | 189 | 130 | 68.8 | 59 | 31.2 | 37 | 19.5 | 22 | 11.7 |
| With No Children < 18 | 10,847 | 8,193 | 75.5 | 2,633 | 24.2 | 1,371 | 12.6 | 1,262 | 11.6 |
| More Than One Adult | 4,461 | 3,509 | 78.7 | 947 | 21.3 | 539 | 12.1 | 408 | 9.2 |
| Women Living Alone | 4,306 | 3,290 | 76.4 | 1,009 | 23.4 | 534 | 12.4 | 475 | 11.0 |
| Men Living Alone | 2,080 | 1,394 | 67.0 | 677 | 32.5 | 298 | 14.3 | 379 | 18.2 |
| Households with Elderly | 5,693 | 4,833 | 84.9 | 854 | 15.0 | 531 | 9.3 | 323 | 5.7 |
| Elderly Living Alone | 3,291 | 2,864 | 87.0 | 427 | 13.0 | 259 | 7.9 | 168 | 5.1 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |  |  |  |
| White Non-Hispanic | 11,354 | 8,320 | 73.3 | 3,014 | 26.5 | 1,703 | 15.0 | 1,311 | 11.5 |
| Black Non-Hispanic | 4,768 | 2,865 | 60.1 | 1,870 | 39.3 | 1,128 | 23.7 | 742 | 15.6 |
| Hispanic | 3,321 | 2,060 | 62.0 | 1,253 | 37.7 | 764 | 23.0 | 489 | 14.7 |
| Other Non-Hispanic | 840 | 622 | 74.1 | 215 | 25.7 | 131 | 15.7 | 84 | 10.0 |

TABLE F. 11 (continued)

TABLE F． 12

[^31]|  | Total （000） | Food Secure |  | Food Insecure |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | All |  | Without Hunger |  | With Hunger |  |
|  |  | （000） | Percent | （000） | Percent | （000） | Percent | （000） | Percent |
| All Households | 20，158 | 14，753 | 73.2 | 5，366 | 26.6 | 3，387 | 16.8 | 1，979 | 9.8 |
| Household Composition |  |  |  |  |  |  |  |  |  |
| With Children＜ 6 | 5，178 | 3，543 | 68.4 | 1，628 | 31.5 | 1，194 | 23.1 | 434 | 8.4 |
| With Children＜ 18 | 9，245 | 6，162 | 66.7 | 3，071 | 33.3 | 2，141 | 23.2 | 930 | 10.1 |
| Married Couple Families | 3，818 | 2，783 | 72.9 | 1，028 | 26.9 | 767 | 20.1 | 261 | 6.8 |
| Female Head，No Spouse | 4，606 | 2，801 | 60.8 | 1，802 | 39.1 | 1，184 | 25.7 | 618 | 13.4 |
| Male Head，No Spouse | 635 | 453 | 71.3 | 182 | 28.6 | 144 | 22.6 | 38 | 6.0 |
| Other Households with Children | 186 | 125 | 67.0 | 59 | 31.4 | 47 | 25.0 | 12 | 6.4 |
| With No Children＜ 18 | 10，913 | 8，591 | 78.7 | 2，295 | 21.0 | 1，246 | 11.4 | 1，049 | 9.6 |
| More Than One Adult | 4，516 | 3，605 | 79.8 | 901 | 20.0 | 524 | 11.6 | 377 | 8.4 |
| Women Living Alone | 4，211 | 3，402 | 80.8 | 797 | 18.9 | 452 | 10.7 | 345 | 8.2 |
| Men Living Alone | 2，186 | 1，583 | 72.4 | 596 | 27.2 | 269 | 12.3 | 327 | 14.9 |
| Households with Elderly | 5，393 | 4，638 | 86.0 | 749 | 13.9 | 528 | 9.8 | 221 | 4.1 |
| Elderly Living Alone | 3，377 | 2，979 | 88.2 | 395 | 11.7 | 259 | 7.7 | 136 | 4.0 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |  |  |  |
| White Non－Hispanic | 10，962 | 8，503 | 77.6 | 2，450 | 22.3 | 1，429 | 13.0 | 1，021 | 9.3 |
| Black Non－Hispanic | 4，815 | 3，224 | 66.9 | 1，580 | 32.8 | 1，026 | 21.3 | 554 | 11.5 |
| Hispanic | 3，515 | 2，367 | 67.3 | 1，132 | 32.2 | 799 | 22.7 | 333 | 9.5 |
| Other Non－Hispanic | 865 | 659 | 76.2 | 204 | 23.5 | 133 | 15.3 | 71 | 8.2 |

TABLE F. 12 (continued)

|  | $\begin{aligned} & \text { Total } \\ & (000) \\ & \hline \end{aligned}$ | Food Secure |  | Food Insecure |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | All |  | Without Hunger |  | With Hunger |  |
|  |  | (000) | Percent | (000) | Percent | (000) | Percent | (000) | Percent |
| Household Income-to-Poverty Ratio |  |  |  |  |  |  |  |  |  |
| Under 0.50 | 6,206 | 4,070 | 65.6 | 2,121 | 34.2 | 1,313 | 21.2 | 808 | 13.0 |
| Under 1.00 | 14,157 | 9,828 | 69.4 | 4,297 | 30.4 | 2,676 | 18.9 | 1621 | 11.5 |
| Under 1.30 | 20,158 | 14,753 | 73.2 | 5,366 | 26.6 | 3,387 | 16.8 | 1979 | 9.8 |
| Under 1.85 | 20,158 | 14,753 | 73.2 | 5,366 | 26.6 | 3,387 | 16.8 | 1979 | 9.8 |
| 1.85 and Over | - | - | 0.0 | - | 0.0 | - | 0.0 | - | 0.0 |
| Income Not Known | - | - | 0.0 | - | 0.0 | - | 0.0 | - | 0.0 |
| Area of Residence |  |  |  |  |  |  |  |  |  |
| Inside Metropolitan Area |  |  |  |  |  |  |  |  |  |
| In Central City | 6,644 | 4,611 | 69.4 | 2,018 | 30.4 | 1,206 | 18.2 | 812 | 12.2 |
| Not in Central City | 5,240 | 3,815 | 72.8 | 1,407 | 26.8 | 945 | 18.0 | 462 | 8.8 |
| Outside Metropolitan Area | 5,241 | 4,029 | 76.9 | 1,209 | 23.1 | 792 | 15.1 | 417 | 8.0 |
| Census Geographical Region |  |  |  |  |  |  |  |  |  |
| Northeast | 3,522 | 2,557 | 72.6 | 952 | 27.1 | 563 | 16.0 | 389 | 11.1 |
| Midwest | 4,021 | 2,953 | 73.4 | 1,059 | 26.3 | 647 | 16.1 | 412 | 10.2 |
| South | 8,088 | 6,038 | 74.7 | 2,044 | 25.2 | 1,345 | 16.6 | 699 | 8.6 |
| West | 4,527 | 3,206 | 70.8 | 1,310 | 29.0 | 832 | 18.4 | 478 | 10.6 |

TABLE F. 13


|  | $\begin{array}{r} \text { Total } \\ (000) \\ \hline \end{array}$ | Food Secure |  | Food Insecure |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | All |  | Without Hunger |  | With Hunger |  |
|  |  | (000) | Percent | (000) | Percent | (000) | Percent | (000) | Percent |
| All Households | 68,628 | 65,165 | 95.0 | 3,444 | 5.0 | 2,288 | 3.3 | 1,156 | 1.7 |
| Household Composition |  |  |  |  |  |  |  |  | 1.4 |
| With Children < 18 | 25,422 | 23,613 | 92.9 | 1,803 | 7.0 | 1,360 | 5.3 | 443 | 1.7 |
| Married Couple Families | 20,742 | 19,615 | 94.6 | 1,121 | 5.4 | 880 | 4.2 | 241 | 1.2 |
| Female Head, No Spouse | 3,311 | 2,734 | 82.6 | 577 | 17.5 | 403 | 12.2 | 174 | 5.3 |
| Male Head, No Spouse | 1,090 | 1,008 | 92.5 | 82 | 7.5 | 66 | 6.1 | 16 | 1.4 |
| Other Households with Children | 279 | 256 | 91.8 | 23 | 8.2 | 11 | 3.8 | 12 | 4.4 |
| With No Children < 18 | 43,206 | 41,552 | 96.2 | 1,641 | 3.8 | 928 | 2.1 | 713 | 1.7 |
| More Than One Adult | 27,630 | 26,799 | 97.0 | 826 | 3.0 | 496 | 1.8 | 330 | 1.2 |
| Women Living Alone | 8,358 | 7,943 | 95.0 | 414 | 4.9 | 226 | 2.7 | 188 | 2.2 |
| Men Living Alone | 7,217 | 6,810 | 94.4 | 402 | 5.6 | 206 | 2.9 | 196 | 2.7 |
| Households with Elderly | 13,638 | 13,370 | 98.0 | 262 | 2.0 | 171 | 1.3 | 91 | 0.7 |
| Elderly Living Alone | 4,628 | 4,543 | 98.2 | 83 | 1.8 | 52 | 1.1 | 31 | 0.7 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |  |  |  |
| White Non-Hispanic | 57,147 | 54,807 | 95.9 | 2,335 | 4.1 | 1,558 | 2.7 | 777 | 1.4 |
| Black Non-Hispanic | 5,972 | 5,300 | 88.8 | 666 | 11.2 | 433 | 7.3 | 233 | 3.9 |
| Hispanic | 3,809 | 3,426 | 90.0 | 374 | 9.8 | 254 | 6.7 | 120 | 3.1 |
| Other Non-Hispanic | 1,701 | 1,632 | 95.9 | 69 | 4.0 | 43 | 2.5 | 26 | 1.5 |

TABLE F. 13 (continued)

TABLE F. 14


|  | $\begin{aligned} & \text { Total } \\ & (000) \end{aligned}$ | Food Secure |  | All |  | Without Hunger |  | With Hunger |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  |  |  |  |  |
|  |  | (000) | Percent | (000) | Percent | (000) | Percent | (000) | Percent |
| All Households | 69,655 | 65,960 | 94.7 | 3,681 | 5.2 | 2,394 | 3.4 | 1,287 | 1.8 |
| Household Composition |  |  |  |  |  |  |  |  |  |
| With Children < 6 | 11,157 | 10,199 | 91.4 | 951 | 8.5 | 738 | 6.6 | 213 | 1.9 |
| With Children < 18 | 25,529 | 23,565 | 92.3 | 1,953 | 7.6 | 1,430 | 5.6 | 523 | 2.0 |
| Married Couple Families | 20,673 | 19,440 | 94.0 | 1,227 | 6.0 | 925 | 4.5 | 302 | 1.5 |
| Female Head, No Spouse | 3,361 | 2,745 | 81.7 | 611 | 18.2 | 429 | 12.8 | 182 | 5.4 |
| Male Head, No Spouse | 1,157 | 1,077 | 93.1 | 79 | 6.8 | 51 | 4.4 | 28 | 2.4 |
| Other Households with Children | 339 | 303 | 89.5 | 36 | 10.5 | 25 | 7.4 | 11 | 3.1 |
| With No Children < 18 | 44,126 | 42,395 | 96.1 | 1,727 | 3.9 | 964 | 2.2 | 763 | 1.7 |
| More Than One Adult | 28,269 | 27,343 | 96.7 | 924 | 3.3 | 528 | 1.9 | 396 | 1.4 |
| Women Living Alone | 8,613 | 8,208 | 95.3 | 404 | 4.7 | 217 | 2.5 | 187 | 2.2 |
| Men Living Alone | 7,244 | 6,843 | 94.5 | 401 | 5.5 | 220 | 3.0 | 181 | 2.5 |
| Households with Elderly | 14,208 | 13,918 | 98.0 | 287 | 2.0 | 198 | 1.4 | 89 | 0.6 |
| Elderly Living Alone | 4,716 | 4,611 | 97.8 | 105 | 2.2 | 68 | 1.4 | 37 | 0.8 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |  |  |  |
| White Non-Hispanic | 57,256 | 54,842 | 95.8 | 2,410 | 4.2 | 1,560 | 2.7 | 850 | 1.5 |
| Black Non-Hispanic | 5,905 | 5,219 | 88.4 | 682 | 11.5 | 461 | 7.8 | 221 | 3.7 |
| Hispanic | 4,201 | 3,710 | 88.3 | 485 | 11.5 | 307 | 7.3 | 178 | 4.2 |
| Other Non-Hispanic | 2,293 | 2,190 | 95.5 | 103 | 4.5 | 66 | 2.9 | 37 | 1.6 |

TABLE F. 14 (continued)

|  | $\begin{aligned} & \text { Total } \\ & (000) \\ & \hline \end{aligned}$ | Food Secure |  | Food Insecure |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | All |  | Without Hunger |  | With Hunger |  |
|  |  | (000) | Percent | (000) | Percent | (000) | Percent | (000) | Percent |
| Household Income-to-Poverty Ratio |  |  |  |  |  |  |  |  |  |
| Under 0.50 | - | - | 0.0 | - | 0.0 | - | 0.0 | 0 | 0 |
| Under 1.00 | - | - | 0.0 | - | 0.0 | - | 0.0 | 0 | 0 |
| Under 1.30 | - | - | 0.0 | - | 0.0 | - | 0.0 | 0 | 0 |
| Under 1.85 | 12,532 | 10,581 | 84.4 | 1,937 | 15.5 | 1,312 | 10.5 | 625 | 5 |
| 1.85 and Over | 57,123 | 55,379 | 96.9 | 1,743 | 3.1 | 1,082 | 1.9 | 661 | 1.2 |
| Income Not Known | - | - | 0.0 | - | 0.0 | - | 0.0 | 0 | 0 |
| Area of Residence |  |  |  |  |  |  |  |  |  |
| Inside Metropolitan Area |  |  |  |  |  |  |  |  |  |
| In Central City | 16,385 | 15,240 | 93.0 | 1,141 | 7.0 | 700 | 4.3 | 441 | 2.7 |
| Not in Central City | 30,467 | 29,109 | 95.5 | 1,350 | 4.5 | 869 | 2.9 | 481 | 1.6 |
| Outside Metropolitan Area | 12,604 | 11,927 | 94.6 | 675 | 5.3 | 471 | 3.7 | 204 | 1.6 |
| Census Geographical Region |  |  |  |  |  |  |  |  |  |
| Northeast | 13,712 | 13,113 | 95.6 | 598 | 4.4 | 406 | 3.0 | 192 | 1.4 |
| Midwest | 16,901 | 16,113 | 95.3 | 785 | 4.6 | 497 | 2.9 | 288 | 1.7 |
| South | 23,544 | 22,189 | 94.2 | 1,347 | 5.7 | 897 | 3.8 | 450 | 1.9 |
| West | 15,498 | 14,544 | 93.8 | 951 | 6.1 | 595 | 3.8 | 356 | 2.3 |

TABLE F. 15


|  | Total (000) | Food Secure |  | Food Insecure |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | All |  | Without Hunger |  | With Hunger |  |
|  |  | (000) | Percent | (000) | Percent | (000) | Percent | (000) | Percent |
| All Households | 70,382 | 67,343 | 95.7 | 3,016 | 4.3 | 2,079 | 3.0 | 937 | 1.3 |
| Household Composition |  |  |  |  |  |  |  |  |  |
| With Children < 6 | 11,033 | 10,369 | 94.0 | 664 | 6.0 | 531 | 4.8 | 133 | 1.2 |
| With Children < 18 | 25,755 | 24,175 | 93.9 | 1,576 | 6.1 | 1,206 | 4.7 | 370 | 1.4 |
| Married Couple Families | 20,706 | 19,797 | 95.6 | 906 | 4.4 | 702 | 3.4 | 204 | 1.0 |
| Female Head, No Spouse | 3,466 | 2,933 | 84.6 | 532 | 15.3 | 406 | 11.7 | 126 | 3.6 |
| Male Head, No Spouse | 1,326 | 1,212 | 91.4 | 115 | 8.7 | 93 | 7.0 | 22 | 1.7 |
| Other Households with Children | 256 | 234 | 91.3 | 22 | 8.7 | 5 | 2.1 | 17 | 6.6 |
| With No Children < 18 | 44,627 | 43,167 | 96.7 | 1,440 | 3.3 | 873 | 2.0 | 567 | 1.3 |
| More Than One Adult | 28,122 | 27,458 | 97.6 | 661 | 2.4 | 445 | 1.6 | 216 | 0.8 |
| Women Living Alone | 8,843 | 8,456 | 95.6 | 380 | 4.3 | 211 | 2.4 | 169 | 1.9 |
| Men Living Alone | 7,662 | 7,253 | 94.7 | 399 | 5.2 | 217 | 2.8 | 182 | 2.4 |
| Households with Elderly | 14,676 | 14,392 | 98.1 | 280 | 1.9 | 187 | 1.3 | 93 | 0.6 |
| Elderly Living Alone | 5,019 | 4,913 | 97.9 | 102 | 2.0 | 60 | 1.2 | 42 | 0.8 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |  |  |  |
| White Non-Hispanic | 57,446 | 55,453 | 96.5 | 1,983 | 3.5 | 1,360 | 2.4 | 623 | 1.1 |
| Black Non-Hispanic | 6,326 | 5,747 | 90.9 | 570 | 9.0 | 376 | 5.9 | 194 | 3.1 |
| Hispanic | 4,204 | 3,811 | 90.6 | 390 | 9.3 | 299 | 7.1 | 91 | 2.2 |
| Other Non-Hispanic | 2,405 | 2,331 | 96.9 | 74 | 3.1 | 45 | 1.9 | 29 | 1.2 |

TABLE F. 15 (continued)

TABLE F. 16

[^32]|  | $\begin{aligned} & \text { Total } \\ & (000) \\ & \hline \end{aligned}$ | Food Secure |  | Food Insecure |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | All |  | Without Hunger |  | With Hunger |  |
|  |  | (000) | Percent | (000) | Percent | (000) | Percent | (000) | Percent |
| All Households | 32,182 | 24,030 | 74.7 | 8,080 | 25.1 | 4,935 | 15.3 | 3,145 | 9.8 |
| Household Composition |  |  |  |  |  |  |  |  |  |
| With Children < 6 | 8,202 | 5,519 | 67.3 | 2,645 | 32.3 | 1,804 | 22.0 | 841 | 10.3 |
| With Children < 18 | 14,634 | 9,860 | 67.4 | 4,736 | 32.3 | 3,105 | 21.2 | 1,631 | 11.1 |
| Married Couple Families | 7,713 | 5,655 | 73.3 | 2,039 | 26.4 | 1,428 | 18.5 | 611 | 7.9 |
| Female Head, No Spouse | 5,799 | 3,424 | 59.0 | 2,368 | 40.9 | 1,477 | 25.5 | 891 | 15.4 |
| Male Head, No Spouse | 840 | 595 | 70.9 | 231 | 27.6 | 156 | 18.6 | 75 | 9.0 |
| Other Households with Children | 282 | 186 | 65.9 | 96 | 34.1 | 43 | 15.1 | 53 | 19.0 |
| With No Children < 18 | 17,548 | 14,170 | 80.8 | 3,344 | 19.0 | 1,830 | 10.4 | 1,514 | 8.6 |
| More Than One Adult | 7,682 | 6,331 | 82.4 | 1,334 | 17.4 | 814 | 10.6 | 520 | 6.8 |
| Women Living Alone | 6,470 | 5,295 | 81.8 | 1,171 | 18.1 | 671 | 10.4 | 500 | 7.7 |
| Men Living Alone | 3,395 | 2,544 | 74.9 | 840 | 24.8 | 345 | 10.2 | 495 | 14.6 |
| Households with Elderly | 8,801 | 7,776 | 88.3 | 1,016 | 11.6 | 632 | 7.2 | 384 | 4.4 |
| Elderly Living Alone | 5,033 | 4,518 | 89.8 | 511 | 10.2 | 286 | 5.7 | 225 | 4.5 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |  |  |  |
| White Non-Hispanic | 20,384 | 16,079 | 78.9 | 4,282 | 21.1 | 2,620 | 12.9 | 1,662 | 8.2 |
| Black Non-Hispanic | 6,449 | 4,326 | 67.1 | 2,109 | 32.7 | 1,241 | 19.2 | 868 | 13.5 |
| Hispanic | 4,384 | 2,871 | 65.5 | 1,481 | 33.8 | 952 | 21.7 | 529 | 12.1 |
| Other Non-Hispanic | 965 | 754 | 78.1 | 210 | 21.7 | 123 | 12.7 | 87 | 9.0 |

TABLE F. 16 (continued)

TABLE F. 17

[^33]|  | Total (000) | Food Secure |  | Food Insecure |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | All |  | Without Hunger |  | With Hunger |  |
|  |  | (000) | Percent | (000) | Percent | (000) | Percent | (000) | Percent |
| All Households | 32,814 | 24,449 | 74.5 | 8,290 | 25.3 | 5,038 | 15.4 | 3,252 | 9.9 |
| Household Composition |  |  |  |  |  |  |  |  |  |
| With Children < 6 | 7,795 | 5,126 | 65.8 | 2,642 | 33.9 | 1,810 | 23.2 | 832 | 10.7 |
| With Children < 18 | 14,321 | 9,458 | 66.0 | 4,813 | 33.6 | 3,162 | 22.1 | 1,651 | 11.5 |
| Married Couple Families | 7,422 | 5,270 | 71.0 | 2,123 | 28.6 | 1,461 | 19.7 | 662 | 8.9 |
| Female Head, No Spouse | 5,765 | 3,396 | 58.9 | 2,349 | 40.8 | 1,464 | 25.4 | 885 | 15.4 |
| Male Head, No Spouse | 846 | 578 | 68.4 | 267 | 31.6 | 185 | 21.9 | 82 | 9.7 |
| Other Households with Children | 288 | 214 | 74.2 | 74 | 25.8 | 52 | 18.1 | 22 | 7.7 |
| With No Children < 18 | 18,492 | 14,990 | 81.1 | 3,477 | 18.8 | 1,876 | 10.1 | 1,601 | 8.7 |
| More Than One Adult | 8,524 | 7,127 | 83.6 | 1,390 | 16.3 | 812 | 9.5 | 578 | 6.8 |
| Women Living Alone | 6,730 | 5,498 | 81.7 | 1,224 | 18.2 | 660 | 9.8 | 564 | 8.4 |
| Men Living Alone | 3,238 | 2,365 | 73.0 | 864 | 26.7 | 405 | 12.5 | 459 | 14.2 |
| Households with Elderly | 9,885 | 8,823 | 89.3 | 1,054 | 10.6 | 676 | 6.8 | 378 | 3.8 |
| Elderly Living Alone | 5,269 | 4,758 | 90.3 | 510 | 9.7 | 315 | 6.0 | 195 | 3.7 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |  |  |  |
| White Non-Hispanic | 20,635 | 16,367 | 79.3 | 4,244 | 20.6 | 2,531 | 12.3 | 1,713 | 8.3 |
| Black Non-Hispanic | 6,347 | 4,098 | 64.6 | 2,214 | 34.9 | 1,357 | 21.4 | 857 | 13.5 |
| Hispanic | 4,617 | 3,043 | 65.9 | 1,561 | 33.9 | 981 | 21.3 | 580 | 12.6 |
| Other Non-Hispanic | 1,215 | 942 | 77.5 | 271 | 22.3 | 169 | 13.9 | 102 | 8.4 |

TABLE F. 17 (continued)

TABLE F. 18

[^34]|  | $\begin{aligned} & \text { Total } \\ & (000) \end{aligned}$ | Food Secure |  | Food Insecure |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | All |  | Without Hunger |  | With Hunger |  |
|  |  | (000) | Percent | (000) | Percent | (000) | Percent | (000) | Percent |
| All Households | 31,560 | 24,629 | 78.0 | 6,878 | 21.8 | 4,471 | 14.2 | 2,407 | 7.6 |
| Household Composition |  |  |  |  |  |  |  |  |  |
| With Children < 6 | 7,460 | 5,406 | 72.5 | 2,046 | 27.4 | 1,515 | 20.3 | 531 | 7.1 |
| With Children < 18 | 13,975 | 10,004 | 71.6 | 3,956 | 28.3 | 2,811 | 20.1 | 1,145 | 8.2 |
| Married Couple Families | 6,953 | 5,399 | 77.7 | 1,544 | 22.2 | 1,145 | 16.5 | 399 | 5.7 |
| Female Head, No Spouse | 5,754 | 3,657 | 63.6 | 2,093 | 36.4 | 1,423 | 24.7 | 670 | 11.7 |
| Male Head, No Spouse | 1,000 | 758 | 75.8 | 242 | 24.2 | 191 | 19.1 | 51 | 5.1 |
| Other Households with Children | 269 | 190 | 70.7 | 76 | 28.1 | 52 | 19.3 | 24 | 8.8 |
| With No Children < 18 | 17,585 | 14,625 | 83.2 | 2,922 | 16.6 | 1,660 | 9.4 | 1,262 | 7.2 |
| More Than One Adult | 7,708 | 6,553 | 85.0 | 1,143 | 14.8 | 708 | 9.2 | 435 | 5.6 |
| Women Living Alone | 6,478 | 5,476 | 84.5 | 987 | 15.2 | 557 | 8.6 | 430 | 6.6 |
| Men Living Alone | 3,399 | 2,597 | 76.4 | 792 | 23.3 | 395 | 11.6 | 397 | 11.7 |
| Households with Elderly | 9,494 | 8,535 | 89.9 | 949 | 10.0 | 664 | 7.0 | 285 | 3.0 |
| Elderly Living Alone | 5,346 | 4,852 | 90.8 | 487 | 9.2 | 318 | 6.0 | 169 | 3.2 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |  |  |  |
| White Non-Hispanic | 19,194 | 15,762 | 82.1 | 3,417 | 17.8 | 2,135 | 11.1 | 1,282 | 6.7 |
| Black Non-Hispanic | 6,345 | 4,467 | 70.4 | 1,863 | 29.3 | 1,214 | 19.1 | 649 | 10.2 |
| Hispanic | 4,753 | 3,383 | 71.2 | 1,351 | 28.5 | 963 | 20.3 | 388 | 8.2 |
| Other Non-Hispanic | 1,268 | 1,017 | 80.2 | 248 | 19.5 | 159 | 12.5 | 89 | 7.0 |

TABLE F. 18 (continued)

|  | $\begin{aligned} & \text { Total } \\ & (000) \\ & \hline \end{aligned}$ | Food Secure |  | Food Insecure |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | All |  | Without Hunger |  | With Hunger |  |
|  |  | (000) | Percent | (000) | Percent | (000) | Percent | (000) | Percent |
| Household Income-to-Poverty Ratio |  |  |  |  |  |  |  |  |  |
| Under 0.50 | 6,206 | 4,070 | 65.6 | 2,121 | 34.2 | 1,313 | 21.2 | 808 | 13.0 |
| Under 1.00 | 14,157 | 9,828 | 69.4 | 4,297 | 30.4 | 2,676 | 18.9 | 1621 | 11.5 |
| Under 1.30 | 20,158 | 14,753 | 73.2 | 5,366 | 26.6 | 3,387 | 16.8 | 1979 | 9.8 |
| Under 1.85 | 31,560 | 24,629 | 78.0 | 6,878 | 21.8 | 4,471 | 14.2 | 2407 | 7.6 |
| 1.85 and Over | - | - | 0.0 | - | 0.0 | - | 0.0 | - | 0.0 |
| Income Not Known | - | - | 0.0 | - | 0.0 | - | 0.0 | - | 0.0 |
| Area of Residence |  |  |  |  |  |  |  |  |  |
| Inside Metropolitan Area |  |  |  |  |  |  |  |  |  |
| In Central City | 9,587 | 7,123 | 74.3 | 2,444 | 25.5 | 1,487 | 15.5 | 957 |  |
| Not in Central City | 9,189 | 7,255 | 79.0 | 1,911 | 20.8 | 1,288 | 14.0 | 623 | 6.8 |
| Outside Metropolitan Area | 8,119 | 6,529 | 80.4 | 1,584 | 19.5 | 1,099 | 13.5 | 485 | 6.0 |
| Census Geographical Region |  |  |  |  |  |  |  |  |  |
| Northeast | 5,449 | 4,282 | 78.6 | 1,155 | 21.2 | 706 | 13.0 | 449 | 8.2 |
| Midwest | 6,740 | 5,322 | 79.0 | 1,408 | 20.8 | 893 | 13.2 | 515 | 7.6 |
| South | 12,422 | 9,756 | 78.5 | 2,655 | 21.4 | 1,790 | 14.4 | 865 | 7.0 |
| West | 6,949 | 5,269 | 75.8 | 1,661 | 23.9 | 1,082 | 15.6 | 579 | 8.3 |

TABLE F. 19

IV

|  | $\begin{array}{r} \text { Total } \\ (000) \\ \hline \end{array}$ | Food Secure |  | All |  | Without Hunger |  | With Hunger |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | (000) | Percent | (000) | Percent | (000) | Percent | (000) | Percent |
| All Households | 58,400 | 56,615 | 96.9 | 1,777 | 3.0 | 1,176 | 2.0 | 601 | 1.0 |
| Household Composition |  |  |  |  |  |  |  |  |  |
| With Children < 6 | 9,002 | 8,645 | 96.0 | 357 | 4.0 | 280 | 3.1 | 77 | 0.9 |
| With Children < 18 | 20,852 | 19,956 | 95.7 | 896 | 4.3 | 674 | 3.2 | 222 | 1.1 |
| Married Couple Families | 17,475 | 16,930 | 96.9 | 545 | 3.1 | 438 | 2.5 | 107 | 0.6 |
| Female Head, No Spouse | 2,311 | 2,012 | 87.1 | 299 | 13.0 | 198 | 8.6 | 101 | 4.4 |
| Male Head, No Spouse | 851 | 808 | 94.9 | 43 | 5.0 | 36 | 4.2 | 7 | 0.8 |
| Other Households with Children | 216 | 206 | 95.7 | 9 | 4.3 | 2 | 1.1 | 7 | 3.2 |
| With No Children < 18 | 37,547 | 36,659 | 97.6 | 881 | 2.3 | 502 | 1.3 | 379 | 1.0 |
| More Than One Adult | 24,719 | 24,247 | 98.1 | 470 | 1.9 | 272 | 1.1 | 198 | 0.8 |
| Women Living Alone | 6,689 | 6,485 | 96.9 | 201 | 3.0 | 107 | 1.6 | 94 | 1.4 |
| Men Living Alone | 6,139 | 5,928 | 96.6 | 209 | 3.4 | 122 | 2.0 | 87 | 1.4 |
| Households with Elderly | 11,331 | 11,213 | 99.0 | 118 | 1.1 | 77 | 0.7 | 41 | 0.4 |
| Elderly Living Alone | 3,473 | 3,437 | 98.9 | 36 | 1.0 | 22 | 0.6 | 14 | 0.4 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |  |  |  |
| White Non-Hispanic | 49,664 | 48,397 | 97.4 | 1,262 | 2.5 | 844 | 1.7 | 418 | 0.8 |
| Black Non-Hispanic | 4,599 | 4,263 | 92.7 | 333 | 7.3 | 214 | 4.7 | 119 | 2.6 |
| Hispanic | 2,720 | 2,576 | 94.7 | 144 | 5.3 | 95 | 3.5 | 49 | 1.8 |
| Other Non-Hispanic | 1,417 | 1,379 | 97.4 | 37 | 2.6 | 23 | 1.6 | 14 | 1.0 |

TABLE F. 19 (continued)

Table F. 20


|  | $\begin{array}{r} \text { Total } \\ (000) \\ \hline \end{array}$ | Food Secure |  | Food Insecure |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | All |  | Without Hunger |  | With Hunger |  |
|  |  | (000) | Percent | (000) | Percent | (000) | Percent | (000) | Percent |
| All Households | 57,123 | 55,379 | 96.9 | 1,743 | 3.1 | 1,082 | 1.9 | 661 | 1.2 |
| Household Composition |  |  |  |  |  |  |  |  |  |
| With Children < 6 | 8,708 | 8,326 | 95.6 | 382 | 4.4 | 288 | 3.3 | 94 | 1.1 |
| With Children < 18 | 20,642 | 19,782 | 95.8 | 860 | 4.1 | 623 | 3.0 | 237 | 1.1 |
| Married Couple Families | 17,525 | 16,939 | 96.7 | 587 | 3.3 | 439 | 2.5 | 148 | 0.8 |
| Female Head, No Spouse | 1,998 | 1,788 | 89.5 | 210 | 10.5 | 152 | 7.6 | 58 | 2.9 |
| Male Head, No Spouse | 878 | 835 | 95.1 | 43 | 4.9 | 23 | 2.6 | 20 | 2.3 |
| Other Households with Children | 240 | 220 | 91.6 | 21 | 8.4 | 10 | 4.0 | 11 | 4.4 |
| With No Children < 18 | 36,481 | 35,597 | 97.6 | 884 | 2.5 | 459 | 1.3 | 425 | 1.2 |
| More Than One Adult | 24,206 | 23,725 | 98.0 | 481 | 2.0 | 255 | 1.1 | 226 | 0.9 |
| Women Living Alone | 6,189 | 6,000 | 96.9 | 190 | 3.1 | 92 | 1.5 | 98 | 1.6 |
| Men Living Alone | 6,086 | 5,873 | 96.5 | 212 | 3.5 | 112 | 1.8 | 100 | 1.7 |
| Households with Elderly | 10,016 | 9,929 | 99.1 | 87 | 0.8 | 52 | 0.5 | 35 | 0.3 |
| Elderly Living Alone | 2,738 | 2,716 | 99.2 | 22 | 0.8 | 12 | 0.4 | 10 | 0.4 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |  |  |  |
| White Non-Hispanic | 47,975 | 46,795 | 97.5 | 1,180 | 2.4 | 732 | 1.5 | 448 | 0.9 |
| Black Non-Hispanic | 4,326 | 3,986 | 92.2 | 339 | 7.9 | 232 | 5.4 | 107 | 2.5 |
| Hispanic | 2,905 | 2,727 | 93.9 | 177 | 6.1 | 90 | 3.1 | 87 | 3.0 |
| Other Non-Hispanic | 1,917 | 1,871 | 97.6 | 47 | 2.4 | 28 | 1.4 | 19 | 1.0 |

TABLE F. 20 (continued)

TABLE F. 21


|  | Total (000) | Food Secure |  | Food Insecure |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | All |  | Without Hunger |  | With Hunger |  |
|  |  | (000) | Percent | (000) | Percent | (000) | Percent | (000) | Percent |
| All Households | 58,980 | 57,467 | 97.4 | 1,504 | 2.6 | 995 | 1.7 | 509 | 0.9 |
| Household Composition |  |  |  |  |  |  |  |  |  |
| With Children < 6 | 8,750 | 8,505 | 97.2 | 245 | 2.8 | 210 | 2.4 | 35 | 0.4 |
| With Children < 18 | 21,025 | 20,334 | 96.7 | 691 | 3.3 | 537 | 2.6 | 154 | 0.7 |
| Married Couple Families | 17,571 | 17,181 | 97.8 | 390 | 2.2 | 324 | 1.8 | 66 | 0.4 |
| Female Head, No Spouse | 2,318 | 2,077 | 89.6 | 241 | 10.4 | 167 | 7.2 | 74 | 3.2 |
| Male Head, No Spouse | 962 | 907 | 94.3 | 54 | 5.7 | 45 | 4.7 | 9 | 1.0 |
| Other Households with Children | 173 | 168 | 97.1 | 5 | 2.9 | - | 0.0 | 5 | 2.9 |
| With No Children < 18 | 37,955 | 37,133 | 97.8 | 813 | 2.1 | 459 | 1.2 | 354 | 0.9 |
| More Than One Adult | 24,930 | 24,510 | 98.3 | 419 | 1.6 | 261 | 1.0 | 158 | 0.6 |
| Women Living Alone | 6,576 | 6,383 | 97.1 | 190 | 2.9 | 106 | 1.6 | 84 | 1.3 |
| Men Living Alone | 6,449 | 6,240 | 96.8 | 203 | 3.1 | 91 | 1.4 | 112 | 1.7 |
| Households with Elderly | 10,575 | 10,495 | 99.2 | 80 | 0.8 | 51 | 0.5 | 29 | 0.3 |
| Elderly Living Alone | 3,050 | 3,040 | 99.7 | 10 | 0.3 | 1 | 0.0 | 9 | 0.3 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |  |  |  |
| White Non-Hispanic | 49,215 | 48,194 | 97.9 | 1,016 | 2.0 | 654 | 1.3 | 362 | 0.7 |
| Black Non-Hispanic | 4,795 | 4,504 | 93.9 | 288 | 6.0 | 189 | 3.9 | 99 | 2.1 |
| Hispanic | 2,967 | 2,796 | 94.2 | 172 | 5.7 | 135 | 4.5 | 37 | 1.2 |
| Other Non-Hispanic | 2,002 | 1,974 | 98.6 | 29 | 1.4 | 18 | 0.9 | 11 | 0.5 |

TABLE F. 21 (continued)

TABLE F. 22
STANDARD ERRORS OF PREVALENCE ESTIMATES OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER BY SELECTED CHARACTERSITICS OF HOUSEHOLDS, 1995

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  |  | Standard |  | Standard |  | Standard |
|  | Percent | Error | Percent | Error | Percent | Error |
| All Households | 89.6 | 0.252 | 6.4 | 0.177 | 3.9 | 0.126 |
| Household Composition |  |  |  |  |  |  |
| With Children < 6 | 82.8 | 0.534 | 11.9 | 0.385 | 5.1 | 0.277 |
| With Children < 18 | 84.4 | 0.424 | 10.5 | 0.342 | 5.0 | 0.208 |
| Married Couple Families | 89.9 | 0.309 | 7.3 | 0.319 | 2.8 | 0.184 |
| Female Head, No Spouse | 67.7 | 1.039 | 20.4 | 0.923 | 11.8 | 0.526 |
| Male Head, No Spouse | 83.6 | 1.759 | 11.0 | 1.393 | 4.8 | 1.050 |
| Other Households with Children | 78.7 | 3.965 | 9.0 | 2.139 | 12.3 | 3.615 |
| With No Children < 18 | 92.7 | 0.221 | 3.9 | 0.129 | 3.3 | 0.143 |
| More Than One Adult | 94.7 | 0.243 | 3.1 | 0.171 | 2.1 | 0.140 |
| Women Living Alone | 90.4 | 0.389 | 5.4 | 0.347 | 4.2 | 0.153 |
| Men Living Alone | 89.1 | 0.536 | 4.7 | 0.394 | 6.0 | 0.548 |
| Households with Elderly | 94.9 | 0.202 | 3.1 | 0.142 | 1.9 | 0.154 |
| Elderly Living Alone | 94.2 | 0.429 | 3.3 | 0.263 | 2.5 | 0.317 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |
| White Non-Hispanic | 92.5 | 0.204 | 4.6 | 0.152 | 2.8 | 0.102 |
| Black Non-Hispanic | 78.3 | 0.832 | 12.8 | 0.690 | 8.8 | 0.534 |
| Hispanic | 77.0 | 1.276 | 14.7 | 0.996 | 7.9 | 0.686 |
| Other Non-Hispanic | 89.8 | 1.309 | 6.0 | 0.990 | 4.2 | 0.586 |

TABLE F. 22 (continued)

TABLE F. 23
STANDARD ERRORS OF PREVALENCE ESTIMATES OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER BY SELECTED CHARACTERSITICS OF HOUSEHOLDS, 1996

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  |  | Standard |  | Standard |  | Standard |
|  | Percent | Error | Percent | Error | Percent | Error |
| All Households | 89.5 | 0.161 | 6.4 | 0.110 | 4.1 | 0.099 |
| Household Composition |  |  |  |  |  |  |
| With Children < 6 | 82.3 | 0.410 | 12.2 | 0.315 | 5.3 | 0.299 |
| With Children < 18 | 84.3 | 0.334 | 10.4 | 0.177 | 5.1 | 0.246 |
| Married Couple Families | 89.5 | 0.279 | 7.3 | 0.270 | 3.1 | 0.130 |
| Female Head, No Spouse | 68.2 | 0.751 | 20.0 | 0.715 | 11.4 | 0.610 |
| Male Head, No Spouse | 82.9 | 2.377 | 11.3 | 1.410 | 5.8 | 1.763 |
| Other Households with Children | 82.3 | 2.289 | 12.1 | 1.873 | 5.6 | 1.346 |
| With No Children < 18 | 92.6 | 0.159 | 3.9 | 0.106 | 3.4 | 0.113 |
| More Than One Adult | 94.5 | 0.196 | 3.0 | 0.135 | 2.4 | 0.168 |
| Women Living Alone | 89.9 | 0.445 | 5.3 | 0.308 | 4.7 | 0.240 |
| Men Living Alone | 89.2 | 0.439 | 5.1 | 0.365 | 5.6 | 0.405 |
| Households with Elderly | 94.9 | 0.220 | 3.2 | 0.201 | 1.8 | 0.112 |
| Elderly Living Alone | 94.1 | 0.436 | 3.6 | 0.359 | 2.3 | 0.238 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |
| White Non-Hispanic | 92.6 | 0.147 | 4.4 | 0.112 | 2.9 | 0.065 |
| Black Non-Hispanic | 77.2 | 0.648 | 13.9 | 0.523 | 8.5 | 0.391 |
| Hispanic | 77.6 | 0.799 | 13.5 | 0.530 | 8.6 | 0.702 |
| Other Non-Hispanic | 89.8 | 0.687 | 6.3 | 0.493 | 3.6 | 0.418 |

Table F. 23 (continued)

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  |  | Standard |  | Standard |  | Standard |
|  | Percent | Error | Percent | Error | Percent | Error |
| Household Income-to-Poverty Ratio |  |  |  |  |  |  |
| Under 0.50 | 61.0 | 0.882 | 20.8 | 0.752 | 17.7 | 0.748 |
| Under 1.00 | 64.5 | 0.783 | 20.3 | 0.555 | 14.8 | 0.508 |
| Under 1.30 | 68.4 | 0.567 | 18.4 | 0.415 | 13.0 | 0.377 |
| Under 1.85 | 74.5 | 0.300 | 15.4 | 0.266 | 9.9 | 0.257 |
| 1.85 and Over | 96.9 | 0.116 | 1.9 | 0.092 | 1.2 | 0.055 |
| Income Not Known | 94.9 | 0.357 | 2.9 | 0.312 | 1.9 | 0.181 |
| Area of Residence |  |  |  |  |  |  |
| Inside Metropolitan Area |  |  |  |  |  |  |
| In Central City | 86.1 | 0.312 | 8.1 | 0.318 | 5.6 | 0.294 |
| Not in Central City | 91.8 | 0.233 | 4.8 | 0.173 | 3.2 | 0.134 |
| Outside Metropolitan Area | 88.7 | 0.320 | 7.3 | 0.311 | 4.0 | 0.224 |
| Census Geographical Region |  |  |  |  |  |  |
| Northeast | 92.0 | 0.266 | 4.8 | 0.220 | 3.1 | 0.125 |
| Midwest | 91.3 | 0.297 | 5.3 | 0.219 | 3.4 | 0.213 |
| South | 87.9 | 0.325 | 7.4 | 0.194 | 4.6 | 0.216 |
| West | 87.7 | 0.300 | 7.3 | 0.238 | 4.9 | 0.116 |

Source: Tabulations of Current Population Survey, Food Security Supplement data.
TABLE F. 24
PREVALENCE OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER BY SELECTED CHARACTERSITICS OF HOUSEHOLDS, 1997

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  |  | Standard |  | Standard |  | Standard |
|  | Percent | Error | Percent | Error | Percent | Error |
| All Households | 91.2 | 0.168 | 5.7 | 0.115 | 3.1 | 0.107 |
| Household Composition |  |  |  |  |  |  |
| With Children < 6 | 86.3 | 0.412 | 10.3 | 0.385 | 3.3 | 0.182 |
| With Children < 18 | 87.0 | 0.303 | 9.3 | 0.231 | 3.6 | 0.196 |
| Married Couple Families | 92.2 | 0.349 | 5.9 | 0.302 | 1.9 | 0.140 |
| Female Head, No Spouse | 71.9 | 0.800 | 19.0 | 0.928 | 8.9 | 0.518 |
| Male Head, No Spouse | 85.2 | 1.426 | 11.8 | 1.199 | 3.0 | 0.791 |
| Other Households with Children | 83.2 | 2.865 | 10.1 | 2.408 | 5.6 | 1.415 |
| With No Children < 18 | 93.6 | 0.153 | 3.6 | 0.116 | 2.7 | 0.112 |
| More Than One Adult | 95.4 | 0.189 | 2.9 | 0.147 | 1.7 | 0.155 |
| Women Living Alone | 91.8 | 0.350 | 4.5 | 0.341 | 3.6 | 0.246 |
| Men Living Alone | 90.2 | 0.665 | 4.7 | 0.344 | 4.9 | 0.500 |
| Households with Elderly | 95.3 | 0.210 | 3.2 | 0.203 | 1.4 | 0.116 |
| Elderly Living Alone | 94.9 | 0.320 | 3.2 | 0.285 | 1.8 | 0.172 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |
| White Non-Hispanic | 93.9 | 0.139 | 3.8 | 0.129 | 2.3 | 0.082 |
| Black Non-Hispanic | 81.1 | 0.785 | 12.2 | 0.570 | 6.5 | 0.438 |
| Hispanic | 80.4 | 0.799 | 14.0 | 0.621 | 5.4 | 0.484 |
| Other Non-Hispanic | 91.9 | 0.634 | 5.2 | 0.555 | 2.8 | 0.430 |

TABLE F. 24 (continued)

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  | Percent | Standard <br> Error | Percent | Standard <br> Error | Percent | Standard <br> Error |
| Household Income-to-Poverty Ratio |  |  |  |  |  |  |
| Under 0.50 | 65.6 | 1.061 | 21.2 | 0.938 | 13.0 | 0.698 |
| Under 1.00 | 69.4 | 0.693 | 18.9 | 0.610 | 11.5 | 0.449 |
| Under 1.30 | 73.2 | 0.633 | 16.8 | 0.512 | 9.8 | 0.428 |
| Under 1.85 | 78.0 | 0.385 | 14.2 | 0.354 | 7.6 | 0.244 |
| 1.85 and Over | 97.4 | 0.126 | 1.7 | 0.082 | 0.9 | 0.087 |
| Income Not Known | 94.8 | 0.529 | 3.1 | 0.336 | 1.8 | 0.375 |
| Area of Residence |  |  |  |  |  |  |
| Inside Metropolitan Area |  |  |  |  |  |  |
| In Central City | 88.1 | 0.446 | 7.1 | 0.292 | 4.6 | 0.273 |
| Not in Central City | 93.2 | 0.238 | 4.5 | 0.184 | 2.3 | 0.174 |
| Outside Metropolitan Area | 90.6 | 0.434 | 6.5 | 0.372 | 2.8 | 0.155 |
| Census Geographical Region |  |  |  |  |  |  |
| Northeast | 92.0 | 0.290 | 5.0 | 0.289 | 2.9 | 0.158 |
| Midwest | 92.4 | 0.301 | 4.7 | 0.260 | 2.9 | 0.117 |
| South | 90.4 | 0.302 | 6.3 | 0.207 | 3.1 | 0.142 |
| West | 90.3 | 0.432 | 6.4 | 0.147 | 3.2 | 0.397 |

Source: Tabulations of Current Population Survey, Food Security Supplement data.
TABLE F. 25
STANDARD ERRORS OF PREVALENCE ESTIMATES OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER BY SELECTED CHARACTERSITICS OF HOUSEHOLDS, 1995 FOR THOSE FAMILIES UNDER 100 PERCENT OF POVERTY THRESHHOLD USING THE COMMON DENOMINATOR SCREEN

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  |  | Standard |  | Standard |  | Standard |
|  | Percent | Error | Percent | Error | Percent | Error |
| All Households | 67.3 | 0.799 | 19.3 | 0.611 | 13.0 | 0.417 |
| Household Composition |  |  |  |  |  |  |
| With Children < 6 | 59.2 | 1.250 | 25.8 | 0.922 | 14.4 | 0.875 |
| With Children < 18 | 59.2 | 1.043 | 25.5 | 0.956 | 14.9 | 0.736 |
| Married Couple Families | 63.7 | 1.424 | 24.2 | 1.368 | 11.7 | 1.102 |
| Female Head, No Spouse | 55.7 | 1.671 | 27.1 | 1.515 | 17.0 | 0.879 |
| Male Head, No Spouse | 59.8 | 3.068 | 23.3 | 3.664 | 14.2 | 2.877 |
| Other Households with Children | 62.6 | 9.016 | 13.5 | 6.507 | 23.9 | 7.308 |
| With No Children < 18 | 75.2 | 1.053 | 13.4 | 0.714 | 11.2 | 0.642 |
| More Than One Adult | 76.5 | 1.273 | 13.6 | 1.141 | 9.6 | 0.900 |
| Women Living Alone | 76.3 | 1.442 | 14.0 | 1.280 | 9.6 | 0.561 |
| Men Living Alone | 70.3 | 2.637 | 11.8 | 1.411 | 17.5 | 2.090 |
| Households with Elderly | 83.8 | 1.035 | 10.4 | 0.988 | 5.8 | 0.620 |
| Elderly Living Alone | 86.2 | 1.117 | 7.9 | 0.972 | 5.9 | 0.754 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |
| White Non-Hispanic | 71.6 | 0.767 | 16.9 | 0.601 | 11.3 | 0.558 |
| Black Non-Hispanic | 63.5 | 1.685 | 20.5 | 1.263 | 15.8 | 0.923 |
| Hispanic | 57.9 | 1.637 | 26.7 | 1.625 | 14.5 | 1.310 |
| Other Non-Hispanic | 73.3 | 2.559 | 14.3 | 1.659 | 12.0 | 1.929 |

TABLE F. 25 (continued)

TABLE F. 26
STANDARD ERRORS OF PREVALENCE ESTIMATES OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER BY SELECTED CHARACTERSITICS OF HOUSEHOLDS, 1996
FOR THOSE FAMILIES UNDER 100 PERCENT OF POVERTY THRESH

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  |  | Standard |  | Standard |  | Standard |
|  | Percent | Error | Percent | Error | Percent | Error |
| All Households | 64.5 | 0.783 | 20.3 | 0.555 | 14.8 | 0.508 |
| Household Composition |  |  |  |  |  |  |
| With Children < 6 | 57.0 | 1.694 | 27.7 | 1.309 | 15.0 | 0.931 |
| With Children < 18 | 56.6 | 1.047 | 26.7 | 0.675 | 16.3 | 0.779 |
| Married Couple Families | 61.0 | 1.249 | 24.8 | 1.117 | 13.5 | 1.085 |
| Female Head, No Spouse | 53.0 | 1.158 | 28.0 | 1.315 | 18.7 | 1.004 |
| Male Head, No Spouse | 56.6 | 4.563 | 29.1 | 3.616 | 14.3 | 3.930 |
| Other Households with Children | 62.2 | 7.103 | 23.8 | 5.762 | 14.0 | 4.306 |
| With No Children < 18 | 72.2 | 0.969 | 14.1 | 0.670 | 13.4 | 0.850 |
| More Than One Adult | 75.4 | 1.361 | 13.0 | 1.206 | 11.4 | 1.280 |
| Women Living Alone | 72.9 | 1.614 | 14.7 | 0.931 | 12.1 | 1.048 |
| Men Living Alone | 64.0 | 1.708 | 15.2 | 1.670 | 20.2 | 2.094 |
| Households with Elderly | 81.2 | 0.829 | 11.5 | 0.817 | 7.1 | 0.566 |
| Elderly Living Alone | 84.1 | 1.063 | 10.0 | 0.801 | 5.9 | 0.762 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |
| White Non-Hispanic | 69.9 | 0.897 | 16.7 | 0.733 | 13.2 | 0.516 |
| Black Non-Hispanic | 57.1 | 1.400 | 24.9 | 1.004 | 17.4 | 0.981 |
| Hispanic | 57.9 | 2.374 | 25.0 | 1.349 | 16.8 | 1.500 |
| Other Non-Hispanic | 71.4 | 2.536 | 17.4 | 1.638 | 10.9 | 1.731 |

TABLE F. 26 (continued)

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  |  | Standard |  | Standard |  | Standard |
|  | Percent | Error | Percent | Error | Percent | Error |
| Household Income-to-Poverty Ratio |  |  |  |  |  |  |
| Under 0.50 | 61.0 | 0.882 | 20.8 | 0.752 | 17.7 | 0.748 |
| Under 1.00 | 64.5 | 0.783 | 20.3 | 0.555 | 14.8 | 0.508 |
| Under 1.30 | 64.5 | 0.783 | 20.3 | 0.555 | 14.8 | 0.508 |
| Under 1.85 | 64.5 | 0.783 | 20.3 | 0.555 | 14.8 | 0.508 |
| 1.85 and Over | 0.0 | na | 0.0 | na | 0.0 | na |
| Income Not Known | 0.0 | na | 0.0 | na | 0.0 | na |
| Area of Residence |  |  |  |  |  |  |
| Inside Metropolitan Area |  |  |  |  |  |  |
| In Central City | 63.0 | 1.025 | 21.5 | 1.086 | 15.1 | 0.854 |
| Not in Central City | 62.9 | 1.714 | 19.8 | 1.183 | 16.6 | 1.040 |
| Outside Metropolitan Area | 67.1 | 1.183 | 20.2 | 1.224 | 12.7 | 0.809 |
| Census Geographical Region |  |  |  |  |  |  |
| Northeast | 71.1 | 1.359 | 15.5 | 1.368 | 13.0 | 1.207 |
| Midwest | 66.7 | 1.373 | 19.2 | 1.358 | 13.8 | 0.659 |
| South | 63.2 | 1.589 | 21.0 | 0.915 | 15.3 | 0.912 |
| West | 60.2 | 1.098 | 23.5 | 0.786 | 16.1 | 1.139 |

TABLE F. 27
STANDARD ERRORS OF PREVALENCE ESTIMATES OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER BY SELECTED CHARACTERSITICS OF HOUSEHOLDS, 1997 FOR THOSE FAMILIES UNDER 100 PERCENT OF POVERTY THRESHHOLD USING THE COMMON DENOMINATOR SCREEN

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  |  | Standard |  | Standard |  | Standard |
|  | Percent | Error | Percent | Error | Percent | Error |
| All Households | 69.4 | 0.693 | 18.9 | 0.610 | 11.5 | 0.449 |
| Household Composition |  |  |  |  |  |  |
| With Children < 6 | 65.9 | 1.135 | 24.5 | 0.996 | 9.4 | 0.721 |
| With Children < 18 | 63.0 | 0.823 | 25.4 | 0.956 | 11.4 | 0.727 |
| Married Couple Families | 70.3 | 1.483 | 21.7 | 1.278 | 7.8 | 0.654 |
| Female Head, No Spouse | 57.8 | 1.461 | 27.7 | 1.698 | 14.3 | 1.001 |
| Male Head, No Spouse | 65.3 | 3.515 | 27.1 | 3.546 | 7.6 | 1.983 |
| Other Households with Children | 60.4 | 6.951 | 28.3 | 6.232 | 8.9 | 5.179 |
| With No Children < 18 | 75.7 | 0.864 | 12.4 | 0.666 | 11.5 | 0.595 |
| More Than One Adult | 76.9 | 1.179 | 12.5 | 1.231 | 10.4 | 1.111 |
| Women Living Alone | 77.2 | 1.184 | 12.3 | 1.012 | 10.2 | 0.702 |
| Men Living Alone | 71.0 | 2.373 | 12.6 | 1.662 | 15.9 | 2.217 |
| Households with Elderly | 83.2 | 1.188 | 11.6 | 1.143 | 5.1 | 0.524 |
| Elderly Living Alone | 86.3 | 1.229 | 8.7 | 1.103 | 5.0 | 0.740 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |
| White Non-Hispanic | 73.3 | 1.048 | 15.1 | 0.973 | 11.4 | 0.502 |
| Black Non-Hispanic | 64.6 | 1.386 | 22.9 | 1.149 | 12.3 | 0.918 |
| Hispanic | 64.7 | 1.984 | 23.9 | 1.715 | 11.0 | 1.215 |
| Other Non-Hispanic | 74.9 | 2.347 | 15.9 | 2.016 | 8.8 | 1.943 |

TABLE F. 27 (continued)

TABLE F. 28
STANDARD ERRORS OF PREVALENCE ESTIMATES OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER BY SELECTED CHARACTERSITICS OF HOUSEHOLDS, 1995
FOR THOSE FAMILIES OVER 100 PERCENT OF POVERTY THRESH

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  |  | Standard |  | Standard |  | Standard |
|  | Percent | Error | Percent | Error | Percent | Error |
| All Households | 93.7 | 0.181 | 4.1 | 0.158 | 2.2 | 0.090 |
| Household Composition |  |  |  |  |  |  |
| With Children < 6 | 90.9 | 0.367 | 7.1 | 0.357 | 2.0 | 0.201 |
| With Children < 18 | 91.1 | 0.309 | 6.4 | 0.298 | 2.5 | 0.137 |
| Married Couple Families | 93.3 | 0.256 | 5.1 | 0.222 | 1.6 | 0.158 |
| Female Head, No Spouse | 78.9 | 1.392 | 13.9 | 1.240 | 7.2 | 0.672 |
| Male Head, No Spouse | 91.7 | 1.101 | 6.9 | 1.085 | 1.4 | 0.466 |
| Other Households with Children | 86.7 | 2.748 | 6.9 | 2.330 | 6.4 | 2.374 |
| With No Children < 18 | 95.2 | 0.158 | 2.7 | 0.139 | 2.1 | 0.125 |
| More Than One Adult | 96.3 | 0.202 | 2.2 | 0.166 | 1.4 | 0.135 |
| Women Living Alone | 93.9 | 0.267 | 3.2 | 0.258 | 2.8 | 0.190 |
| Men Living Alone | 92.6 | 0.607 | 3.5 | 0.384 | 3.8 | 0.492 |
| Households with Elderly | 97.0 | 0.218 | 1.8 | 0.200 | 1.2 | 0.130 |
| Elderly Living Alone | 96.4 | 0.479 | 1.9 | 0.342 | 1.6 | 0.277 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |
| White Non-Hispanic | 94.9 | 0.168 | 3.3 | 0.149 | 1.8 | 0.080 |
| Black Non-Hispanic | 86.4 | 0.839 | 8.7 | 0.702 | 4.8 | 0.552 |
| Hispanic | 87.5 | 0.597 | 7.9 | 0.557 | 4.5 | 0.414 |
| Other Non-Hispanic | 94.1 | 1.324 | 3.8 | 1.045 | 2.1 | 0.414 |

TABLE F. 28 (continued)

TABLE F. 29
STANDARD ERRORS OF PREVALENCE ESTIMATES OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER BY SELECTED CHARACTERSITICS OF HOUSEHOLDS, 1996
FOR THOSE FAMILIES OVER 100 PERCENT OF POVERTY THRESHHOLD USING THE COMMON DENOMINATOR SCREEN

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  |  | Standard |  | Standard |  | Standard |
|  | Percent | Error | Percent | Error | Percent | Error |
| All Households | 93.4 | 0.156 | 4.2 | 0.110 | 2.3 | 0.104 |
| Household Composition |  |  |  |  |  |  |
| With Children < 6 | 89.4 | 0.501 | 7.9 | 0.394 | 2.6 | 0.257 |
| With Children < 18 | 90.6 | 0.243 | 6.8 | 0.215 | 2.6 | 0.164 |
| Married Couple Families | 92.7 | 0.285 | 5.4 | 0.241 | 1.9 | 0.135 |
| Female Head, No Spouse | 79.3 | 0.944 | 14.3 | 1.072 | 6.2 | 0.701 |
| Male Head, No Spouse | 90.3 | 1.461 | 6.5 | 1.010 | 3.2 | 1.062 |
| Other Households with Children | 89.0 | 1.981 | 7.5 | 2.159 | 3.5 | 1.124 |
| With No Children < 18 | 95.1 | 0.154 | 2.7 | 0.111 | 2.2 | 0.109 |
| More Than One Adult | 96.2 | 0.157 | 2.3 | 0.120 | 1.5 | 0.090 |
| Women Living Alone | 93.7 | 0.414 | 3.2 | 0.263 | 3.1 | 0.329 |
| Men Living Alone | 92.9 | 0.488 | 3.8 | 0.363 | 3.4 | 0.335 |
| Households with Elderly | 97.1 | 0.209 | 1.9 | 0.165 | 1.0 | 0.115 |
| Elderly Living Alone | 96.6 | 0.397 | 2.0 | 0.314 | 1.4 | 0.287 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |
| White Non-Hispanic | 94.8 | 0.145 | 3.3 | 0.112 | 1.9 | 0.079 |
| Black Non-Hispanic | 85.9 | 0.758 | 9.4 | 0.575 | 4.5 | 0.509 |
| Hispanic | 86.2 | 0.512 | 8.8 | 0.565 | 4.9 | 0.596 |
| Other Non-Hispanic | 94.4 | 0.518 | 3.5 | 0.453 | 2.1 | 0.350 |

TABLE F. 29 (continued)

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  | Percent | Standard Error | Percent | Standard Error | Percent | Standard Error |
| Household Income-to-Poverty Ratio |  |  |  |  |  |  |
| Under 0.50 | 0.0 | na | 0.0 | na | 0.0 | na |
| Under 1.00 | 0.0 | na | 0.0 | na | 0.0 | na |
| Under 1.30 | 78.1 | 0.839 | 13.5 | 0.668 | 8.2 | 0.879 |
| Under 1.85 | 82.4 | 0.473 | 11.4 | 0.435 | 6.0 | 0.373 |
| 1.85 and Over | 96.9 | 0.116 | 1.9 | 0.092 | 1.2 | 0.055 |
| Income Not Known | 0.0 | na | 0.0 | na | 0.0 | na |
| Area of Residence |  |  |  |  |  |  |
| Inside Metropolitan Area |  |  |  |  |  |  |
| In Central City | 91.5 | 0.277 | 5.1 | 0.314 | 3.3 | 0.288 |
| Not in Central City | 94.6 | 0.228 | 3.5 | 0.187 | 1.9 | 0.111 |
| Outside Metropolitan Area | 93.2 | 0.288 | 4.6 | 0.263 | 2.1 | 0.168 |
| Census Geographical Region |  |  |  |  |  |  |
| Northeast | 94.8 | 0.42 | 3.4 | 0.267 | 1.8 | 0.249 |
| Midwest | 94.4 | 0.392 | 3.5 | 0.234 | 2.1 | 0.242 |
| South | 92.7 | 0.221 | 4.9 | 0.209 | 2.3 | 0.149 |
| West | 92.3 | 0.237 | 4.6 | 0.144 | 3.1 | 0.228 |

Source: Tabulations of Current Population Survey, Food Security Supplement data.
TABLE F. 30
STANDARD ERRORS OF PREVALENCE ESTIMATES OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER BY SELECTED CHARACTERSITICS OF HOUSEHOLDS, 1997
FOR THOSE FAMILIES OVER 100 PERCENT OF POVERTY THRESHHOLD USING THE COMMON DENOMINATOR SCREEN

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  |  | Standard |  | Standard |  | Standard |
|  | Percent | Error | Percent | Error | Percent | Error |
| All Households | 94.6 | 0.164 | 3.7 | 0.108 | 1.7 | 0.095 |
| Household Composition |  |  |  |  |  |  |
| With Children < 6 | 92.5 | 0.430 | 6.0 | 0.382 | 1.5 | 0.183 |
| With Children < 18 | 92.6 | 0.333 | 5.6 | 0.241 | 1.8 | 0.160 |
| Married Couple Families | 94.7 | 0.364 | 4.1 | 0.310 | 1.2 | 0.126 |
| Female Head, No Spouse | 82.9 | 1.041 | 12.5 | 0.912 | 4.6 | 0.528 |
| Male Head, No Spouse | 90.9 | 1.658 | 7.4 | 1.819 | 1.7 | 0.707 |
| Other Households with Children | 88.9 | 3.400 | 5.5 | 2.046 | 5.6 | 2.543 |
| With No Children < 18 | 95.8 | 0.146 | 2.5 | 0.145 | 1.6 | 0.104 |
| More Than One Adult | 96.9 | 0.135 | 2.1 | 0.134 | 1.0 | 0.076 |
| Women Living Alone | 94.5 | 0.350 | 3.1 | 0.361 | 2.3 | 0.289 |
| Men Living Alone | 93.3 | 0.567 | 3.5 | 0.341 | 3.1 | 0.348 |
| Households with Elderly | 97.1 | 0.184 | 2.0 | 0.175 | 0.9 | 0.107 |
| Elderly Living Alone | 96.4 | 0.339 | 2.3 | 0.325 | 1.2 | 0.230 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |
| White Non-Hispanic | 95.8 | 0.130 | 2.8 | 0.114 | 1.4 | 0.063 |
| Black Non-Hispanic | 88.8 | 0.868 | 7.2 | 0.630 | 3.8 | 0.446 |
| Hispanic | 88.2 | 1.245 | 9.1 | 0.940 | 2.5 | 0.497 |
| Other Non-Hispanic | 95.4 | 0.674 | 2.9 | 0.499 | 1.7 | 0.509 |

TABLE F. 30 (continued)

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  |  | Standard |  | Standard |  | Standard |
|  | Percent | Error | Percent | Error | Percent | Error |
| Household Income-to-Poverty Ratio |  |  |  |  |  |  |
| Under 0.50 | 0.0 | na | 0.0 | na | 0.0 | na |
| Under 1.00 | 0.0 | na | 0.0 | na | 0.0 | na |
| Under 1.30 | 82.1 | 0.917 | 11.8 | 0.650 | 6.0 | 0.648 |
| Under 1.85 | 85.0 | 0.451 | 10.3 | 0.401 | 4.5 | 0.230 |
| 1.85 and Over | 97.4 | 0.126 | 1.7 | 0.082 | 0.9 | 0.087 |
| Income Not Known | 0.0 | na | 0.0 | na | 0.0 | na |
| Area of Residence |  |  |  |  |  |  |
| Inside Metropolitan Area |  |  |  |  |  |  |
| In Central City | 93.3 | 0.558 | 4.2 | 0.344 | 2.5 | 0.310 |
| Not in Central City | 95.5 | 0.165 | 3.1 | 0.118 | 1.4 | 0.125 |
| Outside Metropolitan Area | 94.3 | 0.440 | 4.3 | 0.366 | 1.4 | 0.174 |
| Census Geographical Region |  |  |  |  |  |  |
| Northeast | 95.6 | 0.267 | 2.9 | 0.177 | 1.5 | 0.127 |
| Midwest | 95.1 | 0.247 | 3.2 | 0.258 | 1.8 | 0.082 |
| South | 94.1 | 0.395 | 4.1 | 0.207 | 1.7 | 0.285 |
| West | 94.0 | 0.189 | 4.2 | 0.153 | 1.7 | 0.213 |

Source: Tabulations of Current Population Survey, Food Security Supplement data.
TABLE F. 31
STANDARD ERRORS OF PREVALENCE ESTIMATES OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER BY SELECTED CHARACTERSITICS OF HOUSEHOLDS, 1995 FOR THOSE FAMILIES UNDER 130 PERCENT OF POVERTY THRESHHOLD USING THE COMMON DENOMINATOR SCREEN

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  |  | Standard |  | Standard |  | Standard |
|  | Percent | Error | Percent | Error | Percent | Error |
| All Households | 70.5 | 0.702 | 17.4 | 0.483 | 11.8 | 0.435 |
| Household Composition |  |  |  |  |  |  |
| With Children < 6 | 61.8 | 1.138 | 24.6 | 0.771 | 13.0 | 0.800 |
| With Children < 18 | 61.6 | 0.886 | 24.0 | 0.790 | 14.0 | 0.677 |
| Married Couple Families | 66.8 | 1.261 | 22.2 | 1.268 | 10.7 | 0.900 |
| Female Head, No Spouse | 56.3 | 1.395 | 26.5 | 1.315 | 17.0 | 0.852 |
| Male Head, No Spouse | 65.8 | 3.623 | 21.0 | 3.540 | 11.1 | 2.396 |
| Other Households with Children | 62.2 | 7.120 | 15.7 | 5.076 | 22.0 | 6.167 |
| With No Children < 18 | 78.0 | 0.931 | 11.8 | 0.499 | 9.9 | 0.565 |
| More Than One Adult | 79.2 | 1.011 | 12.4 | 0.804 | 8.1 | 0.674 |
| Women Living Alone | 79.9 | 1.284 | 11.5 | 1.080 | 8.5 | 0.431 |
| Men Living Alone | 71.8 | 2.129 | 11.3 | 1.175 | 16.6 | 2.095 |
| Households with Elderly | 86.5 | 0.813 | 8.3 | 0.509 | 5.1 | 0.686 |
| Elderly Living Alone | 88.0 | 1.156 | 6.6 | 0.606 | 5.4 | 0.855 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |
| White Non-Hispanic | 74.9 | 0.620 | 14.8 | 0.465 | 10.1 | 0.471 |
| Black Non-Hispanic | 64.8 | 1.530 | 20.1 | 1.095 | 14.9 | 0.973 |
| Hispanic | 61.3 | 2.274 | 24.0 | 1.552 | 13.9 | 1.454 |
| Other Non-Hispanic | 73.6 | 3.540 | 15.0 | 2.161 | 11.0 | 1.902 |

TABLE F. 31 (continued)

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  |  | Standard |  | Standard |  | Standard |
|  | Percent | Error | Percent | Error | Percent | Error |
| Household Income-to-Poverty Ratio |  |  |  |  |  |  |
| Under 0.50 | 61.9 | 1.155 | 21.0 | 1.190 | 16.7 | 0.722 |
| Under 1.00 | 67.3 | 0.799 | 19.3 | 0.611 | 13.0 | 0.417 |
| Under 1.30 | 70.5 | 0.702 | 17.4 | 0.483 | 11.8 | 0.435 |
| Under 1.85 | 70.5 | 0.702 | 17.4 | 0.483 | 11.8 | 0.435 |
| 1.85 and Over | 0.0 | na | 0.0 | na | 0.0 | na |
| Income Not Known | 0.0 | na | 0.0 | na | 0.0 | na |
| Area of Residence |  |  |  |  |  |  |
| Inside Metropolitan Area |  |  |  |  |  |  |
| In Central City | 66.7 | 1.386 | 19.9 | 0.876 | 13.4 | 1.124 |
| Not in Central City | 71.5 | 0.616 | 15.6 | 0.812 | 12.5 | 0.687 |
| Outside Metropolitan Area | 72.8 | 1.166 | 16.8 | 0.742 | 10.1 | 0.764 |
| Census Geographical Region |  |  |  |  |  |  |
| Northeast | 71.6 | 1.566 | 16.7 | 1.153 | 11.4 | 0.804 |
| Midwest | 72.7 | 0.703 | 16.5 | 0.867 | 10.9 | 0.845 |
| South | 70.8 | 1.290 | 17.4 | 0.894 | 11.7 | 0.737 |
| West | 66.9 | 1.711 | 19.0 | 0.843 | 13.4 | 1.021 |

Source: Tabulations of Current Population Survey, Food Security Supplement data.
TABLE F. 32
STANDARD ERRORS OF PREVALENCE ESTIMATES OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER BY SELECTED CHARACTERSITICS OF HOUSEHOLDS, 1996
FOR THOSE FAMILIES UNDER 130 PERCENT OF POVERTY THRESH

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  |  | Standard |  | Standard |  | Standard |
|  | Percent | Error | Percent | Error | Percent | Error |
| All Households | 68.4 | 0.567 | 18.4 | 0.415 | 13.0 | 0.377 |
| Household Composition |  |  |  |  |  |  |
| With Children < 6 | 60.9 | 1.450 | 25.4 | 1.169 | 13.3 | 0.755 |
| With Children < 18 | 60.2 | 0.878 | 25.0 | 0.551 | 14.5 | 0.754 |
| Married Couple Families | 64.8 | 0.999 | 22.8 | 0.906 | 11.9 | 0.921 |
| Female Head, No Spouse | 55.4 | 0.950 | 26.9 | 1.121 | 17.3 | 0.926 |
| Male Head, No Spouse | 59.2 | 4.384 | 27.7 | 2.965 | 13.1 | 3.535 |
| Other Households with Children | 68.8 | 5.173 | 19.5 | 4.211 | 11.7 | 3.150 |
| With No Children < 18 | 75.5 | 0.639 | 12.6 | 0.501 | 11.6 | 0.501 |
| More Than One Adult | 78.7 | 0.925 | 12.1 | 0.759 | 9.2 | 0.897 |
| Women Living Alone | 76.4 | 1.145 | 12.4 | 0.764 | 11.0 | 0.677 |
| Men Living Alone | 67.0 | 1.438 | 14.3 | 1.332 | 18.2 | 1.607 |
| Households with Elderly | 84.9 | 0.659 | 9.3 | 0.659 | 5.7 | 0.475 |
| Elderly Living Alone | 87.0 | 0.925 | 7.9 | 0.684 | 5.1 | 0.650 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |
| White Non-Hispanic | 73.3 | 0.680 | 15.0 | 0.567 | 11.5 | 0.376 |
| Black Non-Hispanic | 60.1 | 1.239 | 23.7 | 0.832 | 15.6 | 0.924 |
| Hispanic | 62.0 | 1.843 | 23.0 | 1.263 | 14.7 | 1.293 |
| Other Non-Hispanic | 74.1 | 1.972 | 15.7 | 1.363 | 10.0 | 1.377 |

TABLE F. 32 (continued)

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  |  | Standard |  | Standard |  | Standard |
|  | Percent | Error | Percent | Error | Percent | Error |
| Household Income-to-Poverty Ratio |  |  |  |  |  |  |
| Under 0.50 | 61.0 | 0.882 | 20.8 | 0.752 | 17.7 | 0.748 |
| Under 1.00 | 64.5 | 0.783 | 20.3 | 0.555 | 14.8 | 0.508 |
| Under 1.30 | 68.4 | 0.567 | 18.4 | 0.415 | 13.0 | 0.377 |
| Under 1.85 | 68.4 | 0.567 | 18.4 | 0.415 | 13.0 | 0.377 |
| 1.85 and Over | 0.0 | na | 0.0 | na | 0.0 | na |
| Income Not Known | 0.0 | na | 0.0 | na | 0.0 | na |
| Area of Residence |  |  |  |  |  |  |
| Inside Metropolitan Area |  |  |  |  |  |  |
| In Central City | 66.3 | 0.788 | 19.5 | 0.862 | 13.8 | 0.959 |
| Not in Central City | 67.6 | 1.351 | 18.0 | 1.022 | 13.8 | 0.704 |
| Outside Metropolitan Area | 70.9 | 0.910 | 18.1 | 0.959 | 11.0 | 0.623 |
| Census Geographical Region |  |  |  |  |  |  |
| Northeast | 74.4 | 0.953 | 13.9 | 0.915 | 11.4 | 0.389 |
| Midwest | 71.3 | 1.159 | 16.7 | 1.133 | 11.8 | 0.728 |
| South | 66.8 | 1.164 | 19.8 | 0.662 | 13.0 | 0.766 |
| West | 64.0 | 0.526 | 20.7 | 0.649 | 15.2 | 0.611 |

Source: Tabulations of Current Population Survey, Food Security Supplement data.
TABLE F. 33
STANDARD ERRORS OF PREVALENCE ESTIMATES OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER BY SELECTED CHARACTERSITICS OF HOUSEHOLDS, 1997 FOR THOSE FAMILIES UNDER 130 PERCENT OF POVERTY THRESHHOLD USING THE COMMON DENOMINATOR SCREEN

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  | Percent | Standard <br> Error | Percent | Standard <br> Error | Percent | Standard <br> Error |
| All Households | 73.2 | 0.633 | 16.8 | 0.512 | 9.8 | 0.428 |
| Household Composition |  |  |  |  |  |  |
| With Children < 6 | 68.4 | 0.965 | 23.1 | 0.878 | 8.4 | 0.578 |
| With Children < 18 | 66.7 | 0.752 | 23.2 | 0.780 | 10.1 | 0.669 |
| Married Couple Families | 72.9 | 1.282 | 20.1 | 1.078 | 6.8 | 0.632 |
| Female Head, No Spouse | 60.8 | 1.309 | 25.7 | 1.481 | 13.4 | 0.996 |
| Male Head, No Spouse | 71.3 | 2.742 | 22.6 | 2.376 | 6.0 | 1.504 |
| Other Households with Children | 67.0 | 6.129 | 25.0 | 5.697 | 6.4 | 3.462 |
| With No Children < 18 | 78.7 | 0.791 | 11.4 | 0.584 | 9.6 | 0.504 |
| More Than One Adult | 79.8 | 0.955 | 11.6 | 0.973 | 8.4 | 0.802 |
| Women Living Alone | 80.8 | 0.993 | 10.7 | 0.907 | 8.2 | 0.644 |
| Men Living Alone | 72.4 | 1.882 | 12.3 | 1.276 | 14.9 | 1.695 |
| Households with Elderly | 86.0 | 0.895 | 9.8 | 0.786 | 4.1 | 0.398 |
| Elderly Living Alone | 88.2 | 1.037 | 7.7 | 0.891 | 4.0 | 0.526 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |
| White Non-Hispanic | 77.6 | 0.772 | 13.0 | 0.720 | 9.3 | 0.430 |
| Black Non-Hispanic | 66.9 | 1.332 | 21.3 | 1.048 | 11.5 | 0.803 |
| Hispanic | 67.3 | 1.346 | 22.7 | 1.153 | 9.5 | 1.084 |
| Other Non-Hispanic | 76.2 | 1.683 | 15.3 | 1.535 | 8.2 | 1.617 |

TABLE F. 33 (continued)

TABLE F. 34
STANDARD ERRORS OF PREVALENCE ESTIMATES OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER BY SELECTED CHARACTERSITICS OF HOUSEHOLDS, 1995
FOR THOSE FAMILIES OVER 130 PERCENT OF POVERTY THRESH

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  |  | Standard |  | Standard |  | Standard |
|  | Percent | Error | Percent | Error | Percent | Error |
| All Households | 95.0 | 0.152 | 3.3 | 0.123 | 1.7 | 0.079 |
| Household Composition |  |  |  |  |  |  |
| With Children < 6 | 92.8 | 0.355 | 5.7 | 0.301 | 1.4 | 0.141 |
| With Children < 18 | 92.9 | 0.279 | 5.3 | 0.253 | 1.7 | 0.118 |
| Married Couple Families | 94.6 | 0.247 | 4.2 | 0.220 | 1.2 | 0.113 |
| Female Head, No Spouse | 82.6 | 1.255 | 12.2 | 1.077 | 5.3 | 0.665 |
| Male Head, No Spouse | 92.5 | 0.946 | 6.1 | 0.969 | 1.4 | 0.511 |
| Other Households with Children | 91.8 | 2.061 | 3.8 | 1.794 | 4.4 | 1.977 |
| With No Children < 18 | 96.2 | 0.133 | 2.1 | 0.109 | 1.7 | 0.100 |
| More Than One Adult | 97.0 | 0.152 | 1.8 | 0.137 | 1.2 | 0.128 |
| Women Living Alone | 95.0 | 0.374 | 2.7 | 0.226 | 2.2 | 0.246 |
| Men Living Alone | 94.4 | 0.519 | 2.9 | 0.296 | 2.7 | 0.391 |
| Households with Elderly | 98.0 | 0.197 | 1.3 | 0.160 | 0.7 | 0.107 |
| Elderly Living Alone | 98.2 | 0.402 | 1.1 | 0.291 | 0.7 | 0.220 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |
| White Non-Hispanic | 95.9 | 0.140 | 2.7 | 0.111 | 1.4 | 0.078 |
| Black Non-Hispanic | 88.8 | 0.766 | 7.3 | 0.660 | 3.9 | 0.538 |
| Hispanic | 90.0 | 0.648 | 6.7 | 0.416 | 3.1 | 0.295 |
| Other Non-Hispanic | 95.9 | 0.721 | 2.5 | 0.715 | 1.5 | 0.370 |

TABLE F. 34 (continued)

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  |  | Standard |  | Standard |  | Standard |
|  | Percent | Error | Percent | Error | Percent | Error |
| Household Income-to-Poverty Ratio |  |  |  |  |  |  |
| Under 0.50 | 0.0 | na | 0.0 | na | 0.0 | na |
| Under 1.00 | 0.0 | na | 0.0 | na | 0.0 | na |
| Under 1.30 | 0.0 | na | 0.0 | na | 0.0 | na |
| Under 1.85 | 83.6 | 0.676 | 10.9 | 0.259 | 5.4 | 0.356 |
| 1.85 and Over | 96.9 | 0.098 | 2.0 | 0.091 | 1.0 | 0.070 |
| Income Not Known | 0.0 | na | 0.0 | na | 0.0 | na |
| Area of Residence |  |  |  |  |  |  |
| Inside Metropolitan Area |  |  |  |  |  |  |
| In Central City | 93.4 | 0.340 | 4.2 | 0.256 | 2.4 | 0.275 |
| Not in Central City | 95.5 | 0.249 | 2.9 | 0.161 | 1.6 | 0.182 |
| Outside Metropolitan Area | 95.3 | 0.339 | 3.4 | 0.280 | 1.3 | 0.152 |
| Census Geographical Region |  |  |  |  |  |  |
| Northeast | 95.9 | 0.217 | 2.9 | 0.180 | 1.2 | 0.102 |
| Midwest | 95.1 | 0.358 | 3.3 | 0.307 | 1.6 | 0.125 |
| South | 94.9 | 0.200 | 3.5 | 0.240 | 1.6 | 0.118 |
| West | 94.0 | 0.430 | 3.6 | 0.184 | 2.4 | 0.260 |

Source: Tabulations of Current Population Survey, Food Security Supplement data.
TABLE F. 35
STANDARD ERRORS OF PREVALENCE ESTIMATES OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER BY SELECTED CHARACTERSITICS OF HOUSEHOLDS, 1996
FOR THOSE FAMILIES OVER 130 PERCENT OF POVERTY THRESHHOLD USING THE COMMON DENOMINATOR SCREEN

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  | Percent | Standard <br> Error | Percent | Standard Error | Percent | Standard <br> Error |
| All Households | 94.7 | 0.143 | 3.4 | 0.109 | 1.8 | 0.075 |
| Household Composition |  |  |  |  |  |  |
| With Children < 6 | 91.4 | 0.496 | 6.6 | 0.409 | 1.9 | 0.198 |
| With Children < 18 | 92.3 | 0.269 | 5.6 | 0.231 | 2.0 | 0.131 |
| Married Couple Families | 94.0 | 0.306 | 4.5 | 0.269 | 1.5 | 0.125 |
| Female Head, No Spouse | 81.7 | 1.003 | 12.8 | 1.083 | 5.4 | 0.731 |
| Male Head, No Spouse | 93.1 | 1.305 | 4.4 | 0.714 | 2.4 | 1.101 |
| Other Households with Children | 89.5 | 2.255 | 7.4 | 2.362 | 3.1 | 0.882 |
| With No Children < 18 | 96.1 | 0.144 | 2.2 | 0.119 | 1.7 | 0.081 |
| More Than One Adult | 96.7 | 0.147 | 1.9 | 0.107 | 1.4 | 0.095 |
| Women Living Alone | 95.3 | 0.305 | 2.5 | 0.218 | 2.2 | 0.208 |
| Men Living Alone | 94.5 | 0.535 | 3.0 | 0.342 | 2.5 | 0.363 |
| Households with Elderly | 98.0 | 0.187 | 1.4 | 0.147 | 0.6 | 0.088 |
| Elderly Living Alone | 97.8 | 0.335 | 1.4 | 0.245 | 0.8 | 0.198 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |
| White Non-Hispanic | 95.8 | 0.125 | 2.7 | 0.109 | 1.5 | 0.065 |
| Black Non-Hispanic | 88.4 | 0.818 | 7.8 | 0.607 | 3.7 | 0.431 |
| Hispanic | 88.3 | 0.438 | 7.3 | 0.457 | 4.2 | 0.427 |
| Other Non-Hispanic | 95.5 | 0.537 | 2.9 | 0.418 | 1.6 | 0.303 |

TABLE F. 35 (continued)

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  | Percent | Standard Error | Percent | Standard Error | Percent | Standard <br> Error |
| Household Income-to-Poverty Ratio |  |  |  |  |  |  |
| Under 0.50 | 0.0 | na | 0.0 | na | 0.0 | na |
| Under 1.00 | 0.0 | na | 0.0 | na | 0.0 | na |
| Under 1.30 | 0.0 | na | 0.0 | na | 0.0 | na |
| Under 1.85 | 84.4 | 0.573 | 10.5 | 0.544 | 5.0 | 0.330 |
| 1.85 and Over | 96.9 | 0.116 | 1.9 | 0.092 | 1.2 | 0.055 |
| Income Not Known | 0.0 | na | 0.0 | na | 0.0 | na |
| Area of Residence |  |  |  |  |  |  |
| Inside Metropolitan Area |  |  |  |  |  |  |
| In Central City | 93.0 | 0.304 | 4.3 | 0.308 | 2.7 | 0.212 |
| Not in Central City | 95.5 | 0.231 | 2.9 | 0.215 | 1.6 | 0.101 |
| Outside Metropolitan Area | 94.6 | 0.316 | 3.7 | 0.284 | 1.6 | 0.130 |
| Census Geographical Region |  |  |  |  |  |  |
| Northeast | 95.6 | 0.361 | 3.0 | 0.291 | 1.4 | 0.163 |
| Midwest | 95.3 | 0.352 | 2.9 | 0.200 | 1.7 | 0.200 |
| South | 94.2 | 0.201 | 3.8 | 0.193 | 1.9 | 0.107 |
| West | 93.8 | 0.257 | 3.8 | 0.195 | 2.3 | 0.135 |

Source: Tabulations of Current Population Survey, Food Security Supplement data.
TABLE F. 36
STANDARD ERRORS OF PREVALENCE ESTIMATES OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER BY SELECTED CHARACTERSITICS OF HOUSEHOLDS, 1997
FOR THOSE FAMILIES OVER 130 PERCENT OF POVERTY THRESHHOLD

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  | Standard |  |  | Standard |  | Standard |
|  | Percent | Error | Percent | Error | Percent | Error |
| All Households | 95.7 | 0.158 | 3.0 | 0.112 | 1.3 | 0.079 |
| Household Composition |  |  |  |  |  |  |
| With Children < 6 | 94.0 | 0.420 | 4.8 | 0.395 | 1.2 | 0.202 |
| With Children < 18 | 93.9 | 0.291 | 4.7 | 0.216 | 1.4 | 0.149 |
| Married Couple Families | 95.6 | 0.298 | 3.4 | 0.265 | 1.0 | 0.119 |
| Female Head, No Spouse | 84.6 | 1.067 | 11.7 | 0.912 | 3.6 | 0.647 |
| Male Head, No Spouse | 91.4 | 1.527 | 7.0 | 1.771 | 1.7 | 0.774 |
| Other Households with Children | 91.3 | 3.693 | 2.1 | 1.281 | 6.6 | 3.124 |
| With No Children < 18 | 96.7 | 0.162 | 2.0 | 0.137 | 1.3 | 0.077 |
| More Than One Adult | 97.6 | 0.151 | 1.6 | 0.134 | 0.8 | 0.064 |
| Women Living Alone | 95.6 | 0.414 | 2.4 | 0.372 | 1.9 | 0.260 |
| Men Living Alone | 94.7 | 0.541 | 2.8 | 0.374 | 2.4 | 0.299 |
| Households with Elderly | 98.1 | 0.194 | 1.3 | 0.169 | 0.6 | 0.114 |
| Elderly Living Alone | 97.9 | 0.334 | 1.2 | 0.280 | 0.8 | 0.249 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |
| White Non-Hispanic | 96.5 | 0.115 | 2.4 | 0.099 | 1.1 | 0.063 |
| Black Non-Hispanic | 90.9 | 0.848 | 5.9 | 0.617 | 3.1 | 0.476 |
| Hispanic | 90.6 | 1.039 | 7.1 | 0.732 | 2.2 | 0.492 |
| Other Non-Hispanic | 96.9 | 0.568 | 1.9 | 0.449 | 1.2 | 0.449 |

TABLE F. 36 (continued)

TABLE F. 37
STANDARD ERRORS OF PREVALENCE ESTIMATES OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER BY SELECTED CHARACTERSITICS OF HOUSEHOLDS, 1995 FOR THOSE FAMILIES UNDER 185 PERCENT OF POVERTY THRESHHOLD USING THE COMMON DENOMINATOR SCREEN

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  |  | Standard |  | Standard |  | Standard |
|  | Percent | Error | Percent | Error | Percent | Error |
| All Households | 74.7 | 0.528 | 15.3 | 0.383 | 9.8 | 0.322 |
| Household Composition |  |  |  |  |  |  |
| With Children < 6 | 67.3 | 0.776 | 22.0 | 0.561 | 10.3 | 0.560 |
| With Children < 18 | 67.4 | 0.591 | 21.2 | 0.535 | 11.1 | 0.485 |
| Married Couple Families | 73.3 | 0.646 | 18.5 | 0.676 | 7.9 | 0.606 |
| Female Head, No Spouse | 59.0 | 1.192 | 25.5 | 1.086 | 15.4 | 0.735 |
| Male Head, No Spouse | 70.9 | 2.806 | 18.6 | 2.699 | 9.0 | 1.857 |
| Other Households with Children | 65.9 | 6.727 | 15.1 | 4.472 | 19.0 | 5.644 |
| With No Children < 18 | 80.8 | 0.692 | 10.4 | 0.427 | 8.6 | 0.422 |
| More Than One Adult | 82.4 | 0.886 | 10.6 | 0.713 | 6.8 | 0.421 |
| Women Living Alone | 81.8 | 1.077 | 10.4 | 0.891 | 7.7 | 0.331 |
| Men Living Alone | 74.9 | 1.423 | 10.2 | 0.938 | 14.6 | 1.571 |
| Households with Elderly | 88.3 | 0.546 | 7.2 | 0.373 | 4.4 | 0.400 |
| Elderly Living Alone | 89.8 | 0.853 | 5.7 | 0.499 | 4.5 | 0.601 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |
| White Non-Hispanic | 78.9 | 0.545 | 12.9 | 0.408 | 8.2 | 0.355 |
| Black Non-Hispanic | 67.1 | 1.192 | 19.2 | 0.955 | 13.5 | 0.749 |
| Hispanic | 65.5 | 1.865 | 21.7 | 1.316 | 12.1 | 1.139 |
| Other Non-Hispanic | 78.1 | 2.211 | 12.7 | 1.569 | 9.0 | 1.201 |

TABLE F. 37 (continued)

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  |  | Standard |  | Standard |  | Standard |
|  | Percent | Error | Percent | Error | Percent | Error |
| Household Income-to-Poverty Ratio |  |  |  |  |  |  |
| Under 0.50 | 61.9 | 1.155 | 21.0 | 1.190 | 16.7 | 0.722 |
| Under 1.00 | 67.3 | 0.799 | 19.3 | 0.611 | 13.0 | 0.417 |
| Under 1.30 | 70.5 | 0.702 | 17.4 | 0.483 | 11.8 | 0.435 |
| Under 1.85 | 74.7 | 0.528 | 15.3 | 0.383 | 9.8 | 0.322 |
| 1.85 and Over | 0.0 | na | 0.0 | na | 0.0 | na |
| Income Not Known | 0.0 | na | 0.0 | na | 0.0 | na |
| Area of Residence |  |  |  |  |  |  |
| Inside Metropolitan Area |  |  |  |  |  |  |
| In Central City | 70.4 | 0.938 | 17.9 | 0.627 | 11.6 | 0.770 |
| Not in Central City | 75.5 | 0.623 | 13.8 | 0.622 | 10.4 | 0.555 |
| Outside Metropolitan Area | 77.3 | 0.971 | 14.5 | 0.692 | 7.9 | 0.562 |
| Census Geographical Region |  |  |  |  |  |  |
| Northeast | 76.2 | 1.282 | 14.5 | 0.917 | 9.1 | 0.579 |
| Midwest | 76.0 | 0.913 | 14.9 | 0.781 | 9.1 | 0.647 |
| South | 75.1 | 0.882 | 15.3 | 0.658 | 9.5 | 0.571 |
| West | 71.2 | 1.329 | 16.6 | 0.767 | 11.6 | 0.717 |

Source: Tabulations of Current Population Survey, Food Security Supplement data.
TABLE F. 38
STANDARD ERRORS OF PREVALENCE ESTIMATES OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER BY SELECTED CHARACTERSITICS OF HOUSEHOLDS, 1996
FOR THOSE FAMILIES UNDER 185 PERCENT OF POVERTY THRESH

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  |  | Standard |  | Standard |  | Standard |
|  | Percent |  | Percent | Error | Percent | Error |
| All Households | 74.5 | 0.300 | 15.4 | 0.266 | 9.9 | 0.257 |
| Household Composition |  |  |  |  |  |  |
| With Children < 6 | 65.8 | 0.891 | 23.2 | 0.790 | 10.7 | 0.576 |
| With Children < 18 | 66.0 | 0.502 | 22.1 | 0.390 | 11.5 | 0.498 |
| Married Couple Families | 71.0 | 0.635 | 19.7 | 0.726 | 8.9 | 0.403 |
| Female Head, No Spouse | 58.9 | 0.889 | 25.4 | 1.061 | 15.4 | 0.841 |
| Male Head, No Spouse | 68.4 | 3.917 | 21.9 | 2.453 | 9.7 | 2.621 |
| Other Households with Children | 74.2 | 3.847 | 18.1 | 3.073 | 7.7 | 2.306 |
| With No Children < 18 | 81.1 | 0.424 | 10.1 | 0.351 | 8.7 | 0.352 |
| More Than One Adult | 83.6 | 0.707 | 9.5 | 0.549 | 6.8 | 0.715 |
| Women Living Alone | 81.7 | 0.758 | 9.8 | 0.559 | 8.4 | 0.441 |
| Men Living Alone | 73.0 | 1.114 | 12.5 | 0.918 | 14.2 | 1.230 |
| Households with Elderly | 89.3 | 0.405 | 6.8 | 0.409 | 3.8 | 0.237 |
| Elderly Living Alone | 90.3 | 0.639 | 6.0 | 0.516 | 3.7 | 0.404 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |
| White Non-Hispanic | 79.3 | 0.356 | 12.3 | 0.315 | 8.3 | 0.253 |
| Black Non-Hispanic | 64.6 | 1.004 | 21.4 | 0.684 | 13.5 | 0.754 |
| Hispanic | 65.9 | 1.215 | 21.3 | 1.114 | 12.6 | 0.962 |
| Other Non-Hispanic | 77.5 | 1.820 | 13.9 | 1.310 | 8.4 | 1.005 |

TABLE F. 38 (continued)

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  |  | Standard |  | Standard |  | Standard |
|  | Percent | Error | Percent | Error | Percent | Error |
| Household Income-to-Poverty Ratio |  |  |  |  |  |  |
| Under 0.50 | 61.0 | 0.882 | 20.8 | 0.752 | 17.7 | 0.748 |
| Under 1.00 | 64.5 | 0.783 | 20.3 | 0.555 | 14.8 | 0.508 |
| Under 1.30 | 68.4 | 0.567 | 18.4 | 0.415 | 13.0 | 0.377 |
| Under 1.85 | 74.5 | 0.300 | 15.4 | 0.266 | 9.9 | 0.257 |
| 1.85 and Over | 0.0 | na | 0.0 | na | 0.0 | na |
| Income Not Known | 0.0 | na | 0.0 | na | 0.0 | na |
| Area of Residence |  |  |  |  |  |  |
| Inside Metropolitan Area |  |  |  |  |  |  |
| In Central City | 71.3 | 0.694 | 16.9 | 0.873 | 11.5 | 0.608 |
| Not in Central City | 74.6 | 0.822 | 14.9 | 0.605 | 10.1 | 0.477 |
| Outside Metropolitan Area | 77.0 | 0.606 | 14.8 | 0.690 | 8.1 | 0.439 |
| Census Geographical Region |  |  |  |  |  |  |
| Northeast | 79.5 | 0.767 | 12.3 | 0.625 | 8.0 | 0.365 |
| Midwest | 77.7 | 0.759 | 13.2 | 0.563 | 8.9 | 0.717 |
| South | 72.6 | 0.435 | 16.8 | 0.438 | 10.3 | 0.461 |
| West | 71.0 | 0.559 | 17.2 | 0.568 | 11.7 | 0.303 |

Source: Tabulations of Current Population Survey, Food Security Supplement data.
TABLE F. 39
STANDARD ERRORS OF PREVALENCE ESTIMATES OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER BY SELECTED CHARACTERSITICS OF HOUSEHOLDS, 1997 FOR THOSE FAMILIES UNDER 185 PERCENT OF POVERTY THRESHHOLD USING THE COMMON DENOMINATOR SCREEN

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  |  | Standard |  | Standard |  | Standard |
|  | Percent | Error | Percent | Error | Percent | Error |
| All Households | 78.0 | 0.385 | 14.2 | 0.354 | 7.6 | 0.244 |
| Household Composition |  |  |  |  |  |  |
| With Children < 6 | 72.5 | 0.764 | 20.3 | 0.672 | 7.1 | 0.444 |
| With Children < 18 | 71.6 | 0.603 | 20.1 | 0.566 | 8.2 | 0.432 |
| Married Couple Families | 77.7 | 0.811 | 16.5 | 0.719 | 5.7 | 0.378 |
| Female Head, No Spouse | 63.6 | 1.136 | 24.7 | 1.260 | 11.7 | 0.830 |
| Male Head, No Spouse | 75.8 | 2.286 | 19.1 | 1.841 | 5.1 | 1.224 |
| Other Households with Children | 70.7 | 4.533 | 19.3 | 4.110 | 8.8 | 2.365 |
| With No Children < 18 | 83.2 | 0.466 | 9.4 | 0.421 | 7.2 | 0.337 |
| More Than One Adult | 85.0 | 0.600 | 9.2 | 0.611 | 5.6 | 0.491 |
| Women Living Alone | 84.5 | 0.683 | 8.6 | 0.689 | 6.6 | 0.471 |
| Men Living Alone | 76.4 | 1.593 | 11.6 | 0.809 | 11.7 | 1.350 |
| Households with Elderly | 89.9 | 0.481 | 7.0 | 0.480 | 3.0 | 0.240 |
| Elderly Living Alone | 90.8 | 0.624 | 6.0 | 0.609 | 3.2 | 0.317 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |
| White Non-Hispanic | 82.1 | 0.483 | 11.1 | 0.458 | 6.7 | 0.193 |
| Black Non-Hispanic | 70.4 | 1.163 | 19.1 | 0.985 | 10.2 | 0.670 |
| Hispanic | 71.2 | 0.852 | 20.3 | 0.853 | 8.2 | 0.645 |
| Other Non-Hispanic | 80.2 | 1.480 | 12.5 | 1.174 | 7.0 | 1.159 |

TABLE F. 39 (continued)

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  |  | Standard |  | Standard |  | Standard |
|  | Percent | Error | Percent | Error | Percent | Error |
| Household Income-to-Poverty Ratio |  |  |  |  |  |  |
| Under 0.50 | 65.6 | 1.061 | 21.2 | 0.938 | 13.0 | 0.698 |
| Under 1.00 | 69.4 | 0.693 | 18.9 | 0.610 | 11.5 | 0.449 |
| Under 1.30 | 73.2 | 0.633 | 16.8 | 0.512 | 9.8 | 0.428 |
| Under 1.85 | 78.0 | 0.385 | 14.2 | 0.354 | 7.6 | 0.244 |
| 1.85 and Over | 0.0 | na | 0.0 | na | 0.0 | na |
| Income Not Known | 0.0 | na | 0.0 | na | 0.0 | na |
| Area of Residence |  |  |  |  |  |  |
| Inside Metropolitan Area |  |  |  |  |  |  |
| In Central City | 74.3 | 0.896 | 15.5 | 0.623 | 10.0 | 0.553 |
| Not in Central City | 79.0 | 0.618 | 14.0 | 0.634 | 6.8 | 0.459 |
| Outside Metropolitan Area | 80.4 | 0.925 | 13.5 | 0.788 | 6.0 | 0.395 |
| Census Geographical Region |  |  |  |  |  |  |
| Northeast | 78.6 | 0.976 | 13.0 | 1.051 | 8.2 | 0.520 |
| Midwest | 79.0 | 1.114 | 13.2 | 1.042 | 7.6 | 0.536 |
| South | 78.5 | 0.547 | 14.4 | 0.473 | 7.0 | 0.164 |
| West | 75.8 | 0.589 | 15.6 | 0.418 | 8.3 | 0.833 |

Source: Tabulations of Current Population Survey, Food Security Supplement data.
TABLE F. 40
STANDARD ERRORS OF PREVALENCE ESTIMATES OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER BY SELECTED CHARACTERSITICS OF HOUSEHOLDS, 1995
FOR THOSE FAMILIES OVER 185 PERCENT OF POVERTY THRESHH

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  |  | Standard |  | Standard |  | Standard |
|  | Percent | Error | Percent | Error | Percent | Error |
| All Households | 96.9 | 0.098 | 2.0 | 0.091 | 1.0 | 0.070 |
| Household Composition |  |  |  |  |  |  |
| With Children < 6 | 96.0 | 0.326 | 3.1 | 0.277 | 0.9 | 0.150 |
| With Children < 18 | 95.7 | 0.233 | 3.2 | 0.218 | 1.1 | 0.088 |
| Married Couple Families | 96.9 | 0.179 | 2.5 | 0.191 | 0.6 | 0.085 |
| Female Head, No Spouse | 87.1 | 1.232 | 8.6 | 1.031 | 4.4 | 0.553 |
| Male Head, No Spouse | 94.9 | 0.801 | 4.2 | 0.867 | 0.8 | 0.449 |
| Other Households with Children | 95.7 | 1.964 | 1.1 | 1.088 | 3.2 | 1.764 |
| With No Children < 18 | 97.6 | 0.104 | 1.3 | 0.074 | 1.0 | 0.090 |
| More Than One Adult | 98.1 | 0.130 | 1.1 | 0.087 | 0.8 | 0.117 |
| Women Living Alone | 96.9 | 0.399 | 1.6 | 0.211 | 1.4 | 0.264 |
| Men Living Alone | 96.6 | 0.370 | 2.0 | 0.292 | 1.4 | 0.271 |
| Households with Elderly | 99.0 | 0.125 | 0.7 | 0.119 | 0.4 | 0.076 |
| Elderly Living Alone | 98.9 | 0.242 | 0.6 | 0.198 | 0.4 | 0.146 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |
| White Non-Hispanic | 97.4 | 0.103 | 1.7 | 0.084 | 0.8 | 0.067 |
| Black Non-Hispanic | 92.7 | 0.676 | 4.7 | 0.628 | 2.6 | 0.436 |
| Hispanic | 94.7 | 0.634 | 3.5 | 0.605 | 1.8 | 0.281 |
| Other Non-Hispanic | 97.4 | 0.753 | 1.6 | 0.697 | 1.0 | 0.338 |

TABLE F. 40 (continued)

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  | Percent | Standard Error | Percent | Standard Error | Percent | Standard Error |
| Household Income-to-Poverty Ratio |  |  |  |  |  |  |
| Under 0.50 | 0.0 | na | 0.0 | na | 0.0 | na |
| Under 1.00 | 0.0 | na | 0.0 | na | 0.0 | na |
| Under 1.30 | 0.0 | na | 0.0 | na | 0.0 | na |
| Under 1.85 | 0.0 | na | 0.0 | na | 0.0 | na |
| 1.85 and Over | 96.9 | 0.098 | 2.0 | 0.091 | 1.0 | 0.070 |
| Income Not Known | 0.0 | na | 0.0 | na | 0.0 | na |
| Area of Residence |  |  |  |  |  |  |
| Inside Metropolitan Area |  |  |  |  |  |  |
| In Central City | 95.9 | 0.165 | 2.5 | 0.211 | 1.5 | 0.186 |
| Not in Central City | 97.1 | 0.168 | 1.9 | 0.131 | 0.9 | 0.120 |
| Outside Metropolitan Area | 97.4 | 0.266 | 1.7 | 0.178 | 0.8 | 0.154 |
| Census Geographical Region |  |  |  |  |  |  |
| Northeast | 97.3 | 0.189 | 1.8 | 0.182 | 0.8 | 0.121 |
| Midwest | 97.2 | 0.249 | 1.9 | 0.259 | 0.9 | 0.136 |
| South | 96.9 | 0.138 | 2.1 | 0.142 | 1.0 | 0.136 |
| West | 96.4 | 0.223 | 2.2 | 0.114 | 1.4 | 0.155 |

Source: Tabulations of Current Population Survey, Food Security Supplement data.
TABLE F. 41
STANDARD ERRORS OF PREVALENCE ESTIMATES OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER BY SELECTED CHARACTERSITICS OF HOUSEHOLDS, 1996
FOR THOSE FAMILIES OVER 185 PERCENT OF POVERTY THRESH

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  |  | Standard |  | Standard |  | Standard |
|  | Percent | Error | Percent | Error | Percent | Error |
| All Households | 96.9 | 0.116 | 1.9 | 0.092 | 1.2 | 0.055 |
| Household Composition |  |  |  |  |  |  |
| With Children < 6 | 95.6 | 0.452 | 3.3 | 0.383 | 1.1 | 0.167 |
| With Children < 18 | 95.8 | 0.254 | 3.0 | 0.196 | 1.1 | 0.109 |
| Married Couple Families | 96.7 | 0.279 | 2.5 | 0.206 | 0.8 | 0.114 |
| Female Head, No Spouse | 89.5 | 1.014 | 7.6 | 0.831 | 2.9 | 0.795 |
| Male Head, No Spouse | 95.1 | 1.451 | 2.6 | 0.811 | 2.3 | 1.269 |
| Other Households with Children | 91.6 | 1.660 | 4.0 | 2.084 | 4.4 | 1.227 |
| With No Children < 18 | 97.6 | 0.131 | 1.3 | 0.105 | 1.2 | 0.071 |
| More Than One Adult | 98.0 | 0.128 | 1.1 | 0.121 | 0.9 | 0.074 |
| Women Living Alone | 96.9 | 0.330 | 1.5 | 0.227 | 1.6 | 0.273 |
| Men Living Alone | 96.5 | 0.484 | 1.8 | 0.325 | 1.7 | 0.374 |
| Households with Elderly | 99.1 | 0.109 | 0.5 | 0.106 | 0.3 | 0.059 |
| Elderly Living Alone | 99.2 | 0.242 | 0.4 | 0.207 | 0.4 | 0.124 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |
| White Non-Hispanic | 97.5 | 0.117 | 1.5 | 0.089 | 0.9 | 0.060 |
| Black Non-Hispanic | 92.2 | 0.938 | 5.4 | 0.698 | 2.5 | 0.436 |
| Hispanic | 93.9 | 0.430 | 3.1 | 0.391 | 3.0 | 0.340 |
| Other Non-Hispanic | 97.6 | 0.489 | 1.4 | 0.430 | 1.0 | 0.244 |

TABLE F. 41 (continued)

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  | Percent | Standard Error | Percent | Standard Error | Percent | Standard Error |
| Household Income-to-Poverty Ratio |  |  |  |  |  |  |
| Under 0.50 | 0.0 | na | 0.0 | na | 0.0 | na |
| Under 1.00 | 0.0 | na | 0.0 | na | 0.0 | na |
| Under 1.30 | 0.0 | na | 0.0 | na | 0.0 | na |
| Under 1.85 | 0.0 | na | 0.0 | na | 0.0 | na |
| 1.85 and Over | 96.9 | 0.116 | 1.9 | 0.092 | 1.2 | 0.055 |
| Income Not Known | 0.0 | na | 0.0 | na | 0.0 | na |
| Area of Residence |  |  |  |  |  |  |
| Inside Metropolitan Area |  |  |  |  |  |  |
| In Central City | 96.0 | 0.237 | 2.4 | 0.238 | 1.6 | 0.234 |
| Not in Central City | 97.3 | 0.137 | 1.6 | 0.139 | 1.0 | 0.112 |
| Outside Metropolitan Area | 97.2 | 0.347 | 1.8 | 0.265 | 1.0 | 0.160 |
| Census Geographical Region |  |  |  |  |  |  |
| Northeast | 97.2 | 0.166 | 1.6 | 0.163 | 1.1 | 0.052 |
| Midwest | 97.1 | 0.260 | 1.8 | 0.233 | 1.0 | 0.128 |
| South | 97.0 | 0.250 | 1.9 | 0.183 | 1.1 | 0.091 |
| West | 96.4 | 0.178 | 2.1 | 0.097 | 1.5 | 0.146 |

Source: Tabulations of Current Population Survey, Food Security Supplement data.
TABLE F. 42
STANDARD ERRORS OF PREVALENCE ESTIMATES OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER BY SELECTED CHARACTERSITICS OF HOUSEHOLDS, 1997
FOR THOSE FAMILIES OVER 185 PERCENT OF POVERTY THRESHHOLD

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  | Percent | Standard <br> Error | Percent | Standard Error | Percent | Standard Error |
| All Households | 97.4 | 0.126 | 1.7 | 0.082 | 0.9 | 0.087 |
| Household Composition |  |  |  |  |  |  |
| With Children < 6 | 97.2 | 0.289 | 2.4 | 0.241 | 0.4 | 0.115 |
| With Children < 18 | 96.7 | 0.215 | 2.6 | 0.142 | 0.7 | 0.155 |
| Married Couple Families | 97.8 | 0.199 | 1.8 | 0.161 | 0.4 | 0.090 |
| Female Head, No Spouse | 89.6 | 0.967 | 7.2 | 0.864 | 3.2 | 0.784 |
| Male Head, No Spouse | 94.3 | 1.179 | 4.7 | 1.403 | 1.0 | 0.688 |
| Other Households with Children | 97.1 | 2.460 | 0.0 | na | 2.9 | 2.460 |
| With No Children < 18 | 97.8 | 0.128 | 1.2 | 0.108 | 0.9 | 0.079 |
| More Than One Adult | 98.3 | 0.116 | 1.0 | 0.114 | 0.6 | 0.076 |
| Women Living Alone | 97.1 | 0.424 | 1.6 | 0.343 | 1.3 | 0.248 |
| Men Living Alone | 96.8 | 0.367 | 1.4 | 0.265 | 1.7 | 0.292 |
| Households with Elderly | 99.2 | 0.123 | 0.5 | 0.120 | 0.3 | 0.101 |
| Elderly Living Alone | 99.7 | 0.180 | 0.0 | 0.043 | 0.3 | 0.174 |
| Race and Hispanic Ethnicity |  |  |  |  |  |  |
| White Non-Hispanic | 97.9 | 0.093 | 1.3 | 0.073 | 0.7 | 0.064 |
| Black Non-Hispanic | 93.9 | 0.713 | 3.9 | 0.470 | 2.1 | 0.447 |
| Hispanic | 94.2 | 1.009 | 4.5 | 0.771 | 1.2 | 0.526 |
| Other Non-Hispanic | 98.6 | 0.384 | 0.9 | 0.234 | 0.5 | 0.290 |

TABLE F. 42 (continued)

|  | Food Secure |  | Food Insecure |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Without Hunger |  | With Hunger |  |
|  | Percent | Standard Error | Percent | Standard Error | Percent | Standard Error |
| Household Income-to-Poverty Ratio |  |  |  |  |  |  |
| Under 0.50 | 0.0 | na | 0.0 | na | 0.0 | na |
| Under 1.00 | 0.0 | na | 0.0 | na | 0.0 | na |
| Under 1.30 | 0.0 | na | 0.0 | na | 0.0 | na |
| Under 1.85 | 0.0 | na | 0.0 | na | 0.0 | na |
| 1.85 and Over | 97.4 | 0.126 | 1.7 | 0.082 | 0.9 | 0.087 |
| Income Not Known | 0.0 | na | 0.0 | na | 0.0 | na |
| Area of Residence |  |  |  |  |  |  |
| Inside Metropolitan Area |  |  |  |  |  |  |
| In Central City | 96.7 | 0.394 | 1.9 | 0.246 | 1.3 | 0.257 |
| Not in Central City | 97.6 | 0.224 | 1.6 | 0.150 | 0.8 | 0.101 |
| Outside Metropolitan Area | 98.0 | 0.208 | 1.4 | 0.165 | 0.6 | 0.122 |
| Census Geographical Region |  |  |  |  |  |  |
| Northeast | 97.6 | 0.170 | 1.6 | 0.077 | 0.8 | 0.127 |
| Midwest | 97.7 | 0.154 | 1.3 | 0.152 | 0.9 | 0.105 |
| South | 97.4 | 0.305 | 1.7 | 0.193 | 0.9 | 0.194 |
| West | 97.0 | 0.243 | 2.1 | 0.137 | 0.9 | 0.209 |

Tabulations of Current Population Survey, Food Security Supplement data.


[^0]:    ${ }^{1}$ The 1995-1997 files used in the current report, like the corresponding public use files available on the CPS website, have one further adjustment which is designed to make them more comparable to the 1998 data. In the 1998 survey, households were not asked relatively severe food security questions if they had already consistently answered "no" to blocks of less severe questions. The 1995-1997 data were edited to emulate this screening by replacing answers to the more severe questions with missing value codes for questions which would have been skipped in 1998. See: www.bls.census.gov/cps/foodsecu/1997/agnote.htm ("Food Security, Scales and Screener Variables").

[^1]:    ${ }^{2}$ Note, however, that there has not yet been an opportunity to assess temporal stability in the context of anything other than a strong economy.

[^2]:    ${ }^{1}$ The food security supplement has been fielded by the U.S. Census Bureau on an annual basis, including the August 1998, April 1999, September 2000, and April 2001 CPS. Current plans call for food security data to be collected henceforth in the same month each year, beginning in December 2001.

[^3]:    ${ }^{3}$ The sample attrition at this stage is due mainly to households in the CPS being told that they are about to start a new module, and declining to do so.

[^4]:    ${ }^{5}$ Some researchers view Rasch modeling as a subset of IRT theory; others disagree with that characterization. In any event, they clearly are closely related.

[^5]:    ${ }^{1}$ Under the contract for the current research, the work presented in this report includes analysis only of the surveys done through 1997.
    ${ }^{2}$ Although it is unnecessary, Rasch model software could be used, if desired.

[^6]:    Source: Bickel et al. 2000.

[^7]:    ${ }^{1}$ A methodologically less stringent ("more liberal") determination would hold that any one or two indications of the recognized and measurable conditions of food insecurity should be deemed sufficient to classify the household as food-insecure. Similarly, a methodologically more stringent determination ("more conservative") might require four or five separate indicators as the observable threshold for the condition. In setting three indicators as the operational rule for identifying the severity threshold (and scale cutpoint) for food insecurity, the technical team attempted to lean in the direction of methodological conservatism. This was done in order to hold to a very low level the likelihood of "false positive" classifications occurring with the method adopted.

[^8]:    ${ }^{2}$ Essentially, the working group defined each cutpoint as being between two adjacent groups of households on the continuous scale, such as the group of households with two affirmative answers and the group with three affirmative answers. Technical staff then chose specific points between these boundaries.

[^9]:    ${ }^{1}$ Note, however, that there has not yet been an opportunity to assess temporal stability in the context of anything other than a strong economy.

[^10]:    ${ }^{1}$ For more complete discussions of item response theory models, see Hambleton (1993) and Wright and Masters (1982).

[^11]:    ${ }^{2}$ The assumption of a single underlying dimension can be tested using principal component/factor analysis models.

[^12]:    ${ }^{3}$ In more formal terms, all model parameters are estimated using maximum likelihood procedures. The form of the likelihood function allows the item parameters to be integrated out of the expression when estimating the respondent parameters. In similar fashion, the respondent parameters can be integrated out of the likelihood expression when the item parameters are estimated. We discuss estimation of the model parameters later, in Section 3.

[^13]:    ${ }^{4}$ However, as we discuss below, using test items that properly target the ability level of the respondent results in more efficient estimates (that is, estimates with smaller standard errors) of that respondent's ability than do test items with difficulty levels that are far from the respondent's ability.

[^14]:    ${ }^{7}$ Some IRT models relax this assumption, allowing for multipart test questions. These partialcredit models take into account structural dependencies among test items. For example, situations where one question can be answered correctly only if a prior question was answered correctly can be accommodated.
    ${ }^{8}$ The assumption that the probability of a correct response does not depend on any respondent characteristics other than $b$ (in traditional applications, this is the respondent's ability; here, it is a household's level of food insecurity) can be tested in a variety of ways. In traditional testing applications, such evaluations are the basis for assessing whether test items are culturally biased, for example. Here, such an evaluation would indicate whether members of different demographic groups respond to comparable levels of food insecurity in different ways.

[^15]:    ${ }^{9}$ Actually, there is no difference in the two situations. In both cases, the vertical axis measures the expected value of the outcome. In the case of the item characteristic curves, those expected values are between 0 and 1 . For test characteristic curves, the expected values range between 0 and the number of items in the test. Item characteristic curves can be summed to create a test characteristic curve because expected values are additive.

[^16]:    ${ }^{10}$ It is appropriate, however, to recognize that a test item provides more information about respondents with locations on the latent scales close to the difficulty of the item, and less information about respondents with locations far from the item's difficulty. The key point to recognize here is that a test item provides information about a range of abilities, not a single point. We will discuss this issue in more detail later.

[^17]:    ${ }^{11}$ This is also the formula for Fisher's Information for the IRT models discussed here.

[^18]:    ${ }^{12}$ This can be seen by examining the property noted at the end of Section 2 above-that the test characteristics curves can be used to assign respondent scores based on the number of correct answers to test items.
    ${ }^{13}$ While the IRT model provides no basis for assigning scores beyond the range of the test, other statistical techniques can sometimes be used. For example, if student scale scores are believed to have a specific distributional form, scale scores can be assigned to students with extreme raw scores so that, when the student scores are combined with the scores assigned by the model, the scores do, in fact, have the assumed distribution.

[^19]:    ${ }^{14}$ In principle, it is possible to have different tests with no common items. For example, when tests are used to monitor student progress through elementary school, it is possible to test first-grade students and sixth-grade students using separate, grade-appropriate tests that measure student performance on a common scale. For this to work, there must be at least one test with some overlapping content at an earlier stage. In the example of monitoring student progress through elementary school, there would likely be at least one test per year with some overlapping content in each pair of adjacent tests. Item calibrations from those tests establish the relative difficulty of test items. Then, using those calibrations as fixed reference points, all other test items can be calibrated relative to those reference points. Once all items have been calibrated with respect to each other, the scores assigned to students will all be measured on a common scale.
    ${ }^{15}$ In April 1997, essentially two additional tests were administered. Because of an error in the questionnaire, two of the CPS rotation groups-about a quarter of the sample-were not asked one of the food security questions. For the households affected, this resulted in a 17 -question version of the test being administered to households with children, and a 9-question version administered to households without children.

[^20]:    ${ }^{16}$ For instance: BIGSTEPS Rasch-Model Computer Program, MESA Press, 5835 Kimbark Ave., Chicago IL. 60637-1609. BILOG Scientific Software International, 7833 N. Lincoln Ave., Suite 100, Lincolnwood, IL 60712-1704.

[^21]:    ${ }^{17}$ For example, the maximum likelihood estimator for the variance of a normally distributed variable is $\frac{1}{N} \sum_{i=1}^{N}\left(\bar{x}-x_{i}\right)^{2}$. An unbiased estimator, however, is $\frac{1}{N-1} \sum_{i=1}^{N}\left(\bar{x}-x_{i}\right)^{2}$.

[^22]:    ${ }^{2}$ The DHHS Poverty Guidelines are widely used to determine program eligibility (for example, by the Food Stamp Program, WIC, and the National School Lunch Program).

[^23]:    ${ }^{1}$ The percentage in the text ignores the problem created by answers to questions on one item in 1997, where the item was incorrectly administered to about 25 percent of the sample.

[^24]:    ${ }^{2}$ The approach used in the current study, as described later in the text, also ignores this information.

[^25]:    ${ }^{3}$ In converting the continuous scores of these households into food security categories, we used the midpoints of the boundary categories as the cutpoints.

[^26]:    ${ }^{2}$ In contrast, the Rasch model would assign this household the same insecurity score as any other household with six "yes" answers (and no missing responses), effectively treating it as though it had the modal response pattern of "yes" to the first 6 items and "no" to the last 12.

[^27]:    ${ }^{1}$ The horizontal coordinate of the triangle discussed in the text is the location of the least severe item (Item 53) when parameters are estimated using the full sample. The vertical coordinate of the diamond is the location of the parameter when the estimation is done only over the Black subsample.

[^28]:    ${ }^{2}$ The ten items applicable to households without children have an average value of 6.077 in the estimates for all households with children, as compared to the average of 7.000 for the entire set of 18 items.

[^29]:    PREVALENCE OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER
    
    FOR THOSE FAMILIES UNDER 130 PERCENT OF POVERTY THRESHHOLD

[^30]:    FOR THOSE FAMILIES UNDER 130 PERCENT OF POVERTY THRESHHOLD

[^31]:    PREVALENCE OF FOOD SECURITY，FOOD INSECURITY，AND HUNGER
    L66I＇SGTOHESกOH HO SDILISษGLOVYVHD đヨLつヨTヨS Xя
    FOR THOSE FAMILIES UNDER 130 PERCENT OF POVERTY THRESHHOLD

[^32]:    PREVALENCE OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER
    
    FOR THOSE FAMILIES UNDER 185 PERCENT OF POVERTY THRESHHOLD

[^33]:    PREVALENCE OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER
    BY SELECTED CHARACTERSITICS OF HOUSEHOLDS, 1996
    FOR THOSE FAMILIES UNDER 185 PERCENT OF POVERTY THRESH

[^34]:    PREVALENCE OF FOOD SECURITY, FOOD INSECURITY, AND HUNGER
    BY SELECTED CHARACTERSITICS OF HOUSEHOLDS, 1997
    FOR THOSE FAMILIES UNDER 185 PERCENT OF POVERTY THRESH

